

What to do today (Oct 28, 2020)?

1. *Introduction*

2. *Probability and Distribution (Chp 1-3)*

3. Essential Topics in Mathematical Statistics

3.1 Elementary Statistical Inferences (Chp 4)

▶ *3.1.1 Sampling and Statistics*

▶ *3.1.2 Confidence Interval*

▶ *3.1.3 Order Statistics*

▶ *3.1.4 Hypothesis Testing*

▶ **3.1.5 Statistical Simulation and Bootstrap**

3.2 Consistency and Limiting Distributions (Chp 5)

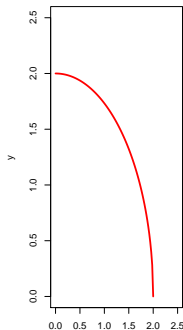
3.3 Maximum Likelihood Methods (Chp 6)

3.1.5 Statistical Simulation and Bootstrap: Monte Carlo Integration

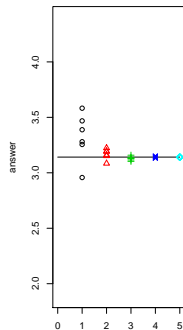
How to calculate $\int_a^b g(x)dx$? For example,

$$\int_0^2 \sqrt{4-x^2} dx = 2 \int_0^2 \sqrt{4-x^2} \left(\frac{1}{2}\right) dx = 2E\{\sqrt{4-X^2}\} \quad (\pi)$$

provided that $X \sim U(0, 2)$.



(a) $y=(4-x^2)^{.5}$



(b) true answer= π

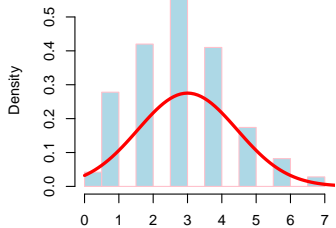
(b) Approximates to the integral by generating n observations from $U(0, 2)$, with $n = 10^k$ for $k = 1, \dots, 5$.

3.1.5 Statistical Simulation and Bootstrap: Simulation Example 1

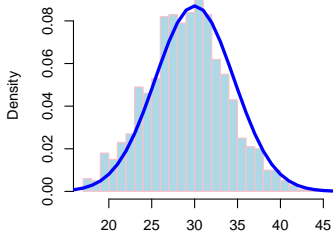
To verify the normal approximation to binomial distn:

```
xtmp=matrix( ifelse( runif(1000*10)<.3,1,0), ncol=10)
x=apply(xtmp,1,sum)
hist(x, freq=FALSE, ...)
y=rbinom(n=1000, size=10, prob=0.3)
hist(y, freq=FALSE, breaks=11, ...)
```

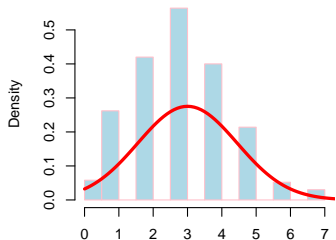
```
xtmp=matrix( ifelse( runif(1000*100)<.3,1,0), ncol=100)
x=apply(xtmp,1,sum)
hist(x, freq=FALSE, breaks=20, ...)
curve(exp(-(x-30)^2/2/(30*.7))/(2*pi*21)^.5, 0, 100, lty=1, col=4,
      lwd=3, add=TRUE)
y=rbinom(n=1000, size=100, prob=0.3)
hist(y, freq=FALSE, breaks=20, ...)
curve(exp(-(x-30)^2/2/(30*.7))/(2*pi*21)^.5, 0, 100, lty=1, col=4,
      lwd=3, add=TRUE)
```



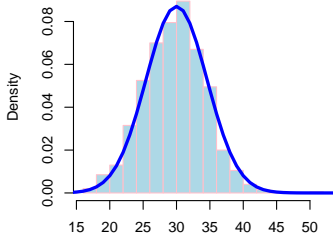
(a1). $X \sim B(10, 0.3)$



(a2). $X \sim B(100, 0.3)$



(b1). $Y \sim B(10, 0.3)$



(b2). $Y \sim B(100, 0.3)$

3.1.5 Statistical Simulation and Bootstrap: Simulation Example 2

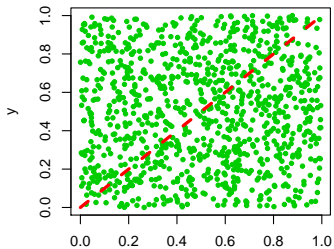
What can data missing result in?

```
x=runif(1000);  
y=runif(1000);  
w=ifelse(x<y,1,0); sum(w)/1000
```

