

Exam #2 Answer Key

Economics 435: Quantitative Methods

Fall 2011

1 Short answers

a) Hopefully you've done this a few times before, but here it is again:

$$\begin{aligned} E(u) &= E(E(u|x)) && \text{(law of iterated expectations)} \\ &= E(0) && \text{(given)} \\ &= 0 \\ E(xu) &= E(E(xu|x)) && \text{(law of iterated expectations)} \\ &= E(xE(u|x)) && \text{(conditioning)} \\ &= E(x0) && \text{(given)} \\ &= 0 \\ \text{cov}(x, u) &= E(xu) - E(x)E(u) && \text{(a result we use all the time)} \\ &= 0 - E(x)0 \\ &= 0 \end{aligned}$$

b)

$$\begin{aligned} \text{plim } e^{\bar{x}} &= e^{\text{plim } \bar{x}} && \text{(Slutsky theorem)} \\ &= e^{\mu} && \text{(law of large numbers)} \end{aligned}$$

c) If $g(\theta) = e^\theta$, then the rules of differentiation imply that $g'(\theta) = e^\theta$ as well. So:

$$\sqrt{n}(e^{\bar{x}} - e^\mu) \rightarrow^D N(0, e^{2\mu}\sigma^2)$$

d) The Central Limit Theorem tells us that the sample average of a large random sample is approximately normally distributed. This is useful for performing inference (hypothesis tests and confidence intervals) because inference requires us to know the probability distribution of our statistics - most of the time the finite-sample distribution is intractable, so having a convenient approximation is very helpful.

2 The 2011 Census

a) We now have a random sample of size $n = NSR$, so

$$\text{var}(\bar{y}) = \frac{\sigma_y^2}{NSR}$$

b)

$$\begin{aligned} \sqrt{\frac{\text{var}(\bar{y}; S = 0.333333, R = 0.5)}{\text{var}(\bar{y}; S = 0.2, R = 0.95)}} &= \sqrt{\frac{\frac{\sigma_y^2}{N \times 0.333 \times 0.5}}{\frac{\sigma_y^2}{N \times 0.2 \times 0.95}}} \\ &= \sqrt{\frac{0.19}{0.1667}} \\ &= \sqrt{1.14} \\ &= 1.07 \end{aligned}$$

It raises the variance by about 14%, and the standard error by about 7%.

c) No.

d) Since $E(y_i|r_i = 0) \in [a, b]$, we have

$$E(y_i|r_i = 1)R + a(1 - R) \leq E(y_i) \leq E(y_i|r_i = 1)R + b(1 - R)$$

e) The lower bound would be:

$$\bar{y}R + a(1 - R)$$

and the upper bound would be:

$$\bar{y}R + b(1 - R)$$

f) The percentage lies between $0.10 * 0.95$ and $0.10 * 0.95 + 0.05$, or 9.5% and 14.5%.

g) The percentage lies between $0.10 * 0.5$ and $0.10 * 0.5 + 0.5$ or between 5% and 55%.

h)

$$\begin{aligned} \theta &= E(y_i|x_i = 1) - E(y_i|x_i = 0) \\ &= (E(y_i|x_i = 1, r_i = 1) \Pr(r_i = 1) + E(y_i|x_i = 1, r_i = 0) \Pr(r_i = 0)) \\ &\quad - (E(y_i|x_i = 0, r_i = 1) \Pr(r_i = 1) + E(y_i|x_i = 0, r_i = 0) \Pr(r_i = 0)) \\ &= (E(y_i|x_i = 1, r_i = 1) - E(y_i|x_i = 1, r_i = 0)) \Pr(r_i = 1) \\ &\quad + (E(y_i|x_i = 1, r_i = 0) - E(y_i|x_i = 0, r_i = 0)) \Pr(r_i = 0) \\ &= \beta_1 R + (E(y_i|x_i = 1, r_i = 0) - E(y_i|x_i = 0, r_i = 0)) (1 - R) \end{aligned}$$

We can determine that $(E(y_i|x_i = 1, r_i = 0) - E(y_i|x_i = 0, r_i = 0))$ lies between $(0 - 1) = -1$ and $(1 - 0) = 1$, so

$$\beta_1 R - (1 - R) \leq \theta \leq \beta_1 R + (1 - R)$$

i) The bounds would be:

$$\hat{\beta}_1 R - (1 - R) \leq \theta \leq \hat{\beta}_1 R + (1 - R)$$

j)

$$-0.05 = (0 \times 0.95) - (1 - 0.95) \leq \theta \leq (0 \times 0.95) + (1 - 0.95) = 0.05$$

So being in a female-headed household may raise or lower the probability of living in poverty by as much as 5%.

k)

$$-0.5 = (0 \times 0.5) - (1 - 0.5) \leq \theta \leq (0 \times 0.5) + (1 - 0.5) = 0.5$$

So being in a female-headed household may raise or lower the probability of living in poverty by as much as 50%.

