

## STAT 450: Statistical Theory

### Expectation, moments

Two elementary definitions of expected values:

**Defn:** If  $X$  has density  $f$  then

$$E(g(X)) = \int g(x)f(x) dx.$$

**Defn:** If  $X$  has discrete density  $f$  then

$$E(g(X)) = \sum_x g(x)f(x).$$

**FACT:** If  $Y = g(X)$  for a smooth  $g$

$$\begin{aligned} E(Y) &= \int y f_Y(y) dy \\ &= \int g(x) f_Y(g(x)) g'(x) dy \\ &= E(g(X)) \end{aligned}$$

by the change of variables formula for integration. This is good because otherwise we might have two different values for  $E(e^X)$ .

In general, there are random variables which are neither absolutely continuous nor discrete. See STAT 801 for general definition of E.

**Defn:** We call  $X$  **integrable** if

$$E(|X|) < \infty.$$

Facts:  $E$  is a linear, monotone, positive operator:

1. **Linear:**  $E(aX + bY) = aE(X) + bE(Y)$  provided  $X$  and  $Y$  are integrable.
2. **Positive:**  $P(X \geq 0) = 1$  implies  $E(X) \geq 0$ .
3. **Monotone:**  $P(X \geq Y) = 1$  and  $X, Y$  integrable implies  $E(X) \geq E(Y)$ .

Major technical theorems:

**Monotone Convergence:** If  $0 \leq X_1 \leq X_2 \leq \dots$  and  $X = \lim X_n$  (which has to exist) then

$$E(X) = \lim_{n \rightarrow \infty} E(X_n).$$

**Dominated Convergence:** If  $|X_n| \leq Y_n$  and  $\exists$  rv  $X$  such that  $X_n \rightarrow X$  (technical details of this convergence later in the course) and a random variable  $Y$  such that  $Y_n \rightarrow Y$  with  $E(Y_n) \rightarrow E(Y) < \infty$  then

$$E(X_n) \rightarrow E(X).$$

Often used with all  $Y_n$  the same rv  $Y$ .

**Theorem:** With this definition of  $E$  if  $X$  has density  $f(x)$  (even in  $R^p$  say) and  $Y = g(X)$  then

$$E(Y) = \int g(x)f(x)dx.$$

(Could be a multiple integral.) If  $X$  has pmf  $f$  then

$$E(Y) = \sum_x g(x)f(x).$$

Firts conclusion works, e.g., even if  $X$  has a density but  $Y$  doesn't.

**Defn:** The  $r^{\text{th}}$  moment (about the origin) of a real rv  $X$  is  $\mu'_r = E(X^r)$  (provided it exists). We generally use  $\mu$  for  $E(X)$ .

**Defn:** The  $r^{\text{th}}$  central moment is

$$\mu_r = E[(X - \mu)^r].$$

We call  $\sigma^2 = \mu_2$  the variance.

**Defn:** For an  $R^p$  valued random vector  $X$

$$\mu_X = E(X)$$

is vector whose  $i^{\text{th}}$  entry is  $E(X_i)$  (provided all entries exist).

**Defn:** The  $(p \times p)$  variance covariance matrix of  $X$  is

$$\text{Var}(X) = E[(X - \mu)(X - \mu)^t]$$

which exists provided each component  $X_i$  has a finite second moment.

Moments and probabilities of rare events are closely connected as will be seen in a number of important probability theorems.

**Example:** Markov's inequality

$$\begin{aligned} P(|X - \mu| \geq t) &= E[1(|X - \mu| \geq t)] \\ &\leq E\left[\frac{|X - \mu|^r}{t^r} 1(|X - \mu| \geq t)\right] \\ &\leq \frac{E[|X - \mu|^r]}{t^r} \end{aligned}$$

Intuition: if moments are small then large deviations from average are unlikely.

Special Case: Chebyshev's inequality

$$P(|X - \mu| \geq t) \leq \frac{\text{Var}(X)}{t^2}.$$

**Example moments:** If  $Z$  is standard normal then

$$\begin{aligned} E(Z) &= \int_{-\infty}^{\infty} z e^{-z^2/2} dz / \sqrt{2\pi} \\ &= \left. \frac{-e^{-z^2/2}}{\sqrt{2\pi}} \right|_{-\infty}^{\infty} \\ &= 0 \end{aligned}$$

and (integrating by parts)

$$\begin{aligned} E(Z^r) &= \int_{-\infty}^{\infty} z^r e^{-z^2/2} dz / \sqrt{2\pi} \\ &= \left. \frac{-z^{r-1} e^{-z^2/2}}{\sqrt{2\pi}} \right|_{-\infty}^{\infty} \\ &\quad + (r-1) \int_{-\infty}^{\infty} z^{r-2} e^{-z^2/2} dz / \sqrt{2\pi} \end{aligned}$$

so that

$$\mu_r = (r-1)\mu_{r-2}$$

for  $r \geq 2$ . Remembering that  $\mu_1 = 0$  and

$$\mu_0 = \int_{-\infty}^{\infty} z^0 e^{-z^2/2} dz / \sqrt{2\pi} = 1$$

we find that

$$\mu_r = \begin{cases} 0 & r \text{ odd} \\ (r-1)(r-3)\cdots 1 & r \text{ even} . \end{cases}$$

If now  $X \sim N(\mu, \sigma^2)$ , that is,  $X \sim \sigma Z + \mu$ , then  $E(X) = \sigma E(Z) + \mu = \mu$  and

$$\mu_r(X) = E[(X - \mu)^r] = \sigma^r E(Z^r) .$$

In particular, we see that our choice of notation  $N(\mu, \sigma^2)$  for the distribution of  $\sigma Z + \mu$  is justified;  $\sigma$  is indeed the variance.

Similarly for  $X \sim MVN(\mu, \Sigma)$  we have  $X = AZ + \mu$  with  $Z \sim MVN(0, I)$  and

$$E(X) = \mu$$

and

$$\begin{aligned} \text{Var}(X) &= E \{ (X - \mu)(X - \mu)^t \} \\ &= E \{ AZ(AZ)^t \} \\ &= AE(ZZ^t)A^t \\ &= AIA^t = \Sigma . \end{aligned}$$

Note use of easy calculation:  $E(Z) = 0$  and

$$\text{Var}(Z) = E(ZZ^t) = I .$$

### Moments and independence

**Theorem:** If  $X_1, \dots, X_p$  are independent and each  $X_i$  is integrable then  $X = X_1 \cdots X_p$  is integrable and

$$E(X_1 \cdots X_p) = E(X_1) \cdots E(X_p) .$$