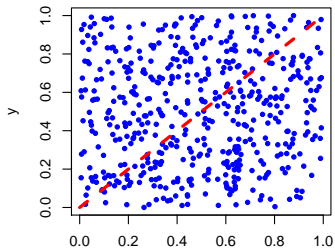
```
r0=ifelse(runif(1000)<.5,1,0)  
x0=x[r0==1]; y0=y[r0==1];  
w0=ifelse(x0<y0,1,0); sum(w0)/sum(r0);
```

```
r1=rep(0,1000)  
r1[x<y]=rbinom(length(x[x<y]), size=1, prob=.8)  
r1[x>=y]=rbinom(length(x[x>=y]), size=1, prob=.2)  
x1=x[r1==1]; y1=y[r1==1];  
w1=ifelse(x1<y1,1,0); sum(w1)/sum(r1);
```

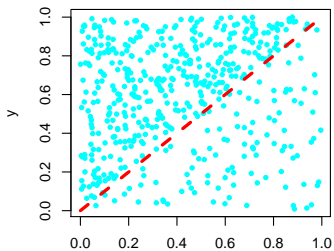
```
r2=rep(0,1000)  
r2[x<y]=rbinom(length(x[x<y]), size=1, prob=.3)  
r2[x>=y]=rbinom(length(x[x>=y]), size=1, prob=.7)  
x2=x[r2==1]; y2=y[r2==1];  
w2=ifelse(x2<y2,1,0); sum(w2)/sum(r2);
```



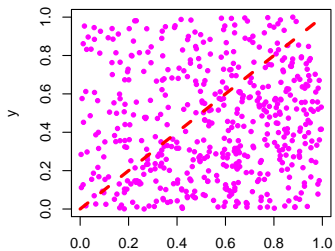
x
(a).all obs: rate of ' $x < y$ '=.495



x
(b).random missing 50%: rate of ' $x < y$ '=.486



x
(c).nonrandom missing 50%: rate of ' $x < y$ '=.789



x
(d).nonrandom missing 50%: rate of ' $x < y$ '=.337

3.1.5 Statistical Simulation and Bootstrap: Preparation for Bootstrap

Consider rv $X \sim F(\cdot)$: iid observations X_1, \dots, X_n

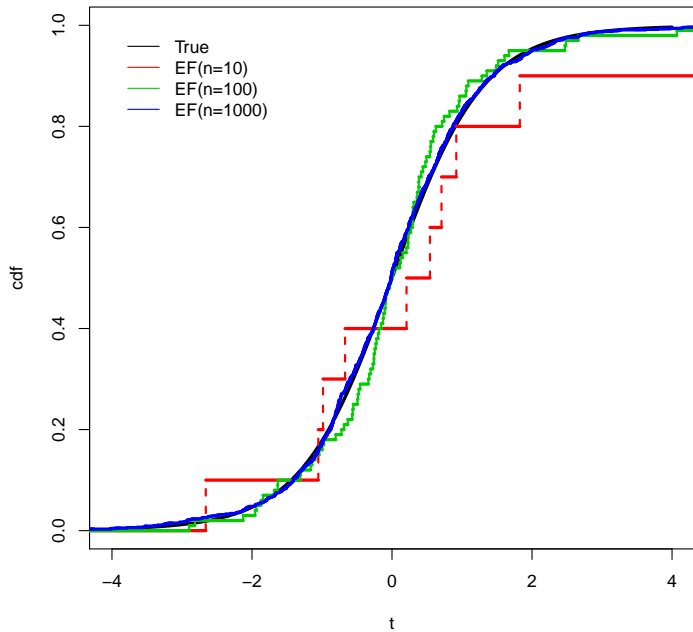
Definition. The following is the **empirical function** with the random sample:

$$F_n(x) = \frac{1}{n} \sum_{i=1}^n I(X_i \leq x) \quad \text{for } -\infty < x < \infty.$$

For all $x \in (-\infty, \infty)$,

- ▶ $E\{F_n(x)\} = F(x)$ and $\text{Var}\{F_n(x)\} = F(x)[1 - F(x)]/n$.
- ▶ more

$F_n(\cdot)$ is a very good estimator of $F(\cdot)$.



$$T \sim t(6).$$

3.1.5 Statistical Simulation and Bootstrap:

Bootstrap

Consider rv $X \sim F(\cdot)$: iid observations X_1, \dots, X_n

- ▶ When to use a point estimator $\hat{\theta}(X_1, \dots, X_n)$ of a population parameter θ , how to estimate its variance $\text{Var}(\hat{\theta})$?