3 Cluster data with fixed effects

a) First we show:

$$\begin{aligned} \bar{y}_s &= \frac{1}{n_s} \sum_{i:s(i)=s} y_i \\ &= \frac{1}{n_s} \sum_{i:s(i)=s} a_{s(i)} + \beta_1 FDK_{s(i)t(i)} + \beta_2 x_i + u_i \\ &= \frac{1}{n_s} \sum_{i:s(i)=s} a_s + \beta_1 \frac{1}{n_s} \sum_{i:s(i)=s} FDK_{st(i)} + \frac{1}{n_s} \sum_{i:s(i)=s} \beta_2 x_i + \frac{1}{n_s} \sum_{i:s(i)=s} u_i \\ &= a_s + \beta_1 \overline{FDK}_s + \beta_2 \bar{x}_s + \bar{u}_s \\ y_i - \bar{y}_{s(i)} &= (a_{s(i)} + \beta_1 FDK_{s(i)t(i)} + \beta_2 x_i + u_i) - (a_{s(i)} + \beta_1 \overline{FDK}_{s(i)} + \beta_2 \bar{x}_{s(i)} + \bar{u}_{s(i)}) \\ &= \beta_1 (FDK_{s(i)t(i)} - \overline{FDK}_{s(i)}) + \beta_2 (x_i - \bar{x}_{s(i)}) + (u_i - \bar{u}_{s(i)}) \end{aligned}$$

Then we show:

$$\begin{aligned}
& E(u_i - \bar{u}_{s(i)} | (FDK_{s(i)t(i)} - \overline{FDK}_{s(i)}), (x_i - \bar{x}_{s(i)})) \\
= & E\left(E\left(u_i - \bar{u}_{s(i)} | a_{s(i)}, \{FDK_{s(i)t}\}_{t=2009}^{2011}, \{x_j\}_{j:s(j)=s(i)}\right) \middle| (FDK_{s(i)t(i)} - \overline{FDK}_{s(i)}), (x_i - \bar{x}_{s(i)})\right) \\
& \text{(by the Law of Iterated Expectations)} \\
= & E\left(E\left(u_i - \frac{1}{n_{s(i)}} \sum_{j:s(j)=s(i)} u_j \middle| a_{s(i)}, \{FDK_{s(i)t}\}_{t=2009}^{2011}, \{x_j\}_{j:s(j)=s(i)}\right) \middle| (FDK_{s(i)t(i)} - \overline{FDK}_{s(i)}), (x_i - \bar{x}_{s(i)})\right) \\
& \text{(by substitution)} \\
= & E\left(E\left(u_i | a_{s(i)}, \{FDK_{s(i)t}\}_{t=2009}^{2011}, \{x_j\}_{j:s(j)=s(i)}\right) \middle| (FDK_{s(i)t(i)} - \overline{FDK}_{s(i)}), (x_i - \bar{x}_{s(i)})\right) \\
& - \frac{1}{n_{s(i)}} \sum_{j:s(j)=s(i)} E\left(E\left(u_j | a_{s(i)}, \{FDK_{s(i)t}\}_{t=2009}^{2011}, \{x_j\}_{j:s(j)=s(i)}\right) \middle| (FDK_{s(i)t(i)} - \overline{FDK}_{s(i)}), (x_i - \bar{x}_{s(i)})\right) \\
& \text{(by linearity of the expected value)} \\
= & E\left(0 | (FDK_{s(i)t(i)} - \overline{FDK}_{s(i)}), (x_i - \bar{x}_{s(i)})\right) - \frac{1}{n_{s(i)}} \sum_{j:s(j)=s(i)} E\left(0 | (FDK_{s(i)t(i)} - \overline{FDK}_{s(i)}), (x_i - \bar{x}_{s(i)})\right) \\
& \text{(by strict exogeneity)} \\
= & 0 - \frac{1}{n_{s(i)}} \sum_{j:s(j)=s(i)} 0 \\
= & 0
\end{aligned}$$

b) First we show:

