## STAT-285 Homework 7 Solutions

## Section 10.2 Question $15 \quad / 4$

Study Objective: Determine which of the six different concrete mixtures have differ with respect to their resistivity.

Formulation: Let

- $X_{i j}$ denote the $j$ th measurement for the $i$ th concrete mixture, with $i=1, \cdots, 6$, and $j=1, \cdots, 26$ (ie $I=6$ and $J=26$ ).

We assume that $X_{i j} \sim N\left(\mu_{i}, \sigma^{2}\right)$, and we are given that

$$
\begin{gathered}
\bar{X}_{1 .}=14.18, \bar{X}_{2 .}=17.94, \bar{X}_{3 .}=18, \\
\bar{X}_{4 .}=18, \bar{X}_{5 .}=25.74, \bar{X}_{6 .}=27.67
\end{gathered}
$$

Method: Apply Tukey's method to identify significant differences.

Using $\alpha=0.05$, we find that $Q_{0.05,6,150}=4.08$
(in R: qtukey $(0.95$, nmeans $=6, \mathrm{df}=150)$ ). Since the question also gives $M S E=$ 13.929, we have

$$
W=Q_{0.05,6,150} \sqrt{\frac{M S E}{J}}=4.08 \sqrt{\frac{13.929}{26}}=2.9863
$$

Table 1 presents the sample mean differences between the six concrete mixtures. Note that if $\bar{X}_{j .}-\bar{X}_{i .}<W$, this implies that zero lies in the corresponding confidence interval. That is, using an underscoring pattern, we summarize our findings as

| $i$ | 1 | 2 | 3 | 4 | 5 | 6 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\bar{X}_{i .}$ | 14.18 | 17.94 | 18 | 18 |  | 25.74 |
|  |  |  | 27.67 |  |  |  |

To interpret this, we see that there is no significant differences between

- concrete mixture 2 from concrete mixtures 3 or 4
- concrete mixture 5 from concrete mixture 6

Table 1: $\bar{X}_{j .}-\bar{X}_{i \text {. for }} j>i$. The bold-faced elements correspond to the values less than $W$.

| $i \backslash j$ | 1 | 2 | 3 | 4 | 5 | 6 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 1 |  | 3.76 | 3.82 | 3.82 | 11.56 | 13.49 |
| 2 |  |  | $\mathbf{0 . 0 6}$ | $\mathbf{0 . 0 6}$ | 7.8 | 9.73 |
| 3 |  |  |  | $\mathbf{0}$ | 7.74 | 9.67 |
| 4 |  |  |  |  | 7.74 | 9.67 |
| 5 |  |  |  |  |  | $\mathbf{1 . 9 3}$ |
| 6 |  |  |  |  |  |  |

## Section 11.1 Question $4 \quad / 14$

Study Objective: Determine if there exists a difference between the mean coverage of light-bulb interior latex paint between brands of paint and rollers.

Formulation: Let

- $X_{i j}$ denote the observation for the $i$ th brand of paint and $j$ th roller brand, with $i=$ $1,2,3,4$, and $j=1,2,3$ (ie $I=4$ and $J=3$ ).

We assume that $X_{i j} \sim N\left(\mu_{i j}, \sigma^{2}\right)$, where $\mu_{i j}=\mu+\alpha_{i}+\beta_{j}$, with $\sum_{i=1}^{4} \alpha_{i}=0$ and $\sum_{j=1}^{3} \beta_{j}=0$. We see that

- $\bar{X}_{. .}=\sum_{i=1}^{4} \sum_{j=1}^{3} X_{i j} / 12=44.75$
- $\bar{X}_{i .}=\sum_{j=1}^{3} X_{i j} / 3$

$$
\bar{X}_{1 .}=50.33, \bar{X}_{2 .}=45.67, \bar{X}_{3 .}=41.67, \bar{X}_{4 .}=41.33
$$

$$
\bar{X}_{. j}=\sum_{i=1}^{4} X_{i j} / 4
$$

$$
\bar{X}_{.1}=45.75, \bar{X}_{.2}=42.25, \bar{X}_{.3}=46.25
$$

## Part A /6

We are to fill out the following ANOVA table:

