

Lecture 3 - Random walks at large N

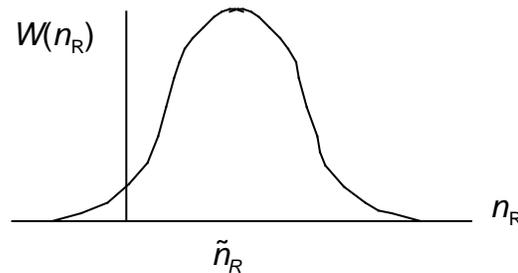
What's Important:

- random walks in the continuum limit
- Text:* Reif

1D Random walks at large N

As the number of steps N in a walk becomes very large, the individual step sizes in n_R become relatively small. Further, the difference in values of $W(n_R)$ between successive values of n_R becomes small as well. The proof of how the difference scales with N is given in Reif. For our purposes, the important point is that we can regard n_R and $W(n_R)$ as continuous at large N .

To obtain a continuous description of $W(n_R)$, consider its form near its most likely value at \tilde{n}_R



At the peak of the distribution, its derivative with respect to \tilde{n}_R vanishes

$$\left. \frac{dW}{dn_R} \right|_{\tilde{n}_R} = 0 \quad \text{or} \quad \left. \frac{d \ln W}{dn_R} \right|_{\tilde{n}_R} = 0$$

Let's expand $W(n_R)$ or its logarithm around \tilde{n}_R using η as an expansion parameter, where

$$n_R = \tilde{n}_R + \eta.$$

Choosing the logarithm, for example, we have

$$\ln W(n_R) = \ln W(\tilde{n}_R) + B_1 \eta + \frac{1}{2} B_2 \eta^2 + \dots \quad (3.1)$$

where

$$B_i = \left. \frac{d^i \ln W}{dn_R^i} \right|_{\tilde{n}_R}$$

At the maximum in W , the coefficient $B_1 = 0$ by definition. The coefficients B_i can be obtained from the discrete expression for W in terms of factorials. Start with the definition

$$\ln W(n_R) = \ln N! - \ln n_R! - \ln(N - n_R)! + n_R \ln p + (N - n_R) \ln q,$$

so

$$\frac{d \ln W(n_R)}{dn_R} = 0 - \frac{d}{dn_R} \ln n_R! - \frac{d}{dn_R} \ln(N - n_R)! + \ln p - \ln q. \quad (3.2)$$

Now, the derivatives of the factorials can be approximated by

$$\frac{d \ln n!}{dn} = \frac{\ln(n+1)! - \ln n!}{1} = \ln \frac{(n+1)!}{n!} = \ln(n+1) \quad (3.3)$$

Thus, Eq. (3.2) becomes

$$\frac{d \ln W(n_R)}{dn_R} = -\ln n_R + \ln(N - n_R) + \ln \frac{p}{q} \quad (3.4)$$

where care has been taken with the sign of the derivative of $\ln(N - n_R)!$

Eq. (3.4) gives an expression for B_1 , which can be used to obtain

$$\tilde{n}_R = Np \quad (3.5)$$

by setting $B_1 = 0$. Of course, we already know Eq. (3.5) from general considerations of the random walk.

Pressing on to extract B_2 , we first take another derivative of W :

$$\begin{aligned} \frac{d^2 \ln W(n_R)}{dn_R^2} &= \frac{d}{dn_R} \left(-\ln n_R + \ln(N - n_R) + \ln \frac{p}{q} \right) \\ &= -\frac{1}{n_R} + (-1) \frac{1}{N - n_R} + 0 \\ &= -\frac{1}{n_R} - \frac{1}{N - n_R} \end{aligned} \quad (3.6)$$

We evaluate this at $n_R = \tilde{n}_R = Np$ to determine B_2

$$\begin{aligned} B_2 &= \left. \frac{d^2 \ln W(n_R)}{dn_R^2} \right|_{\tilde{n}_R} = -\frac{1}{\tilde{n}_R} - \frac{1}{N - \tilde{n}_R} \\ &= -\frac{1}{Np} - \frac{1}{N - Np} \\ &= -\frac{1}{N} \frac{1}{p} + \frac{1}{1-p} \\ &= -\frac{1}{N} \cdot \frac{1-p+p}{p(1-p)} \\ &= -\frac{1}{N} \cdot \frac{1}{pq} \end{aligned} \quad (3.7)$$

Higher-order terms in B are smaller by successive powers of N and are not important for our level of approximation.

