## What to do today (Nov 2, 2020)?

- 1. Introduction
- 2. Probability and Distribution (Chp 1-3)
- 3. Essential Topics in Mathematical Statistics
- 3.1 Elementary Statistical Inferences (Chp 4)
  - ▶ 3.1.1 Sampling and Statistics
  - 3.1.2 Confidence Interval
  - 3.1.3 Order Statistics
  - 3.1.4 Hypothesis Testing
  - 3.1.5 Statistical Simulation and Bootstrap
- 3.2 Consistency and Limiting Distributions (Chp 5)
  - ► 3.2.1 Convergence in Probability
  - ► 3.2.2 Convergence in Distribution
- 3.3 Maximum Likelihood Methods (Chp 6)

## 3.1.5 Statistical Simulation and Bootstrap: Preparation for Bootstrap

Consider rv  $X \sim F(\cdot)$ : iid observations  $X_1, \ldots, X_n$ 

**Definition.** The following is the **empirical function** with the random sample:

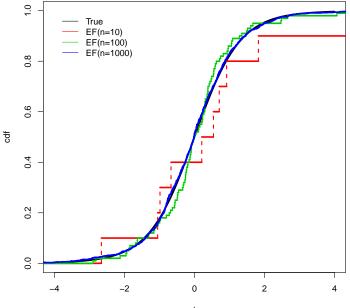
$$F_n(x) = rac{1}{n} \sum_{i=1}^n I(X_i \le x) ext{ for } -\infty < x < \infty.$$

For all  $x \in (-\infty, \infty)$ ,

•  $E\{F_n(x)\} = F(x) \text{ and } Var\{F_n(x)\} = F(x)[1-F(x)]/n.$ 

more ... ...

 $F_n(\cdot)$  is a very good estimator of  $F(\cdot)$ .



 $T \sim t(6)$ .

t

# **3.1.5 Statistical Simulation and Bootstrap: Bootstrap**

Consider rv  $X \sim F(\cdot)$ : iid observations  $X_1, \ldots, X_n$ 

When to use a point estmator θ(X<sub>1</sub>,...,X<sub>n</sub>) of a population parameter θ, how to estimate its variance Var(θ)?

Thinking ... ...

 If we could have a random sample θ
<sub>b</sub> for b = 1,..., B from the same population as θ
<sub>i</sub>, we can estimate the variance with

$$s_{\widehat{ heta}}^2 = \sum_{b=1}^B (\widehat{ heta}_b - \overline{\widehat{ heta}})^2 / (B-1)$$

with  $\overline{\hat{\theta}} = \sum_{b=1}^{B} \widehat{\theta}_b / B$ .

► That can be achieved if there are X<sub>1b</sub>,..., X<sub>nb</sub> iid from F(·). However, F(·) is unknown. How to overcome it?

## 3.1.5 Statistical Simulation and Bootstrap:

### Bootstrap

Consider rv  $X \sim F(\cdot)$ : iid observations  $X_1, \ldots, X_n$ 

When to use a point estmator θ̂(X<sub>1</sub>,...,X<sub>n</sub>) of a population parameter θ, how to estimate its variance Var(θ̂)?

#### Bootstrap variance estimation:

Step 1. Generate  $X_{1b}^*, \ldots, X_{nb}^*$  iid from the empirical function  $F_n(\cdot)$ .

(Resample with size n from  $X_1, \ldots, X_n$  with replacement.)

- Step 2. Calculate  $\widehat{\theta}(X_{1b}^*, \dots, X_{nb}^*)$ , denoted by  $\widehat{\theta}_b^*$ .
- ▶ Repeat Steps 1. and 2. *B* times and obtain  $\{\widehat{\theta}_b^* : b = 1, \dots, B\}.$
- ► With  $\overline{\hat{\theta}^*} = \sum_{b=1}^{B} \widehat{\theta}_b^* / B$ , calculate  $s_{\widehat{\theta}^*}^2 = \sum_{b=1}^{B} (\widehat{\theta}_b^* - \overline{\widehat{\theta}^*})^2 / (B - 1)$ .

• Use 
$$s_{\hat{\theta}^*}^2$$
 to estimate  $Var(\hat{\theta})$ .

