

STAT 330

Tutorial 7

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Outline

1. The Method of Monte Carlo

- Transformations of commonly used rvs
- Accept-Reject Algorithm

2. Bootstrap Procedures

The Method of Monte Carlo

4.8.5. Determine a method to generate random observations for the logistic pdf .

Review:

Monte Carlo methods are a class of computational algorithms that obtain numerical results by repeated random sampling.

How to generate random variables ?

Use transformations of commonly used rvs: for example,

If given a cdf $F(\cdot)$, $Y = F^{-1}(U)$ with $U \sim U(0,1)$ has

$Y \sim F(\cdot)$.

Answer :

The cdf of the logistic distribution is

$$F(x) = \frac{1}{1 + e^{-x}}, \quad -\infty < x < \infty.$$

To determine the inverse of this function,

set $u = \frac{1}{1 + e^{-x}}$ and then solve for x .

Take inverse of both sides, we have

$$u^{-1} = 1 + e^{-x}$$

Then we take logarithm of both sides and get

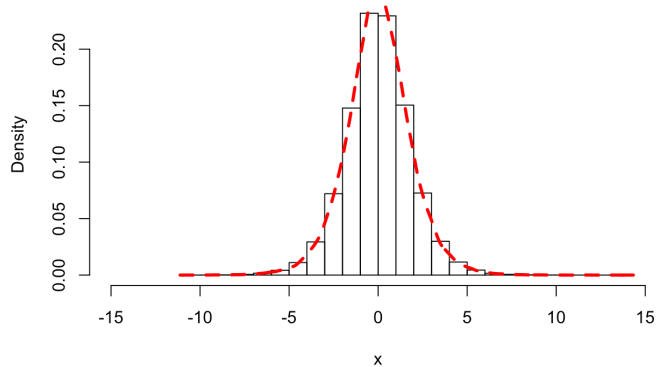
$$-x = \log(u^{-1} - 1).$$

Thus, $F^{-1}(u) = x = \log\left(\frac{u}{1-u}\right)$.

Hence, if U is uniform $(0,1)$ then $\log\frac{U}{1-U}$ has a logistic distribution with cdf $F(x)$.

The Method of Monte Carlo

Histogram of Random Observations using Generators in R



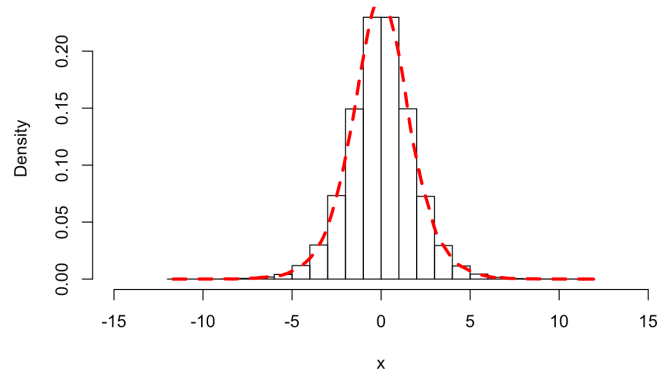
```
Fx = function(n){  
  U = runif(n)  
  Fx = log(U/(1-U))  
  Fx  
}
```

```
par(mfrow=c(1,2))
```

```
logistic<-rlogis(100000, location = 0, scale = 1)  
hist(logistic,freq=FALSE,main="Histogram of Random Observations using Generators in R",xlab="x",xlim=c(-15,15))  
lines(density(logistic),col=2,lty=2,lwd=3)
```

```
mclogistic<-Fx(100000)  
hist(mclogistic,freq=FALSE,main="Histogram of Random Observations using Transformations",xlab="x",xlim=c(-15,15))  
lines(density(mclogistic),col=2,lty=2,lwd=3)
```

Histogram of Random Observations using Transformations



The Method of Monte Carlo

4.8.20. For $\alpha > 0$ and $\beta > 0$, consider the following accept-reject algorithm:

1. Generate U_1 and U_2 iid uniform(0, 1) random variables. Set $V_1 = U_1^{1/\alpha}$ and $V_2 = U_2^{1/\beta}$.
2. Set $W = V_1 + V_2$. If $W \leq 1$, set $X = V_1/W$; else go to step 1.
3. Deliver X .

Show that X has a beta distribution with parameters α and β .

Review: Accept - Reject Algorithm

If $f(\cdot)$ is a pdf and $f(x) \leq M g(x)$ with M a constant and $g(\cdot)$ the instrumental pdf.

Step 1. Generate $Y \sim g(\cdot)$ and $U \sim U(0, 1)$ independently.

Step 2. If $U \leq \frac{f(Y)}{[Mg(Y)]}$, take $X=Y$ and go to Step 3;

otherwise, return to Step 1.

Step 3. Obtain X , which follows $f(\cdot)$.

The Method of Monte Carlo

Answer:

$$\begin{aligned} \text{We note that } F_X(x) &= P_c(X \leq x) = P_c\left(\frac{V_1}{W} \leq x \mid W \leq 1\right) \\ &= \frac{P_c(V_1 \leq xW, W \leq 1)}{P_c(W \leq 1)} \\ &\propto P_c(V_1 \leq xW, W \leq 1) \end{aligned}$$

First, we want to find the cdf of V_1 and V_2 .

