

Factor Analysis

Suppose $\mathbf{Y} \sim MVN_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$. Imagine that \mathbf{Y} is a vector of scores on various tests.

Idea: study structure of $\boldsymbol{\Sigma}$ or of R , look to see if \mathbf{Y} is “explained” by a small number of factors common to all variables together with some variability in individual variability which is independent between variables.

This amounts to writing

$$\mathbf{Y} = \sum_1^m \boldsymbol{\lambda}_i X_i + \boldsymbol{\epsilon}$$

where we assume that $\boldsymbol{\epsilon}$ has independent components and the $\boldsymbol{\lambda}_i$ are p vectors. In matrix form we assume

$$\mathbf{Y} = \boldsymbol{\Lambda}\mathbf{X} + \boldsymbol{\epsilon}$$

where $\boldsymbol{\Lambda}$ is $p \times m$ (hopefully with m small) and $\boldsymbol{\epsilon}$ has diagonal variance covariance

$\boldsymbol{\Psi}$

We also assume that \mathbf{X} is independent of $\boldsymbol{\epsilon}$.

Then the variance covariance of \mathbf{Y} has the form

$$\Sigma_{\mathbf{Y}} = \Lambda\Lambda^T + \Psi$$

Jargon

Specificities: (or specific variances or uniquenesses) — the ψ_i , the diagonal entries in Ψ .

Common factors: the variables X_1, \dots, X_m .

Loading: Λ_{ij} is the “loading of the i th response on the j th common factor”.

Communalities: The parts of the variances of the Y_i which arise from the common factors, that is,

$$\sigma_i^2 = (\Sigma_{\mathbf{Y}})_{ii} - \psi_i = (\Lambda\Lambda^T)_{ii}$$

What can be estimated?

Given $\mathbf{Y}_1, \dots, \mathbf{Y}_n$ iid $MVN_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$.

Assume $\boldsymbol{\Sigma}$ has factor structure for m factors.

Log likelihood is

$$\begin{aligned} \ell(\boldsymbol{\mu}, \boldsymbol{\Lambda}, \boldsymbol{\Psi}) = & -n \log \det(\boldsymbol{\Lambda}\boldsymbol{\Lambda}^T + \boldsymbol{\Psi}) \\ & - \frac{1}{2} \sum (\mathbf{Y}_i - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{Y}_i - \boldsymbol{\mu}) \end{aligned}$$

Notice that $\boldsymbol{\Sigma}$ depends only on $\boldsymbol{\Lambda}\boldsymbol{\Lambda}^T$ so that if

$$\boldsymbol{\Lambda}_1\boldsymbol{\Lambda}_1^T = \boldsymbol{\Lambda}_2\boldsymbol{\Lambda}_2^T$$

then

$$\ell(\boldsymbol{\mu}, \boldsymbol{\Lambda}_1, \boldsymbol{\Psi}) = \ell(\boldsymbol{\mu}, \boldsymbol{\Lambda}_2, \boldsymbol{\Psi})$$

This means that Λ is **not identifiable**.

If two parameter values each give the same density for the data then the data do not distinguish between the two parameter values.

In factor analysis we use subject matter understanding to try to pick one particular Λ from the collection which maximize ℓ on the basis of **external** or **a priori** criteria.

But: MLE of Ψ and $\Lambda\Lambda^T$ possible.

Factor Analysis example

Data: Table 9.12 in Johnson and Wichern; 3 measurements of sales performance and 4 test scores of 50 salespeople for a large firm. The data begin:

Case	Growth	Profit	New	Creat	Mech	Abst	Math
1	93.0	96.0	97.8	9	12	9	20
2	88.8	91.8	96.8	7	10	10	15
3	95.0	100.3	99.0	8	12	9	26

SAS examples: First: principal components factor analysis, no rotation, all output printed, SAS selects m , number of factors:

```
data sales;
  infile "T9-12.DAT";
  input growth profit new
          create mech abst math;
proc factor method=prin
          rotate=none all;
run;
```

The (edited) output is

Initial Factor Method: Principal Components

Inverse Correlation Matrix

	GROWTH	PROFIT	NEW	CREATE	MECH	ABST	MATH
GROWTH	35.14	-8.26	15.74	-13.90	-1.55	-13.86	-23.71
PROFIT	-8.26	31.62	-2.70	1.41	-8.18	6.58	-19.50
NEW	15.74	-2.70	21.69	-13.24	-0.13	-10.46	-19.06
CREATE	-13.90	1.41	-13.24	10.53	-0.12	7.64	14.25
MECH	-1.55	-8.18	-0.13	-0.12	4.56	-0.88	7.22
ABST	-13.86	6.58	-10.46	7.64	-0.88	8.68	7.99
MATH	-23.71	-19.50	-19.06	14.25	7.22	7.99	43.09