Bootstrap variance estimation:

- ▶ Step 1. Generate $X_{1b}^*, \dots, X_{nb}^*$ iid from the empirical function $F_n(\cdot)$.
(Resample with size n from X_1, \dots, X_n with replacement.)
- ▶ Step 2. Calculate $\hat{\theta}(X_{1b}^*, \dots, X_{nb}^*)$, denoted by $\hat{\theta}_b^*$.
- ▶ Repeat Steps 1 and 2 B times and obtain $\{\hat{\theta}_b^* : b = 1, \dots, B\}$.
- ▶ Calculate

$$s_{\hat{\theta}^*}^2 = \sum_{b=1}^B (\hat{\theta}_b^* - \bar{\hat{\theta}^*})^2 / (B - 1)$$

with $\bar{\hat{\theta}^*} = \sum_{b=1}^B \hat{\theta}_b^* / B$.

- ▶ Use $s_{\hat{\theta}^*}^2$ to estimate $\text{Var}(\hat{\theta})$.

3.1.5 Statistical Simulation and Bootstrap: Bootstrap

Consider rv $X \sim F(\cdot)$: iid observations X_1, \dots, X_n

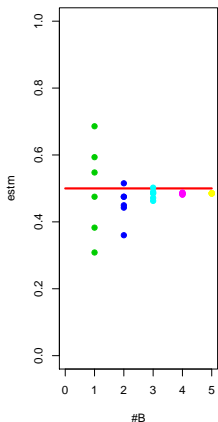
- ▶ How to obtain an interval estimator of a population parameter θ based on a point estimator $\hat{\theta}(X_1, \dots, X_n)$?

Bootstrap confidence interval:

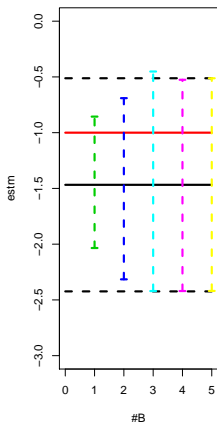
- ▶ Step 1. Generate $X_{1b}^*, \dots, X_{nb}^*$ iid from the empirical function $F_n(\cdot)$.
(Resample with size n from X_1, \dots, X_n with replacement.)
- ▶ Step 2. Calculate $\hat{\theta}(X_{1b}^*, \dots, X_{nb}^*)$, denoted by $\hat{\theta}_b^*$.
- ▶ Repeat Steps 1 and 2 B times and obtain $\{\hat{\theta}_b^* : b = 1, \dots, B\}$.
- ▶ Sort the sequence as $\hat{\theta}_{(1)}^* \leq \dots \leq \hat{\theta}_{(B)}^*$, and obtain bootstrap percentiles: $\hat{\theta}_{((\alpha/2)100)}^*$ and $\hat{\theta}_{((1-\alpha/2)100)}^*$.
- ▶ Use $(\hat{\theta}_{((\alpha/2)100)}^*, \hat{\theta}_{((1-\alpha/2)100)}^*)$ as a $(1 - \alpha)100\%$ CI for θ .

Bootstrap example

Consider $X \sim F(\cdot)$ with $\mu = E(X)$ and iid obs X_1, \dots, X_n .



(a) Bootstrap Estm for SD of Xbar



(b) Bootstrap CI Estm for Population Mean

$X \sim N(-1, 5^2)$ with $n = 100$ and $B = 10^k$ for $k = 1, \dots, 5$.

What will we study next class?

1. *Introduction*

2. *Probability and Distribution (Chp 1-3)*

3. Essential Topics in Mathematical Statistics (Chp 4-6)

- ▶ *3.1 Elementary Statistical Inferences (Chp 4)*
- ▶ **3.2 Consistency and Limiting Distributions (Chp 5)**
 - ▶ **3.2.1 Convergence in Probability**
 - ▶ **3.2.2 Convergence in Distribution**
- ▶ *3.3 Maximum Likelihood Methods (Chp 6)*

4. *Further Topics, Selected from Chp 7-11*