- degrees of freedom for paint brand is $I-1=3$
- degrees of freedom for roller brand is $J-1=2$
- degrees of freedom for error is $(I-1)(J-1)=6$

| Source | df | Sum of Squares | Mean Square | $F$ |
| :--- | :--- | :--- | :--- | :--- |
| Paint Brand (Factor A) | $I-1$ | $S S A$ | $M S A$ | $F_{A}$ |
| Roller Brand (Factor B) | $J-1$ | $S S B$ | $M S B$ | $F_{B}$ |
| Error | $(I-1)(J-1)$ | $S S E$ | $M S E$ |  |
| Total | $I J-1$ | $S S T$ |  |  |

- total degrees of freedom is $I J-1=11$
- $S S T=\sum_{i=1}^{4} \sum_{j=1}^{3}\left(X_{i j}-\bar{X}_{. .}\right)^{2}=238.25$
- $S S A=J \sum_{i=1}^{4}\left(\bar{X}_{i .}-\bar{X}_{. .}\right)^{2}=159.5833$
- $S S B=I \sum_{j=1}^{3}\left(\bar{X}_{. j}-\bar{X}_{. .}\right)^{2}=38$
- $S S E=S S T-S S A-S S B=40.6667$
- $M S A=S S A /(I-1)=53.1944$
- $M S B=S S B /(J-1)=19$
- $M S E=S S E /((I-1)(J-1))=6.7778$
- $F_{A}=M S A / M S E=7.85$
- $F_{B}=M S B / M S E=2.80$


## Part B /2

Hypothesis Test: $H_{0 A}: \alpha_{1}=\alpha_{2}=\alpha_{3}=\alpha_{4}=0$ vs. $H_{a A}$ : At least one $\alpha_{i} \neq 0$

## Test Statistic:

$$
F_{A}=\frac{M S A}{M S E} \sim F(3,6)
$$

From Part A, we have $F_{A, o b s}=7.85$

Method 1 - p-value: $P_{H_{0}}\left(F>F_{A, \text { obs }}\right)=0.0169$, where $F \sim F(3,6)$.
Since $0.0169<0.05$, we reject $H_{0 A}$.

## Method 2 - Rejection Region:

$$
\begin{aligned}
\mathcal{R}_{A, 0.05} & =\left\{f: f>f_{0.05}(3,6)\right\} \\
& =\{f: f>4.76\} .
\end{aligned}
$$

Since $F_{A, o b s} \in \mathcal{R}_{A, 0.05}$, we reject $H_{0 A}$.

## Part C /2

Hypothesis Test: $H_{0 B}: \beta_{1}=\beta_{2}=\beta_{3}=0$ vs. $H_{a B}$ : At least one $\beta_{j} \neq 0$

## Test Statistic:

$$
F_{B}=\frac{M S B}{M S E} \sim F(2,6)
$$

From Part A, we have $F_{B, o b s}=2.80$

Method 1 - p-value: $P_{H_{0}}\left(F>F_{B, o b s}\right)=0.1381$, where $F \sim F(2,6)$.
Since $0.05<0.1381$, we fail to reject $H_{0 B}$.

## Method 2 - Rejection Region:

$$
\begin{aligned}
\mathcal{R}_{B, 0.05} & =\left\{f: f>f_{0.05}(2,6)\right\} \\
& =\{f: f>5.14\} .
\end{aligned}
$$

Since $F_{B, o b s} \notin \mathcal{R}_{B, 0.05}$, we fail to reject $H_{0 B}$.

## Part D /4

Since we fail to reject $H_{0 B}$ in Part C, we only need to use Tukey's method to identify significant differences among the paint brands.

We start by computing

$$
W=Q_{0.05, I,(I-1)(J-1)} \sqrt{\frac{M S E}{J}}=4.90 \sqrt{\frac{6.7778}{3}}=7.3651
$$

Table 2 presents the sample mean differences between the paint brands. We summarize our findings with the following underscoring pattern. (Note that I subtracted off 400 from each observation to simplify the computing!)

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\]

Although there is not a significant difference between paint brands 1 and 2, we can see paint brand 1 appears to be the preferable paint brand.

