

# Midterm Answer Key

Economics 435: Quantitative Methods

Fall 2010

## 1 Cluster samples

a) Yes.

$$\begin{aligned} E(\bar{x}) &= E\left(\frac{1}{nS} \sum_{s=1}^S \sum_{i=1}^n x_{is}\right) \\ &= \frac{1}{nS} \sum_{s=1}^S \sum_{i=1}^n E(x_{is}) \\ &= \frac{1}{nS} \sum_{s=1}^S \sum_{i=1}^n E(\mu + v_s + u_{is}) \\ &= \frac{1}{nS} \sum_{s=1}^S \sum_{i=1}^n \mu \\ &= \frac{1}{nS} (Sn\mu) \\ &= \mu \end{aligned}$$

b)

$$\begin{aligned} \text{var}(x_{is}) &= \text{var}(\mu + v_s + u_{is}) \\ &= \text{var}(v_s) + \text{var}(u_{is}) + 2\text{cov}(v_s, u_{is}) \\ &= \sigma_v^2 + \sigma_u^2 + 0 \\ &= \sigma_v^2 + \sigma_u^2 \end{aligned}$$

c) First we need to find the covariance:

$$\begin{aligned} \text{cov}(x_{is}, x_{js}) &= \text{cov}(\mu + v_s + u_{is}, \mu + v_s + u_{js}) \\ &= \text{cov}(v_s, v_s) + \text{cov}(v_s, u_{js}) + \text{cov}(u_{is}, v_s) + \text{cov}(u_{is}, u_{js}) \\ &= \sigma_v^2 + 0 + 0 + 0 \\ &= \sigma_v^2 \end{aligned}$$

and then we plug into the usual formula for the correlation:

$$\begin{aligned} \text{corr}(x_{is}, x_{js}) &= \frac{\text{cov}(x_{is}, x_{js})}{\sqrt{\text{var}(x_{is})\text{var}(x_{js})}} \\ &= \frac{\sigma_v^2}{\sqrt{(\sigma_v^2 + \sigma_u^2)(\sigma_v^2 + \sigma_u^2)}} \\ &= \frac{\sigma_v^2}{\sigma_v^2 + \sigma_u^2} \end{aligned}$$

d)

$$\begin{aligned} \text{var}(\bar{x}) &= \text{var}\left(\frac{1}{nS} \sum_{s=1}^S \sum_{i=1}^n x_{is}\right) \\ &= \text{var}\left(\frac{1}{nS} \sum_{s=1}^S \sum_{i=1}^n \mu + v_s + u_{is}\right) \\ &= \frac{1}{(nS)^2} \text{var}\left(\sum_{s=1}^S \sum_{i=1}^n \mu + \sum_{s=1}^S \sum_{i=1}^n v_s + \sum_{s=1}^S \sum_{i=1}^n u_{is}\right) \\ &= \frac{1}{(nS)^2} \text{var}\left(nS\mu + \sum_{s=1}^S nv_s + \sum_{s=1}^S \sum_{i=1}^n u_{is}\right) \\ &= \frac{1}{(nS)^2} \left(\text{var}(nS\mu) + \sum_{s=1}^S \text{var}(nv_s) + \sum_{s=1}^S \sum_{i=1}^n \text{var}(u_{is})\right) \\ &= \frac{1}{(nS)^2} (0 + n^2 S \sigma_v^2 + nS \sigma_u^2) \\ &= \frac{\sigma_v^2}{S} + \frac{\sigma_u^2}{nS} \end{aligned}$$

e) Yes. I didn't ask for proof, but here it is. In a random sample of size  $nS$ :

$$\begin{aligned} \text{var}(\bar{x}) &= \frac{\text{var}(x)}{nS} \\ &= \frac{\sigma_v^2 + \sigma_u^2}{nS} \\ &= \frac{\sigma_v^2}{nS} + \frac{\sigma_u^2}{nS} \\ &< \frac{\sigma_v^2}{S} + \frac{\sigma_u^2}{nS} \end{aligned}$$

f) Yes. To see this note that:

$$\begin{aligned} \lim_{S \rightarrow \infty} \text{var}(\bar{x}) &= \lim_{S \rightarrow \infty} \left(\frac{\sigma_v^2}{S} + \frac{\sigma_u^2}{nS}\right) \\ &= 0 \end{aligned}$$

This implies that  $\bar{x} \xrightarrow{m.s.} \mu$ , so  $\bar{x} \xrightarrow{p} \mu$  as well.

g) No. To see this note that:

$$\begin{aligned}\lim_{n \rightarrow \infty} \text{var}(\bar{x}) &= \lim_{n \rightarrow \infty} \left( \frac{\sigma_v^2}{S} + \frac{\sigma_u^2}{nS} \right) \\ &= \frac{\sigma_v^2}{S} \neq 0\end{aligned}$$

## 2 Permanent income and the black-white test score gap

a)

$$\begin{aligned}\text{GAP1}(C) &= E(s|b=1, c=C) - E(s|b=0, c=C) \\ &= (\gamma_0 + \gamma_1 + \gamma_2 C) - (\gamma_0 + \gamma_2 C) \\ &= \gamma_1\end{aligned}$$

b)

$$\begin{aligned}\text{GAP2}(P) &= E(s|b=1, p=P) - E(s|b=0, p=P) \\ &= (\beta_0 + \beta_1 + \beta_2 P) - (\beta_0 + \beta_2 P) \\ &= \beta_1\end{aligned}$$

c) Yes.

d)

$$\widehat{\text{GAP1}}(C) = \hat{\gamma}_1$$

e) Let's try the hint out:

$$\begin{aligned}s &= \beta_0 + \beta_1 b + \beta_2 p + u \\ &= \beta_0 + \beta_1 b + \beta_2(\delta_0 + \delta_1 b + \delta_2 c + w) + u \\ &= (\beta_0 + \beta_2 \delta_0) + (\beta_1 + \beta_2 \delta_1) b + \beta_2 \delta_2 c + (\beta_2 w + u)\end{aligned}$$

Now  $E(\beta_2 w + u|b, c) = \beta_2 E(w|b, c) + E(u|b, c) = \beta_2 * 0 + 0 = 0$ . This means that the model above fits equation (2) where  $\gamma_0 = (\beta_0 + \beta_2 \delta_0)$ ,  $\gamma_1 = (\beta_1 + \beta_2 \delta_1)$ ,  $\gamma_2 = \beta_2 \delta_2$ , and  $v = (\beta_2 w + u)$ . Solving for  $\beta_1$  we get:

$$\beta_1 = \gamma_1 - \delta_1 \gamma_2 / \delta_2$$

This yields the estimator:

$$\begin{aligned}\widehat{\text{GAP2}}(P) &= \hat{\beta}_1 \\ &= \hat{\gamma}_1 - \hat{\delta}_1 \hat{\gamma}_2 / \hat{\delta}_2\end{aligned}$$

f)

$$\begin{aligned}\widehat{\text{GAP1}}(C) &= \hat{\gamma}_1 \\ &= -0.426 \\ \widehat{\text{GAP2}}(P) &= \hat{\gamma}_1 - \hat{\delta}_1 \hat{\gamma}_2 / \hat{\delta}_2 \\ &= -0.426 - (-0.306 * 0.205 / 0.524) \\ &= -0.306\end{aligned}$$

By the way, you might have noticed that  $\widehat{\text{GAP2}}(P)$  and  $\hat{\delta}_2$  are both -0.306, but that's just a coincidence.

g) Yes.