Thus, inserting B_1 and B_2 into the expansion (3.1) gives

$$\ln W(n_R) = \ln W(\tilde{n}_R) + 0 + \frac{1}{2} \frac{1}{Npq} \eta^2 + \dots$$

With the definition

$$\tilde{W} = W(\tilde{n}_R)$$

this expression yields

$$W(n_R) = \tilde{W} e^{-\eta^2 / 2Npq} \tag{3.8}$$

The implications of this Gaussian form will be discussed momentarily. But first, we obtain an expression for the prefactor by means of the normalization condition

$$\int_{-\infty}^{\infty} W(\tilde{n}_R + \eta) d\eta = 1$$

where we assume that N is so large that the integration limits are effectively infinite. Substituting Eq. (3.8) for $W(n_R)$ gives

$$\begin{aligned} 1 &= \tilde{W} \int_{-\infty}^{\infty} e^{-\eta^2 / 2Npq} d\eta \\ &= \tilde{W} (2Npq)^{1/2} \int_{-\infty}^{\infty} e^{-x^2} dx \\ &= \tilde{W} (2Npq)^{1/2} \sqrt{\pi} \end{aligned}$$

Inverting

$$\tilde{W} = (2 Npq)^{-1/2} \tag{3.9}$$

The proof for the value of the integral is

Define $C = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \exp(-y^2) dy$ so that

$$\begin{aligned} C^2 &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \exp(-x^2) dx \int_{-\infty}^{\infty} \exp(-y^2) dy \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \exp[-(x^2+y^2)] dx dy \\ &= \int_0^{2\pi} d\theta \int_0^{\infty} \exp(-r^2) r dr \\ &= 2 \int_0^{\infty} (1/2) \exp(-z) dz \quad (z = r^2) \\ &= \int_0^{\infty} \exp(-z) dz \end{aligned}$$

Thus, $C = \sqrt{\pi}$.

At long last, we have the complete form, after substituting $\eta = n_R - \tilde{n}_R = n_R - Np$:

$$W(n_R) = \frac{1}{(2 Npq)^{1/2}} \exp \left[-\frac{(n_R - Np)^2}{2Npq} \right] \tag{3.10}$$

An alternate form is to write this in terms of n_R and its dispersion

$$\overline{n_R^2} = Npq$$

so that

$$W(n_R) = \frac{1}{(2 \overline{n_R^2})^{1/2}} \exp -\frac{(n_R - \overline{n_R})^2}{2 \overline{n_R^2}} \quad (3.11)$$

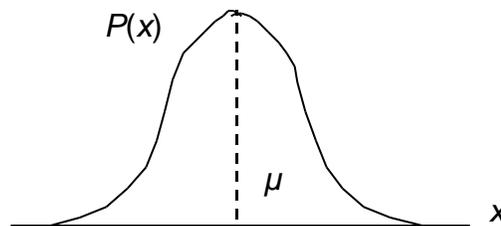
This distribution is of a Gaussian form, peaked at its mean value. We've done this for a *number*, which is a special case of the Gaussian probability *density*

$$P(x) = \frac{1}{\sqrt{2} \sigma} \exp -\frac{(x - \mu)^2}{2\sigma^2} \quad (3.12)$$

defined such that the probability of a variable x having a value between x and $x + dx$ is $P(x) dx$.

Identifying parameters in the distributions leads to

$$\sigma^2 = \overline{n_R^2} \quad \mu = \overline{n_R}. \quad (3.13)$$



Note that the full width at half maximum **IS NOT 2σ but rather 2.355σ** .

The distribution in displacements is also Gaussian. To find $P(m)$, which is the probability of m steps to the right, use

$$N = n_R + n_L \quad \text{and} \quad m = n_R - n_L$$

to give

$$n_R = (N + m) / 2.$$

Thus,

$$\begin{aligned} P(m) &= W(n_R = [N + m] / 2) = \frac{1}{(2 Npq)^{1/2}} \exp -\frac{([N + m] / 2 - Np)^2}{2Npq} \\ &= \frac{1}{(2 Npq)^{1/2}} \exp -\frac{(m - N[p - q])^2}{8Npq} \end{aligned} \quad (3.14)$$

Now, $P(m)$ is the probability of finding a given value of m in an ensemble of walks with N steps. If we wish m to be continuous, rather than discrete, then we must work with probability densities:

$$\begin{aligned} &[\text{probability of finding } y \text{ in the range } y \text{ to } y + dy] \\ &= [\text{probability density at } y] \cdot [\text{range } dy] \end{aligned}$$

or

$$[\text{probability of finding } y \text{ in the range } y \text{ to } y + dy] = P(y) dy.$$

As applied to random walks, this becomes

$$P(n_R) dn_R = W(n_R) dn_R = (1/2)W(n_R) dm$$

where the last equality arises from

$$\frac{dm}{dn_R} = \frac{d}{dn_R} (2n_R - N) = 2$$

In other words, because m increases by 2 units for every increase in n_R by one unit:

$$\begin{aligned} & [\text{probability of finding } n_R \text{ between } n_r \text{ and } n_R + 1] \\ & = [\text{probability of finding } m \text{ between } m \text{ and } m + 2] \end{aligned}$$

The same argument can be repeated to give probabilities as a function of x , which has dimensions of length. Let the step size be l , so

$$\begin{aligned} & [\text{probability of finding } n_R \text{ between } n_R \text{ and } n_R + 1] \\ & = [\text{probability of finding } x \text{ between } x \text{ and } x + 2l] \end{aligned}$$

or

$$P(x) dx = [W(n_R) / 2l] dx. \tag{3.15}$$