## What to do today (Nov 18, 2020)?

- 1. Introduction
- 2. Probability and Distribution (Chp 1-3)

#### 3. Essential Topics in Mathematical Statistics

- 3.1 Elementary Statistical Inferences (Chp 4)
- 3.2 Consistency and Limiting Distributions (Chp 5)
- 3.3 Maximum Likelihood Methods (Chp 6)
  - ▶ 3.3.1 Maximum Likelihood Estimation
  - ▶ 3.3.2 Likelihood-Based Tests
  - 3.3.3 EM Algorithm

#### 4. Further Topics, Selected from Chp 7-11

# 3.3.3 Expectation-Maximization (EM) Algorithm

#### What does it do?

An iterative procedure (algorithm) to calculate MLEs of the population parameters  $\theta$  when it is hard to maximize the likelihood function of  $\theta$  with the available data.

Original References:

- Dempster, Laird, and Rubin (1977). "Maximum Likelihood from Incomplete Data via the EM Algorithm". Journal of the Royal Statistical Society, Series B. 39 (1): 1–38.
- C.F. Jeff Wu (1983). "On the Convergence Properties of the EM Algorithm". Annals of Statistics. 11 (1): 95–103.
- B.W. Turbull (1976) "The Empirical Distribution Function with Arbitrarily Grouped, Censored and Truncated Data". Journal of the Royal Statistical Society, Series B. 38 (3): 290–295.

**3.3.3 Expectation-Maximization (EM) Algorithm Goal**: Consider to maximize  $L(\theta|\mathbf{X})$ , the (observed) likelihood function with the available data  $\mathbf{X}$  to obtain the MLE of population parameter  $\theta$ .

Suppose that it is relatively easier to maximize  $L_C(\theta | \mathbf{X}, \mathbf{Z})$ , the (complete) likelihood function of  $\theta$  with the "augmented" data  $(\mathbf{X}, \mathbf{Z})$ .

Define  $Q(\theta|\theta_0, \mathbf{x}) = E_{\theta_0} [\log L_C(\theta|\mathbf{X}, \mathbf{Z})|\theta_0, \mathbf{X} = \mathbf{x}].$ 

**EM Algorithm.** Let  $\hat{\theta}^{(m)}$  be the estimate on the *m*th step with  $m \ge 0$ . The following compute the estimate on the (m+1)th step:

Expectation-Step. Compute

$$Q(\boldsymbol{\theta}|\hat{\boldsymbol{\theta}}^{(m)},\mathbf{x}) = E_{\hat{\boldsymbol{\theta}}^{(m)}} \big[ \log L_C(\boldsymbol{\theta}|\mathbf{X},\mathbf{Z}) | \hat{\boldsymbol{\theta}}^{(m)},\mathbf{X} = \mathbf{x} \big].$$

Maximization-Step. Let the updated estimate be

$${\hat{oldsymbol{ heta}}}^{(m+1)} = {\it argmax}_{oldsymbol{ heta} \in \Omega} Q(oldsymbol{ heta}| {\hat{oldsymbol{ heta}}}^{(m)}, {f x}).$$

## 3.3.3 Expectation-Maximization (EM) Algorithm

- Starting with an initial estimate (guess) of θ, say, θ<sup>(0)</sup>, repeat the *E-Step* and *M-Step* and record the sequence of estimates {θ<sup>(m)</sup> : m = 1, 2, ...}.
- ▶ Under some assumptions, the sequence converges in probability to the MLE  $\hat{\theta}$  as  $m \to \infty$ .
- ► The EM algorithm works to improve Q(θ|θ̂<sup>(m)</sup>, x). This implies improvements to L(θ|x): L(θ̂<sup>(m+1)</sup>|x) ≥ L(θ̂<sup>(m)</sup>|x).

EM algorithm has a broad range of applications!

... and a lot of variants, eg MCEM Algorithm, ES Algorithm, ...