Resampling methods: Jackknife (J.W. Tukey, 1958); Bootstrap (Bradley Efron, 1979)

# **3.1.5 Statistical Simulation and Bootstrap: Bootstrap**

Consider rv  $X \sim F(\cdot)$ : iid observations  $X_1, \ldots, X_n$ 

• How to obtain an interval estimator of a population parameter  $\theta$  based on a point estmator  $\hat{\theta}(X_1, \ldots, X_n)$ ?

#### Bootstrap confidence interval:

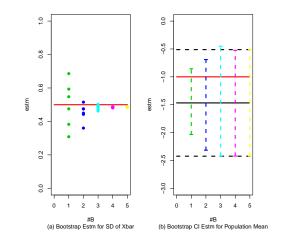
Step 1. Generate  $X_{1b}^*, \ldots, X_{nb}^*$  iid from the empirical function  $F_n(\cdot)$ .

(Resample with size n from  $X_1, \ldots, X_n$  with replacement.)

- ▶ Step 2. Calculate  $\hat{\theta}(X_{1b}^*, \dots, X_{nb}^*)$ , denoted by  $\hat{\theta}_b^*$ .
- Repeat Steps 1. and 2. *B* times and obtain  $\{\widehat{\theta}_b^*: b = 1, \dots, B\}.$
- Sort the sequence as θ<sup>\*</sup><sub>(1)</sub> ≤ ... ≤ θ<sup>\*</sup><sub>(B)</sub>, and obtain bootstrap percentiles: θ<sup>\*</sup><sub>((α/2)100)</sub> and θ<sup>\*</sup><sub>((1-α/2)100)</sub>.
- Use  $(\widehat{\theta}^*_{((\alpha/2)100)}, \widehat{\theta}^*_{((1-\alpha/2)100)})$  as a  $(1-\alpha)100\%$  CI for  $\theta$ .

#### **Bootstrap** example

Consider  $X \sim F(\cdot)$  with  $\mu = E(X)$  and iid obs  $X_1, \ldots, X_n$ .



 $X \sim N(-1, 5^2)$  with n = 100 and  $B = 10^k$  for  $k = 1, \dots, 5$ .

## 3.2.1 Convergence in Probability

**Definition.** We say a sequence of random variables (rvs)  $\{Y_n : n = 1, 2, ...\}$  converges in probability to rv Y if, for any  $(\forall) \ \epsilon > 0$ ,

$$\lim_{n\to\infty} P(|Y_n-Y|\geq\epsilon)=0.$$

Denote it by " $Y_n \to Y$  in probability" as  $n \to \infty$ , or " $Y_n \xrightarrow{P} Y$  as  $n \to \infty$ ".

• A special case is that  $Y_n \xrightarrow{P} c$ , a constant.

**Theorem.** (Weak Law of Large Numbers (WLLN)). Let  $\{X_n\}$  be a sequence of iid rvs with the common mean  $\mu$  and variance  $\sigma^2 < \infty$ . Let  $Y_n = \bar{X}_n = \sum_{i=1}^n X_n/n$ . Then the sequence  $\{Y_n\}$  converges in probability to  $\mu$ . That is,  $\bar{X}_n \xrightarrow{P} \mu$  as  $n \to \infty$ .

**Definition.** Let  $X_1, \ldots, X_n$  be a sample from  $F(x; \theta), \theta \in \Omega$ . A statistic of the sample, denoted by  $T_n$ , is a **consistent** estimator of  $\theta$  if  $T_n \xrightarrow{P} \theta$  as  $n \to \infty$ .

Comments:

- ▶ By WLLN,  $\bar{X}_n \xrightarrow{P} \mu = E(X)$  as  $n \to \infty$  if the observations on X are iid. That is,  $\bar{X}_n$  is a consistent estimator of  $\mu$ .
- The sequence of sample proportions converges in probability to the population proportion. (*View it as the theoretical support to the frequentist definition for probability.*)

**Theorem.** Suppose  $X_n \xrightarrow{P} X$  and  $Y_n \xrightarrow{P} Y$  as  $n \to \infty$ . The following results hold:

- $X_n + Y_n \xrightarrow{P} X + Y$ , and  $X_n Y_n \xrightarrow{P} XY$ .
- $g(X_n) \xrightarrow{P} g(X)$  for any continuous function  $g(\cdot)$ .

eg, the sample variance  $\left[\sum_{i=1}^{n} X_i^2 - n\bar{X}^2\right]/(n-1)$  is a consistent estimator of  $\sigma^2 = Var(X)$  if  $X_1, \ldots, X_n$  is a random sample from the population.