$$\begin{aligned}
\bar{y}_{st} &= \frac{1}{n_{st}} \sum_{i:s(i)=s, t(i)=t} y_i \\
&= \frac{1}{n_{st}} \sum_{i:s(i)=s, t(i)=t} a_{s(i)} + \beta_1 FDK_{s(i)t(i)} + \beta_2 x_i + u_i \\
&= \frac{1}{n_{st}} \sum_{i:s(i)=s, t(i)=t} a_s + \beta_1 \frac{1}{n_{st}} \sum_{i:s(i)=s, t(i)=t} FDK_{st} + \frac{1}{n_{st}} \sum_{i:s(i)=s, t(i)=t} \beta_2 x_i + \frac{1}{n_{st}} \sum_{i:s(i)=s, t(i)=t} u_i \\
&= a_s + \beta_1 FDK_{st} + \beta_2 \bar{x}_{st} + \bar{u}_{st} \\
\bar{y}_{st} - \bar{y}_s &= (a_s + \beta_1 FDK_{st} + \beta_2 \bar{x}_{st} + \bar{u}_{st}) - (a_s + \beta_1 \overline{FDK}_s + \beta_2 \bar{x}_s + \bar{u}_s) \\
&= \beta_1 (FDK_{st} - \overline{FDK}_s) + \beta_2 (\bar{x}_{st} - \bar{x}_s) + (\bar{u}_{st} - \bar{u}_s)
\end{aligned}$$

Then we show:

$$\begin{aligned}
& E(\bar{u}_{st} - \bar{u}_s | (FDK_{st} - \overline{FDK}_s), (\bar{x}_{st} - \bar{x}_s)) \\
&= E\left(E\left(\bar{u}_{st} - \bar{u}_s | a_s, \{FDK_{st}\}_{t=2009}^{2011}, \{x_j\}_{j:s(j)=s}\right) \middle| (FDK_{st} - \overline{FDK}_s), (\bar{x}_{st} - \bar{x}_s)\right) \\
&\quad \text{(by the Law of Iterated Expectations)} \\
&= E\left(E\left(\frac{1}{n_{st}} \sum_{j:s(j)=s, t(j)=t} u_j - \frac{1}{n_s} \sum_{j:s(j)=s} u_j \middle| a_s, \{FDK_{st}\}_{t=2009}^{2011}, \{x_j\}_{j:s(j)=s}\right) \middle| (FDK_{st} - \overline{FDK}_s), (\bar{x}_{st} - \bar{x}_s)\right) \\
&\quad \text{(by substitution)} \\
&= \frac{1}{n_{st}} \sum_{j:s(j)=s, t(j)=t} E\left(E\left(u_j | a_s, \{FDK_{st}\}_{t=2009}^{2011}, \{x_j\}_{j:s(j)=s}\right) \middle| (FDK_{st} - \overline{FDK}_s), (\bar{x}_{st} - \bar{x}_s)\right) \\
&\quad - \frac{1}{n_s} \sum_{j:s(j)=s} E\left(E\left(u_j | a_s, \{FDK_{st}\}_{t=2009}^{2011}, \{x_j\}_{j:s(j)=s}\right) \middle| (FDK_{st} - \overline{FDK}_s), (\bar{x}_{st} - \bar{x}_s)\right) \\
&\quad \text{(by linearity of the expected value)} \\
&= \frac{1}{n_{st}} \sum_{j:s(j)=s, t(j)=t} E(0 | (FDK_{st} - \overline{FDK}_s), (\bar{x}_{st} - \bar{x}_s)) \\
&\quad - \frac{1}{n_s} \sum_{j:s(j)=s} E(0 | (FDK_{st} - \overline{FDK}_s), (\bar{x}_{st} - \bar{x}_s)) \\
&\quad \text{(by strict exogeneity)} \\
&= \left(\frac{1}{n_{st}} \sum_{j:s(j)=s, t(j)=t} 0\right) - \left(\frac{1}{n_s} \sum_{j:s(j)=s} 0\right) \\
&= 0
\end{aligned}$$

c) Estimate the OLS regression of $(y_i - \bar{y}_{s(i)})$ on $(FDK_{s(i)t(i)} - \overline{FDK}_{s(i)})$ and $(x_i - \bar{x}_{s(i)})$.

d) Estimate the OLS regression of $(y_{st} - \bar{y}_s)$ on $(FDK_{st} - \overline{FDK}_s)$ and $(\bar{x}_{st} - \bar{x}_s)$.