

Topics for Today

- 1.) Comparing Models to Data
- 2.) Time Series + Stochastic Processes
 - Stationarity + Ergodicity
- 3.) Markov Chains + Stochastic Difference Equations
- 4.) Invariant Distributions
- 5.) Impulse Response Functions + Spectral Densities

Comparing Models to Data

- The models we develop in this course generally take one of two forms:

- 1.) Discrete-State Markov Chain
- 2.) Continuous-State Stochastic Difference Equation

How do we know whether a model is any good?

2 comments

- 1.) Models are designed to address specific questions. They should be evaluated based on their answers to the question at hand. A model may be good for some purposes, but bad for others. (Of course, the more questions a model can address, the better is the model!).
 - 2.) It is important to keep in mind that the equilibria of our models are stochastic processes, not fixed numbers. Hence, they cannot be judged based on a single outcome.
- Instead, we must compare the statistical properties of the data our model hypothetically would generate, with the same statistical properties of the observed data.

- The time-series properties of our models are determined by a (small) set of parameters characterizing preferences, technology, and market structure.
- There are 2 basic strategies for comparing models to data:

1.) Econometrics: Pick parameters to minimize the "distance" between model data + actual data, and test using χ^2 or LR statistics.

Problem: All models are false. Rejections can be uninformative.

2.) Calibration: Pick the parameters to match some pre-selected moments (usually long-run means) and/or take them from other studies. Then, compare the remaining moments of interest to those in the data.

Problem: Must rely on "eyeball metric", sensitivity to selected moments.

- Either way, to carry out this program, we obviously need to know something about time series; in particular, about Markov Chains + Stochastic Difference Equations.

Key Concepts

① STATE : A complete description of all relevant aspects of the current position of a dynamical system.

Examples :

- 1.) Capital Stock
- 2.) Technology
- 3.) Financial wealth
- 4.) Employment Status
- 5.) Beliefs

Different Kinds of State Variables

- Endogenous vs. Exogenous
 - Discrete vs. Continuous
 - Finite Dimensional vs. Infinite Dimensional
- > capital vs. Technology
- > Employment vs. Wealth Status
- > Wealth vs. Beliefs.

The State plays a role in 2 ways

- $S_{t+1} = f(S_t, X_t, \varepsilon_{t+1})$
 - $X_t = g(S_t)$
- > state transition eg. (e.g., budget constraint) or Bayes Rule
- > policy function

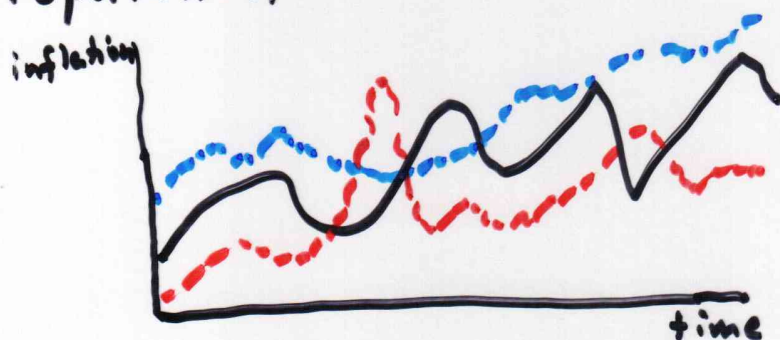
② TIME SERIES / STOCHASTIC PROCESS: A time-indexed sequence of random variables.

Stochastic Processes vs. Random Variables
(paths) (real numbers)

- We need to regard a time series as a single unit

Comments

- 1.) A realization is an entire path, not a number
- 2.) Population = Set of all possible realizations



- 3.) We observe only one realization (unless you're a believer in the "many worlds" interpretation of quantum mechanics!).
- 4.) "Law of Large Numbers" and "Central Limit Theorem" arguments still applicable if observations over time provide enough new information about ensemble averages.