### 3.2.2 Convergence in Distribution

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**Definition.** Consider rv  $X \sim F_X(\cdot)$ . We call a sequence of rvs  $\{X_n\}$  converges in distribution to X if

$$\lim_{n\to\infty}F_{X_n}(x)=\lim_{n\to\infty}P(X_n\leq x)=F(x)$$

for all  $x \in C(F_X)$ , the set of all continuous points of  $F_X(.)$ . Denote it by  $X_n \xrightarrow{D} X$ . In other words, the **limiting distribution** or the **asymptotic distribution** of  $\{X_n\}$  is  $F_X(\cdot)$ .

Recall the Central Limit Theorem (CLT): **Theorem.** (CLT) If  $X_1, \ldots, X_n$  are iid with mean  $\mu$  and variance  $\sigma^2$ ,

$$Y_n = rac{1}{\sqrt{n}} \sum_{i=1}^n \left(rac{X_i - \mu}{\sigma}
ight) \stackrel{D}{
ightarrow} N(0, 1)n 
ightarrow \infty.$$

- ► That is,  $(\bar{X} \mu) / \sqrt{\sigma^2 / n} \xrightarrow{D} \to N(0, 1)$  as  $n \to \infty$ .
- That is,  $\sum_{i=1}^{n} X_i \sim N(n\mu, n\sigma^2)$  approximately as n >> 1.

## 3.2.2 Convergence in Distribution

**Theorem.** Consider a sequence of rvs  $\{X_n\}$ .

• If 
$$X_n \xrightarrow{P} X$$
, then  $X_n \xrightarrow{D} X$ .

• If  $X_n \xrightarrow{D} c$  with c a constant, then  $X_n \xrightarrow{P} c$ .

**Theorem.** Consider sequences of rvs  $\{X_n\}$  with  $X_n \xrightarrow{D} X$ .

- If  $g(\cdot)$  is a continuous function on the support of X, then  $g(X_n) \xrightarrow{D} g(X)$ .
- ▶ If rvs  $A_n \xrightarrow{P} a$  and rvs  $B_n \xrightarrow{P} b$  with both *a* and *b* constant, then

$$A_n + B_n X_n \stackrel{D}{\to} a + b X.$$

(Slutsky's Theorem)

**Example 3.6** Suppose  $X_1, \ldots, X_n$  is a random sample from the uniform distn  $U(0, \theta)$ .

- $Y_n = \max(X_1, \ldots, X_n)$  can be a "good estimator" of  $\theta$ .
- ► *Y<sub>n</sub>*'s distn:

$$F_{Y_n}(y) = \begin{cases} 1, & y > \theta \\ (y/\theta)^n, & 0 < y \le \theta \\ 0, & t \le 0; \end{cases} \quad f_{Y_n}(y) = \begin{cases} ny^{n-1}/\theta^n, & 0 < y \le \theta \\ 0, & elsewhere. \end{cases}$$

• 
$$E(Y_n) = n\theta/(n+1)$$
, a biased estimator of  $\theta$ .

► 
$$F_{Y_n}(y) \to 1$$
 or 0 for  $y \ge \theta$  or  $y < \theta$ , respectively. Thus  $Y_n \xrightarrow{D} \theta$ .

- Since  $\theta$  is a constant,  $Y_n \xrightarrow{P} \theta$ :  $Y_n$  is consistent.
- Further, let  $W_n = n(\theta Y_n)$ . The distn of  $W_n$

$$P(W_n \leq t) = P(Y_n \geq \theta - t/n) = 1 - \left(1 - \frac{t/\theta}{n}\right)^n$$

converges to  $1 - \exp(-t/\theta)$ . That is  $W_n \xrightarrow{D} W$ , which follows the exponential distn  $NE(\theta)$ .

### What will we study next class?

- 1. Introduction
- 2. Probability and Distribution (Chp 1-3)

#### 3. Essential Topics in Mathematical Statistics (Chp 4-6)

- ► 3.1 Elementary Statistical Inferences (Chp 4)
- 3.2 Consistency and Limiting Distributions (Chp 5)
  - ► 3.2.1 Convergence in Probability
  - ► 3.2.2 Convergence in Distribution
- 3.3 Maximum Likelihood Methods (Chp 6)
- 4. Further Topics, Selected from Chp 7-11