Table 2: $\bar{X}_{k .}-\bar{X}_{i}$. for $k>i$. The bold-faced elements correspond to the values less than $W$.

| $i \backslash k$ | 1 | 2 | 3 | 4 |
| :--- | :--- | :--- | :--- | :--- |
| 1 |  | $\mathbf{4 . 6 6 6 7}$ | 8.6667 | 9 |
| 2 |  |  | $\mathbf{4}$ | $\mathbf{4 . 3 3 3 3}$ |
| 3 |  |  |  | $\mathbf{0 . 3 3 3 3}$ |
| 4 |  |  |  |  |

## Section 11.2 Question 16 /14

Study Objective: Determine if there exists an effect of curing time and mixture type on the comprehensive strength of hardened cement cubes

Formulation: Let

- $X_{i j k}$ denote the $k$ th observation for the $i$ th curing time and $j$ th mixture type, with $i=1,2,3, j=1,2,3,4$, and $k=1,2,3$ (ie $I=3, J=4$, and $K=3$ ).
We assume that $X_{i j k} \sim N\left(\mu_{i j}, \sigma^{2}\right)$, where $\mu_{i j}=\mu+\alpha_{i}+\beta_{j}+\gamma_{i j}$, with $\sum_{i=1}^{3} \alpha_{i}=0 \sum_{j=1}^{4} \beta_{j}=0$, $\sum_{i=1}^{3} \gamma_{i j}=0$ for each $j$, and $\sum_{j=1}^{4} \gamma_{i j}$ for each $i$.


## Part A /6

We are to fill out the following ANOVA table:

| Source | df | Sum of Squares | Mean Square | $F$ |
| :--- | :--- | :--- | :--- | :--- |
| Curing Time (Factor A) | $I-1$ | $S S A$ | $M S A$ | $F_{A}$ |
| Mixture Type (Factor B) | $J-1$ | $S S B$ | $M S B$ | $F_{B}$ |
| Interaction | $(I-1)(J-1)$ | $S S A B$ | $M S A B$ | $F_{A B}$ |
| Error | $I J(K-1)$ | $S S E$ | $M S E$ |  |
| Total | $I J K-1$ | $S S T$ |  |  |

- degrees of freedom for curing time is $I-1=2$
- degrees of freedom for mixture type is $J-1=3$
- degrees of freedom for the curing time and mixture type interaction is $(I-1)(J-1)=6$
- degrees of freedom for error is $I J(K-1)=24$
- total degrees of freedom is $I J K-1=35$
- $S S A=30763.0$ (given to us)
- $S S B=34185.6$ (given to us)
- $S S E=97436.8$ (given to us)
- $S S T=205966.6$ (given to us)
- $S S A B=S S T-S S A-S S B-S S E=43581.2$
- $M S A=S S A /(I-1)=15381.5$
- $M S B=S S B /(J-1)=11395.2$
- $M S A B=S S A B /((I-1)(J-1))=7263.533$
- $M S E=S S E /(I J(K-1))=4059.867$
- $F_{A}=M S A / M S E \approx 3.79$
- $F_{B}=M S B / M S E \approx 2.81$
- $F_{A B}=M S A B / M S E \approx 1.79$


## Part B /2

Hypothesis Test: $H_{0 A B}: \gamma_{i j}=0$ for all $i, j$ vs. $H_{a A}$ : At least one $\gamma_{i j} \neq 0$

## Test Statistic:

$$
F_{A B}=\frac{M S A B}{M S E} \sim F(6,24) .
$$

From Part A, we have $F_{A B, o b s} \approx 1.79$

Method 1 - p-value: $P_{H_{0}}\left(F>F_{A B, o b s}\right)=0.1440$, where $F \sim F(6,24)$.
Since $0.1440>0.05$, we fail to reject $H_{0 A B}$.

## Method 2 - Rejection Region:

$$
\begin{aligned}
\mathcal{R}_{A B, 0.05} & =\left\{f: f>f_{0.05}(6,24)\right\} \\
& =\{f: f>2.51\} .
\end{aligned}
$$

Since $F_{A B, \text { obs }} \notin \mathcal{R}_{A B, 0.05}$, we fail to reject $H_{0 A B}$.