③ STATIONARY STOCHASTIC PROCESS / COVARIANCE STATIONARITY

A stochastic process where all joint distributions are independent of time. "Initial conditions have worn off".

$$f(x_t, x_{t+1}, \dots, x_{t+j}) \text{ indpt. of } t \quad \forall j$$

④ MARKOV PROCESS : Restricts conditional distribution function

$$f(\text{future} | \text{present, past}) = f(\text{future} | \text{present})$$

$$\text{Prob}(x_{t+1} | x_t, x_{t-1}, \dots, x_{t-k}) = \text{Prob}(x_{t+1} | x_t) \quad \forall k \geq 1$$

- Whether a process is Markov depends on the dimensionality of the state

Example : $S_t = \alpha_1 S_{t-1} + \alpha_2 S_{t-2} + \alpha_3 S_{t-3} + \epsilon_t$ } 3rd order Markov AR(3)

Redefine state as $Z_t = \begin{pmatrix} S_t \\ S_{t-1} \\ S_{t-2} \end{pmatrix}$

Then $Z_t = A Z_{t-1} + B \epsilon_t$

> 1st order Markov

$$A = \begin{pmatrix} \alpha_1 & \alpha_2 & \alpha_3 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \end{pmatrix} \quad B = \begin{pmatrix} 1 \\ 0 \\ 0 \end{pmatrix}$$

⑤ MARKOV CHAIN: A discrete-state Markov Process

- A Markov Chain is completely characterized by its Transition Matrix, P , where

$$P_{ij} = \text{Prob}(X_{t+1} = X_j \mid X_t = X_i)$$

$$\text{Note, } \sum_{j=1}^N P_{ij} = 1$$

⑥ STATIONARY (INVARIANT) DISTRIBUTION:

A distribution that remains invariant over time (stochastic analog of a limit point).

- Consider the case of a Markov Chain

$$\text{Let } \pi_0 = \text{initial dist. over states} \begin{pmatrix} \pi_0(1) \\ \pi_0(2) \\ \pi_0(3) \end{pmatrix}$$

$$\text{Note, } \pi'_1 = \pi'_0 P = (\pi_0(1) \ \pi_0(2) \ \pi_0(3)) \begin{bmatrix} P_{11} & P_{12} & P_{13} \\ P_{21} & P_{22} & P_{23} \\ P_{31} & P_{32} & P_{33} \end{bmatrix}$$

Hence,

$$\begin{pmatrix} \pi_1(1) \\ \pi_1(2) \\ \pi_1(3) \end{pmatrix} = \begin{pmatrix} \pi_0(1)P_{11} + \pi_0(2)P_{21} + \pi_0(3)P_{31} \\ \pi_0(1)P_{12} + \pi_0(2)P_{22} + \pi_0(3)P_{32} \\ \pi_0(1)P_{13} + \pi_0(2)P_{23} + \pi_0(3)P_{33} \end{pmatrix}$$

Note, if the Markov Chain is "homogeneous", then P does not depend on time, so

$$\pi_2' = \pi_0' P^2 \quad \text{or more generally} \quad \pi_k' = \pi_0' P^k$$

Clearly, the stationary (or invariant) dist. is characterized by:

$$\pi_\infty' = \pi_\infty' P \Rightarrow \pi_\infty' (I - P) = 0$$

$$\Rightarrow (I - P') \pi_\infty = 0$$

π_∞ = eigenvector of P' associated with a unit eigenvalue (normalized so that $\sum \pi_\infty(i) = 1$).

Since P is a "stochastic matrix" [non-negative elements with row sum = 1] it always has at least one unit eigenvalue. [It's unique if $P_{ij} > 0 \forall i, j$].