We have $V_1 = U_1^{1/2}$ and $V_2 = U_2^{1/\beta}$ where U_1 and

U_2 are iid uniform $(0,1)$ random variables.

$$\begin{aligned} F_{V_1}(x) &= P_c(V_1 \leq x) = P_c(U_1^{1/2} \leq x) = P_c(U_1 \leq x^2) \\ &= \int_0^{x^2} 1 dx = x^2. \end{aligned}$$

$$\begin{aligned} F_{V_2}(x) &= P_c(V_2 \leq x) = P_c(U_2^{1/\beta} \leq x) = P_c(U_2 \leq x^\beta) \\ &= \int_0^{x^\beta} 1 dx = x^\beta. \end{aligned}$$

Take the derivative of $F_{V_1}(x)$ and $F_{V_2}(x)$ and we get the pdf of V_1 and V_2 :

$$f_{V_1}(x) = 2x^{2-1}, \quad 0 < x < 1.$$

$$f_{V_2}(x) = \beta x^{\beta-1}, \quad 0 < x < 1.$$

The Jacobian determinant of transformation for $V = V_1, W = V + V_2$ is,

$$|J| = \begin{vmatrix} \frac{\partial V_1}{\partial V_1} & \frac{\partial V_1}{\partial V_2} \\ \frac{\partial W}{\partial V_1} & \frac{\partial W}{\partial V_2} \end{vmatrix} = \begin{vmatrix} 1 & 0 \\ 1 & 1 \end{vmatrix} = 1$$

Thus, the joint distribution is $f_{V_1, W}(v, w) = 2v^{2-1}\beta(w-v)^{\beta-1}$, where $0 < v < 1$ and $v < w < v+1$.

$$\begin{aligned} \text{We have that } F_X(x) &\propto \int_0^1 \int_0^{wx} 2v^{2-1}\beta(w-v)^{\beta-1} dv dw \\ &= \int_0^1 2(wx)^{2-1}\beta(w-wx)^{\beta-1} dw \\ &= x^{2-1}(1-x)^{\beta-1} \int_0^1 2\beta w^{2+\beta-2} dw \end{aligned}$$

Thus, X has a beta distribution with parameters 2 and β .

Bootstrap Procedures

Review:

Consider $n \sim \text{iid}$ observations X_1, \dots, X_n .

Bootstrap variance estimation:

Step 1. Generate $X_{1b}^*, \dots, X_{nb}^*$ iid from
the empirical function $F_n(\cdot)$.
(Resample with size n from X_1, \dots, X_n with replacement)

Step 2. Calculate $\hat{\theta}(X_{1b}^*, \dots, X_{nb}^*)$, denoted by $\hat{\theta}_b^*$.

Repeat step 1 and 2 B times and obtain

$$\{\hat{\theta}_b^* : b = 1, \dots, B\}$$

$$\text{Calculate } S_{\hat{\theta}}^2 = \frac{\sum_{b=1}^B (\hat{\theta}_b^* - \bar{\hat{\theta}}^*)^2}{B-1}$$

$$\text{with } \bar{\hat{\theta}}^* = \frac{\sum_{b=1}^B \hat{\theta}_b^*}{B}.$$

Use $S_{\hat{\theta}}^2$ to estimate $\text{Var}(\hat{\theta})$.

R codes:

```
B<-1000
```

```
n<-100
```

```
X<-rnorm(n)
```

```
thetabstarhat<-rep(0,B)
```

```
for(i in 1:B){
```

```
  Xstar<-sample(X,n,replace=TRUE)
```

```
  thetabstarhat[i]<-mean(Xstar)
```

```
}
```

```
thetastarhatbar<-sum(thetabstarhat)/B
```

```
ssquared<-0
```

```
for(j in 1:B){
```

```
  ssquared<-ssquared+(thetabstarhat[j]-thetastarhatbar)^2/(B-1)
```

```
}
```

```
ssquared
```

```
var(X)/n
```

Results:

```
> ssquared
```

```
[1] 0.01128843
```

```
> var(X)/n
```

```
[1] 0.01149071
```

Bootstrap Procedures

4.9.9. Let $\mathbf{x}' = (x_1, x_2, \dots, x_{n_1})$ and $\mathbf{y}' = (y_1, y_2, \dots, y_{n_2})$ be the realizations of the two random samples. The test statistic is the difference in sample means $\bar{y} - \bar{x}$. The estimated p -value of the test is calculated as follows:

1. Combine the data into one sample $\mathbf{z}' = (\mathbf{x}', \mathbf{y}')$.
2. Obtain all possible samples of size n_1 drawn without replacement from \mathbf{z} . Each such sample automatically gives another sample of size n_2 , i.e., all elements of \mathbf{z} not in the sample of size n_1 . There are $M = \binom{n_1+n_2}{n_1}$ such samples.
3. For each such sample j :
 - (a) Label the sample of size n_1 by \mathbf{x}^* and label the sample of size n_2 by \mathbf{y}^* .
 - (b) Calculate $v_j^* = \bar{y}^* - \bar{x}^*$.
4. The estimated p -value is $\hat{p}^* = \#\{v_j^* \geq \bar{y} - \bar{x}\} / M$.

Bootstrap Procedures

Suppose we have two samples each of size 3 which result in the realizations: $\mathbf{x}' = (10, 15, 21)$ and $\mathbf{y}' = (20, 25, 30)$. Determine the test statistic and the permutation test described above along with the p -value.

```
library(gtools)

x<-c(10,15,21)
y<-c(20,25,30)
z<-c(x,y)

xstar<-combinations(n=6,r=3,v=z)
xstarsum<-apply(xstar[1:20,], 1, sum)
xstarbar<-apply(xstar[1:20,], 1, mean)
zsum<-sum(z)
ystarsum<-zsum-xstarsum
ystarbar<-ystarsum/3
vstar<-ystarbar-xstarbar
ybar<-mean(y)
xbar<-mean(x)
pstarhat<-sum((ystarbar-xstarbar)>=(ybar-xbar))/20
```

[1] 0.1

Questions