## Part C /2

Hypothesis Test: $H_{0 A}: \alpha_{1}=\alpha_{2}=\alpha_{3}=0$ vs. $H_{a A}$ : At least one $\alpha_{i} \neq 0$

## Test Statistic:

$$
F_{A}=\frac{M S A}{M S E} \sim F(2,24) .
$$

From Part A, we have $F_{A, \text { obs }} \approx 3.79$

Method 1 - p-value: $P_{H_{0}}\left(F>F_{A, o b s}\right)=0.0372$, where $F \sim F(2,24)$.
Since $0.0372<0.05$, we reject $H_{0 A}$.

Method 2 - Rejection Region:

$$
\begin{aligned}
\mathcal{R}_{A, 0.05} & =\left\{f: f>f_{0.05}(2,24)\right\} \\
& =\{f: f>3.40\}
\end{aligned}
$$

Since $F_{A, o b s} \in \mathcal{R}_{A, 0.05}$, we reject $H_{0 A}$.

## Part D /2

Hypothesis Test: $H_{0 B}: \beta_{1}=\beta_{2}=\beta_{3}=\beta_{4}=0$ vs. $H_{a B}$ : At least one $\beta_{j} \neq 0$

Test Statistic:

$$
F_{B}=\frac{M S B}{M S E} \sim F(3,24) .
$$

From Part A, we have $F_{B, \text { obs }} \approx 2.81$

Method 1 - p-value: $P_{H_{0}}\left(F>F_{B, o b s}\right)=0.0612$, where $F \sim F(3,24)$.
Since $0.0612>0.05$, we fail to reject $H_{0 B}$.

## Method 2 - Rejection Region:

$$
\begin{aligned}
\mathcal{R}_{B, 0.05} & =\left\{f: f>f_{0.05}(3,24)\right\} \\
& =\{f: f>3.01\}
\end{aligned}
$$

Since $F_{B, o b s} \notin \mathcal{R}_{B, 0.05}$, we fail to reject $H_{0 B}$.

## Part E /2

Note that we can apply Tukey's method, since we failed to reject $H_{0 A B}$ in Part $\mathbf{B}$, and rejected $H_{0 A}$ in Part C.

We start by computing

$$
W=Q_{0.05, I, I J(K-1)} \sqrt{\frac{M S E}{J K}}=3.53 \sqrt{\frac{4059.867}{12}}=64.9292
$$

Table 3 presents the sample mean differences between the curing times. We summarize our findings with the following underscoring pattern

$$
\begin{array}{cccc}
k & 3 & 1 & 2 \\
\bar{X}_{k . .} & 3960.02 & 4010.88 & 4029.10 \\
\hline
\end{array}
$$

Although there is not a significant difference between curing times 1 and 3 , we can see a significant difference between curing times 3 and 2 .

Table 3: $\bar{X}_{l . .}-\bar{X}_{i . .}$ for $l>i$. The bold-faced elements correspond to the values less than $W$.

| $i \backslash l$ | 3 | 1 | 2 |
| :--- | :--- | :--- | :--- |
| 3 |  | $\mathbf{5 0 . 8 6}$ | 69.08 |
| 1 |  |  | $\mathbf{1 8 . 2 2}$ |
| 2 |  |  |  |

## Section 11.2 Question 22 /8

Study Objective: Determine if there exists a difference between the writing lifetimes of four premium brands of pens. However, it is believed that the writing surface might affect the writing lifetime.

Formulation: Let

- $X_{i j k}$ denote the $k$ th observation for the $i$ th brand type and $j$ th writing surface, with $i=1,2,3,4, j=1,2,3$, and $k=1,2$ (ie $I=4, J=3$, and $K=2$ ).

We assume that $X_{i j k} \sim N\left(\mu_{i j}, \sigma^{2}+\sigma_{B}^{2}+\sigma_{G}^{2}\right)$, where $\mu_{i j}=\mu+\alpha_{i}+B_{j}+G_{i j}$, with $\sum_{i=1}^{3} \alpha_{i}=0$, $B_{j} \stackrel{i i d}{\sim} N\left(0, \sigma_{B}^{2}\right)$, and $G_{i j} \stackrel{i i d}{\sim} N\left(0, \sigma_{G}^{2}\right)$.