Example: Suppose $P = \begin{pmatrix} 1-\alpha & \alpha \\ \beta & 1-\beta \end{pmatrix}$

$$\text{Then, } \pi_{\infty} = \left(\frac{\beta}{\alpha+\beta}, \frac{\alpha}{\alpha+\beta} \right)$$

⑦ ERGODIC MARKOV CHAIN: A Markov Chain with a unique invariant distribution

- Having multiple stationary distributions is the stochastic analog of having multiple equilibria.
- Ergodicity is important because it guarantees sample averages over time converge to population moments (Law of Large Numbers).
- A Markov Chain is ergodic when it is possible to go from every state to every state (not necessarily in one jump). This property is sometimes called irreducibility.
- If $P_{ij} > 0 \forall i, j$ then obviously the chain is ergodic. However, this is not necessary.

• Example of non-ergodic Markov Chain:

$$\begin{pmatrix} 1 & 0 & 0 \\ 0 & \frac{1}{2} & \frac{1}{2} \\ 0 & \frac{1}{2} & \frac{1}{2} \end{pmatrix}$$

⑧ LINEAR STOCHASTIC DIFFERENCE EQUATIONS:

$$Z_{t+1} = A Z_t + B \varepsilon_{t+1}$$

where ε_{t+1} is mean zero, iid, with continuous dist.

- A common specification for ε_{t+1} is $N(0, \Sigma)$.
When ε_{t+1} is Gaussian, the first 2 moments of Z_t characterize the entire distribution.
In particular, "covariance stationarity" implies stationarity.

First Moment

Let $\mu_t = E(Z_t)$. Then,

$$\mu_{t+1} = A \mu_t \Rightarrow (I - A) \mu_{\infty} = 0$$

$\Rightarrow$ Unconditional / Stationary mean is the eigenvector associated with (unique) unit eigenvalue.

Sufficient Condition: A has all eigenvalues inside unit circle.

Second Moment

Define

$$C(j) = E(x_{t+j} - \mu)(x_t - \mu)' \quad > \text{Covariogram (indpt. of time)}$$

$$C(0) = \text{Variance}$$

$$= E[(A z_t + B \varepsilon_{t+1})(z_t' A' + \varepsilon_{t+1}' B')]$$

$$\Rightarrow C(0) = A C(0) A' + B \Sigma B' \quad > \text{Lyapunov Eq.}$$

$$C(j) = A^j C(0)$$

- There are 2 common ways to summarize the dynamics of a stochastic difference equation:

a.) Impulse Response Function.

b.) Spectral Density.

⑨ IMPULSE RESPONSE FUNCTION:

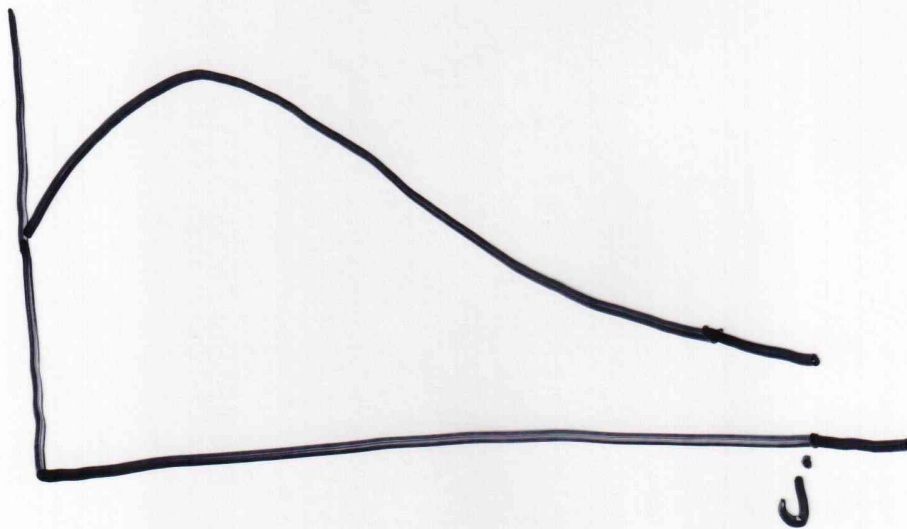
A plot of the system's moving average coefficients
write the stoch. diff. eq. using lag operators,

$$(I - AL)X_t = B\varepsilon_t \quad \text{where } LX_t = X_{t-1}$$

$$\Rightarrow X_t = (I - AL)^{-1} B \varepsilon_t$$
$$= \sum_{j=0}^{\infty} A^j B \varepsilon_{t-j}$$

> converges if A
has all eigenvalues
inside unit circle

IRF
 $= A^j B$



⑩ SPECTRAL DENSITY : Fourier Transform of the covariogram

Basic idea : Decompose autocovariances into orthogonal components

Why do this ?