Example 3.14 Consider a mixture of normal distributions:  $X = (1 - W)Y_1 + WY_2$  with  $Y_j \sim N(\mu_j, \sigma_j^2)$  for j = 1, 2 and  $W \sim B(1, \epsilon)$ . Suppose the observations on a random sample  $\mathbf{X}' = (X_1, \dots, X_n)$  from the mixture distn are available.

- Apply the EM algorithm to estimate the parameter θ' = (μ<sub>1</sub>, σ<sub>1</sub><sup>2</sup>, μ<sub>2</sub>, σ<sub>2</sub><sup>2</sup>, ε)
- Similar to Ex6.6.8, generate data with n = 1000 and  $\theta^{'} = (100, 15^2, 200, 10^2, .8)$ , and then evaluate the MLE of  $\theta^{'}$  by EM algorithm.

Solution. Let  $f_j(y)$  be the pdf of  $Y_j$ . The log-likelihood functn with the observed data is

$$l(\boldsymbol{ heta}|\mathbf{x}) = \sum_{i=1}^{n} \log \left[ (1-\epsilon)f_1(x_i) + \epsilon f_2(x_i) \right].$$

With observations  $\mathbf{w}'$  on a random sample  $W_1, \ldots, W_n$  from  $W \sim B(1, \epsilon)$ , the log-likelihood functn with the complete data is

$$I_C(oldsymbol{ heta}|\mathbf{x},\mathbf{w}) = \sum_{i=1}^n \Big\{ \big[ (1-w_i)\log f_1(x_i) + w_i\log f_2(x_i) \big] + \big[ w_i\log \epsilon + (1-w_i)\log(1-\epsilon) \big] \Big\}.$$

Provided the estimate at the *m*th stage is  $\theta^{(m)}$ ,

*E-Step.* 
$$Q(\boldsymbol{\theta}|\boldsymbol{\theta}^{(m)}, \mathbf{x}) = I_{\mathcal{C}}(\boldsymbol{\theta}|\mathbf{x}, \mathbf{w})\Big|_{\mathbf{w}=\widehat{\mathbf{w}}^{(m)}}$$
 with  
 $\hat{w}_{i}^{(m)} = \mathcal{E}(W_{i}|\boldsymbol{\theta}^{(m)}, \mathbf{x}) = \frac{\epsilon^{(m)}f_{2}(x_{i})^{(m)}}{(1-\epsilon^{(m)})f_{1}(x_{i})^{(m)} + \epsilon^{(m)}f_{2}(x_{i})^{(m)}}.$ 

*M-Step.* Maximizing  $Q(\boldsymbol{\theta}|\boldsymbol{\theta}^{(m)}, \mathbf{x})$  wrt  $\boldsymbol{\theta}$  yields

$$\begin{split} \mu_1^{(m+1)} &= \frac{\sum_{i=1}^n (1 - \hat{w}_i^{(m)}) x_i}{\sum_{i=1}^n (1 - \hat{w}_i^{(m)})}, \quad \sigma_1^{2(m+1)} = \frac{\sum_{i=1}^n (1 - \hat{w}_i^{(m)}) (x_i - \mu_1^{(m+1)})^2}{\sum_{i=1}^n (1 - \hat{w}_i^{(m)})}, \\ \mu_2^{(m+1)} &= \frac{\sum_{i=1}^n \hat{w}_i^{(m)} x_i}{\sum_{i=1}^n \hat{w}_i^{(m)}}, \quad \sigma_2^{2(m+1)} = \frac{\sum_{i=1}^n \hat{w}_i^{(m)} (x_i - \mu_2^{(m+1)})^2}{\sum_{i=1}^n \hat{w}_i^{(m)}}, \\ \text{and } \epsilon^{(m+1)} &= \sum_{i=1}^n \hat{w}_i^{(m)} / n. \end{split}$$



### What will we do next week?

- 1. Introduction
- 2. Probability and Distribution (Chp 1-3)
- 3. Essential Topics in Mathematical Statistics (Chp 4-6)
- 4. Further Topics, Selected from Chp 7-11
  - 4.1 Nonparametric and Robust Statistics (Chp 10)
  - ▶ 4.2 Bayesian Procedures (Chp 11.1)