## Hypothesis Test:

$$
\begin{aligned}
& H_{0 A}: \alpha_{1}=\alpha_{2}=\alpha_{3}=\alpha_{4}=0 \text { vs. } H_{a A}: \text { At least one } \alpha_{i} \neq 0 \\
& H_{0 B}: \sigma_{B}^{2}=0 \text { vs. } H_{a B}: \sigma_{B}^{2}>0 \\
& H_{0 G}: \sigma_{G}^{2}=0 \text { vs. } H_{a G}: \sigma_{G}^{2}>0
\end{aligned}
$$

Note: It is customary to test $H_{0 A}$ and $H_{0 B}$ if we fail to reject $H_{0 G}$. To conduct the hyptohesis tests, let's fill out the following ANOVA table:

| Source | df | Sum of Squares | Mean Square | $F$ |
| :--- | :--- | :--- | :--- | :--- |
| Pen Brand (Factor A) | $I-1$ | $S S A$ | $M S A$ | $F_{A}$ |
| Writing Surface (Factor B) | $J-1$ | $S S B$ | $M S B$ | $F_{B}$ |
| Interaction | $(I-1)(J-1)$ | $S S A B$ | $M S A B$ | $F_{A B}$ |
| Error | $I J(K-1)$ | $S S E$ | $M S E$ |  |
| Total | $I J K-1$ | $S S T$ |  |  |

- degrees of freedom for pen brand is $I-1=3$
- degrees of freedom for writing surface is $J-1=2$
- degrees of freedom for the pen brand and writing surface interaction is $(I-1)(J-1)=6$
- degrees of freedom for error is $I J(K-1)=12$
- total degrees of freedom is $I J K-1=23$
- $S S T=\sum_{i=1}^{I} \sum_{j=1}^{J} \sum_{k=1}^{K}\left(X_{i j k}-\bar{X}_{\ldots}\right)^{2}=20591.83$
- $S S E=\sum_{i=1}^{I} \sum_{j=1}^{J} \sum_{k=1}^{K}\left(X_{i j k}-\bar{X}_{i j} .\right)^{2}=8216.0$
- $S S A=J K \sum_{i=1}^{I}\left(\bar{X}_{i . .}-\bar{X}_{\ldots .}\right)^{2}=1387.5$
- $S S B=I K \sum_{j=1}^{J}\left(\bar{X}_{. j .}-\bar{X}_{. .}\right)^{2}=2888.083$
- $S S A B=S S T-S S A-S S B-S S E=8100.25$
- $M S A=S S A /(I-1)=462.5$
- $M S B=S S B /(J-1)=1444.042$
- $M S A B=S S A B /((I-1)(J-1))=1350.042$
- $M S E=S S E /(I J(K-1))=684.6667$
- $F_{A}=M S A / M S A B \approx 0.34$
- $F_{B}=M S B / M S A B \approx 1.07$
- $F_{A B}=M S A B / M S E \approx 1.97$


## Test Statistics:

$$
\begin{aligned}
F_{A} & =\frac{M S A}{M S A B} \sim F(I-1,(I-1)(J-1)) \\
F_{B} & =\frac{M S B}{M S A B} \sim F(J-1,(I-1)(J-1)) \\
F_{A B} & =\frac{M S A B}{M S E} \sim F((I-1)(J-1), I J(K-1))
\end{aligned}
$$

## Decision:

$P_{H_{0 G}}\left(F_{A B}>F_{A B, o b s}\right)=0.1492$, where $F_{A B} \sim F(6,12)$. Since $0.05<0.1492$, we fail to reject $H_{0 G}$.
$P_{H_{0 A}}\left(F_{A}>F_{A, o b s}\right)=0.7960$, where $F_{A} \sim F(3,6)$ since $0.05<0.7960$, we fail to reject $H_{0 A}$.
$P_{H_{0 B}}\left(F_{B}>F_{B, o b s}\right)=0.4006$, where $F_{A} \sim F(2,6)$ since $0.05<0.4006$, we fail to reject $H_{0 B}$.