1.) Computational - Orthogonality $\Rightarrow$ Non-interference
No cross-terms

Converts functional equations into algebraic eqs.
Solve the algebraic eq. then transform back !

2.) Conceptual - Separate components may be interesting (e.g., business cycles, or seasonal cycles).

Fourier Transforms

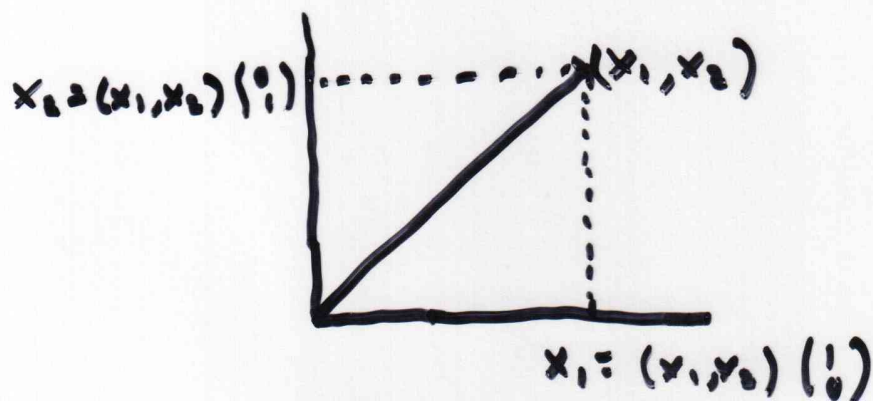
• Let's start slow. You know that any vector in $\mathbb{R}^2$ can be written :

$$\begin{pmatrix} x_1 \\ x_2 \end{pmatrix} = x_1 \cdot \begin{pmatrix} 1 \\ 0 \end{pmatrix} + x_2 \cdot \begin{pmatrix} 0 \\ 1 \end{pmatrix}$$

> sum of
orthog.
components

where the coefficients are given by an inner product:

$$x_1 = (x_1, x_2) \begin{pmatrix} 1 \\ 0 \end{pmatrix} \quad x_2 = (x_1, x_2) \begin{pmatrix} 0 \\ 1 \end{pmatrix}$$



The idea behind Fourier Transforms is exactly the same, but rather than expressing a vector in $\mathbb{R}^n$ (a finite dimensional object) as an orthogonal sum, we want to express an "arbitrary" function as an orthogonal sum

what does this mean?

$$f(x) = \frac{a_0}{2} + \sum_{n=1}^{\infty} (a_n \cos nx + b_n \sin nx) \quad -\pi < x < \pi$$

$$a_n = \frac{1}{\pi} \int_{-\pi}^{\pi} f(x) \cos nx \, dx$$

$$b_n = \frac{1}{\pi} \int_{-\pi}^{\pi} f(x) \sin nx \, dx$$

Same as above
with inner product
defined as
 $(f, g) = \int_{-\pi}^{\pi} f(x)g(x)dx$

Note: (a_n, b_n) are like regression coefficients from projecting f onto sine & cosines.

Orthogonality

$$\int_{-\pi}^{\pi} \cos(mx) \sin(nx) dx = 0 \quad \forall m, n$$

$$\int_{-\pi}^{\pi} \cos(mx) \cos(nx) dx = 0 \quad m \neq n$$

$$\int_{-\pi}^{\pi} \sin(mx) \sin(nx) dx = 0 \quad m \neq n$$

- For functions defined on the whole real line (which are not periodic) we need a fourier integral
$$f(x) = \int_{-\infty}^{\infty} a(\omega) e^{-i\omega x} d\omega$$

Examples

$$S(\omega) = \sum_{j=-\infty}^{\infty} c(j) e^{-i\omega j} \quad \frac{\omega}{2\pi} = \text{frequency (cycles per period)}$$

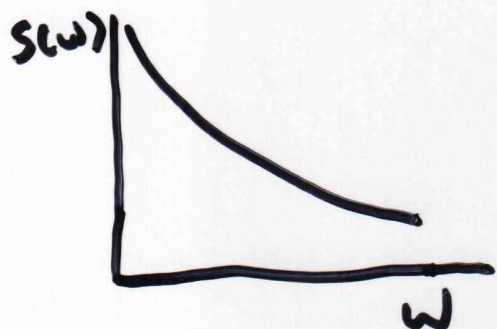
$$= (I - A e^{-i\omega})^{-1} B B' (I - A' e^{i\omega})^{-1}$$

$$c(j) = \frac{1}{2\pi} \int_{-\pi}^{\pi} S(\omega) e^{i\omega j} d\omega$$

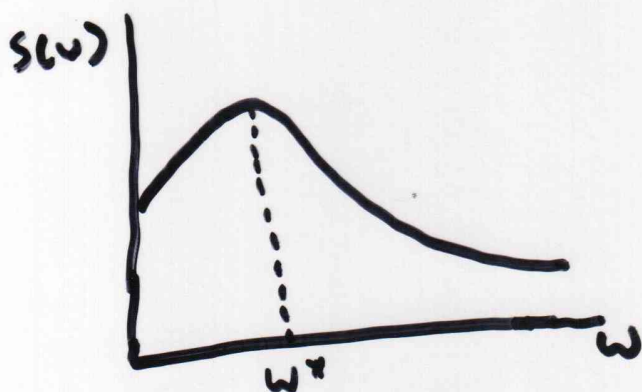
scalar case, $x_t = a x_{t-1} + \varepsilon_t \Rightarrow c_j = |a|^j c_0$

$$\Rightarrow S(\omega) = \frac{c_0(1-a^2)}{(1-a e^{-i\omega})(1-a e^{i\omega})} = \frac{c_0(1-a^2)}{1+a^2-a \cos \omega}$$

- Most economic time series have spectral densities that look like:



- A series with a cycle has a spectral density with a local peak



The implied cycle has period $\frac{2\pi}{\omega^*}$

The theoretical foundation of Fourier analysis is the following theorem:

Riesz-Fischer Theorem: Let $\{c_n\}$ be a square-summable sequence of complex numbers, (i.e., $\sum |c_n|^2 < \infty$). Then there exists a complex-valued function, $g(\omega)$, defined for $\omega \in [-\pi, \pi]$, s.t.

$$(A) \quad g(\omega) = \sum_{j=-\infty}^{\infty} c_j e^{-i\omega j}$$

where convergence is in the mean-square sense,

$$\lim_{n \rightarrow \infty} \int_{-\pi}^{\pi} \left| \sum_{j=-n}^n c_j e^{-i\omega j} - g(\omega) \right|^2 d\omega = 0$$

and $g(\omega)$ is square (Lebesgue) integrable

$$\int_{-\pi}^{\pi} |g(\omega)|^2 d\omega < \infty.$$

Conversely, given a square integrable $g(\omega)$,

$\exists$ a square summable sequence s.t.

$$(B) \quad c_k = \frac{1}{2\pi} \int_{-\pi}^{\pi} g(\omega) e^{i\omega k} d\omega$$

• (A) + (B) form a "Fourier Transform pair" linked via the Parseval Identity

$$\frac{1}{2\pi} \int_{-\pi}^{\pi} |g(\omega)|^2 d\omega = \sum_{j=-\infty}^{\infty} |c_j|^2$$

In words, $\exists$ an isometric isomorphism between ℓ_2 (time domain) and L^2 (freq domain) (i.e., a mapping that is one-to-one and preserves distance + linear structure).