Partial Correlations Controlling all other Variables

	GROWTH	PROFIT	NEW	CREATE	MECH	ABST	MATH
GROWTH	1.000	0.248	-0.570	0.722	0.123	0.793	0.609
PROFIT	0.248	1.000	0.103	-0.077	0.681	-0.397	0.528
NEW	-0.570	0.103	1.000	0.876	0.013	0.762	0.623
CREATE	0.722	-0.077	0.876	1.000	0.018	-0.798	-0.668
MECH	0.123	0.681	0.013	0.018	1.000	0.140	-0.515
ABST	0.793	-0.397	0.762	-0.798	0.140	1.000	-0.413
MATH	0.609	0.528	0.623	-0.668	-0.515	-0.413	1.000

Prior Communalities Estimates: ONE

Eigenvalues of the Correlation Matrix:

	Total = 7				Average = 1		
	1	2	3	4	5	6	7
Eigenvalue	5.0346	0.9335	0.4979	0.4212	0.0810	0.0203	0.0113
Difference	4.1011	0.4356	0.0767	0.3402	0.0607	0.0090	
Proportion	0.7192	0.1334	0.0711	0.0602	0.0116	0.0029	0.0016
Cumulative	0.7192	0.8526	0.9237	0.9839	0.9955	0.9984	1.0000

1 factors will be retained by the MINEIGEN criterion.

Eigenvectors

GROWTH	0.43367
PROFIT	0.42021
NEW	0.42105
CREATE	0.29429
MECH	0.34909
ABST	0.28917
MATH	0.40740

Factor Pattern

	FACTOR1
GROWTH	0.97307
PROFIT	0.94287
NEW	0.94475
CREATE	0.66032
MECH	0.78329
ABST	0.64883
MATH	0.91413

Variance explained by each factor

FACTOR1	5.034598
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Communalities Estimates: Total = 5.034598

GROWTH	PROFIT	NEW	CREATE	MECH	ABST	MATH
0.9468	0.8890	0.8925	0.4360	0.6135	0.4209	0.8356

Note: matrix of loadings called Factor Pattern.

Can estimate latent variables (Factor Scores).

Scoring Coefficients Estimated by Regression
Squared Multiple Correlations of
the Variables with each Factor

FACTOR1
1.000000

Standardized Scoring Coefficients

	FACTOR1
GROWTH	0.19328
PROFIT	0.18728
NEW	0.18765
CREATE	0.13116
MECH	0.15558
ABST	0.12887
MATH	0.18157

Residual Correlations With Uniqueness on the Diagonal

	GROWTH	PROFIT	NEW	CREATE	MECH	ABST	MATH
GROWTH	0.053	0.008	-0.035	-0.070	-0.054	0.043	0.037
PROFIT	0.008	0.110	-0.048	-0.081	0.007	-0.146	0.082
NEW	-0.035	-0.048	0.107	0.076	-0.102	0.028	-0.011
CREATE	-0.070	-0.081	0.076	0.563	0.073	-0.281	-0.190
MECH	-0.054	0.007	-0.102	0.073	0.386	-0.122	-0.141
ABST	0.043	-0.146	0.028	-0.281	-0.122	0.579	-0.026
MATH	0.037	0.082	-0.011	-0.190	-0.141	-0.026	0.164

Root Mean Square Off-diagonal Residuals:

Over-all = 0.10322669

GROWTH	PROFIT	NEW	CREATE	MECH	ABST	MATH
0.0456	0.0787	0.0589	0.1519	0.0947	0.1408	0.1045

Partial Correlations Controlling Factors

	GROWTH	PROFIT	NEW	CREATE	MECH	ABST	MATH
GROWTH	1.000	0.111	-0.467	-0.407	-0.377	0.245	0.404
PROFIT	0.111	1.000	-0.441	-0.324	0.035	-0.577	0.609
NEW	-0.467	-0.441	1.000	0.310	-0.503	0.112	-0.083
CREATE	-0.407	-0.324	0.310	1.000	0.157	-0.492	-0.627
MECH	-0.377	0.035	-0.503	0.157	1.000	-0.258	-0.561
ABST	0.245	-0.577	0.112	-0.492	-0.258	1.000	-0.086
MATH	0.404	0.609	-0.083	-0.627	-0.561	-0.086	1.000

Root Mean Square Off-diagonal Partial:

Over-all = 0.38975152

GROWTH	PROFIT	NEW	CREATE	MECH	ABST	MATH
0.3566	0.4122	0.3612	0.4140	0.3660	0.3472	0.4580

The procedure selects only one factor on which all the variables are fairly highly loaded. The factor is the first principal component of the correlation matrix.

Now insist on two factors, use varimax rotation and plot factor loadings before and after rotation.

```
proc factor m=prin nfactor=2 rotate=v
           preplot plot all;
```

The output

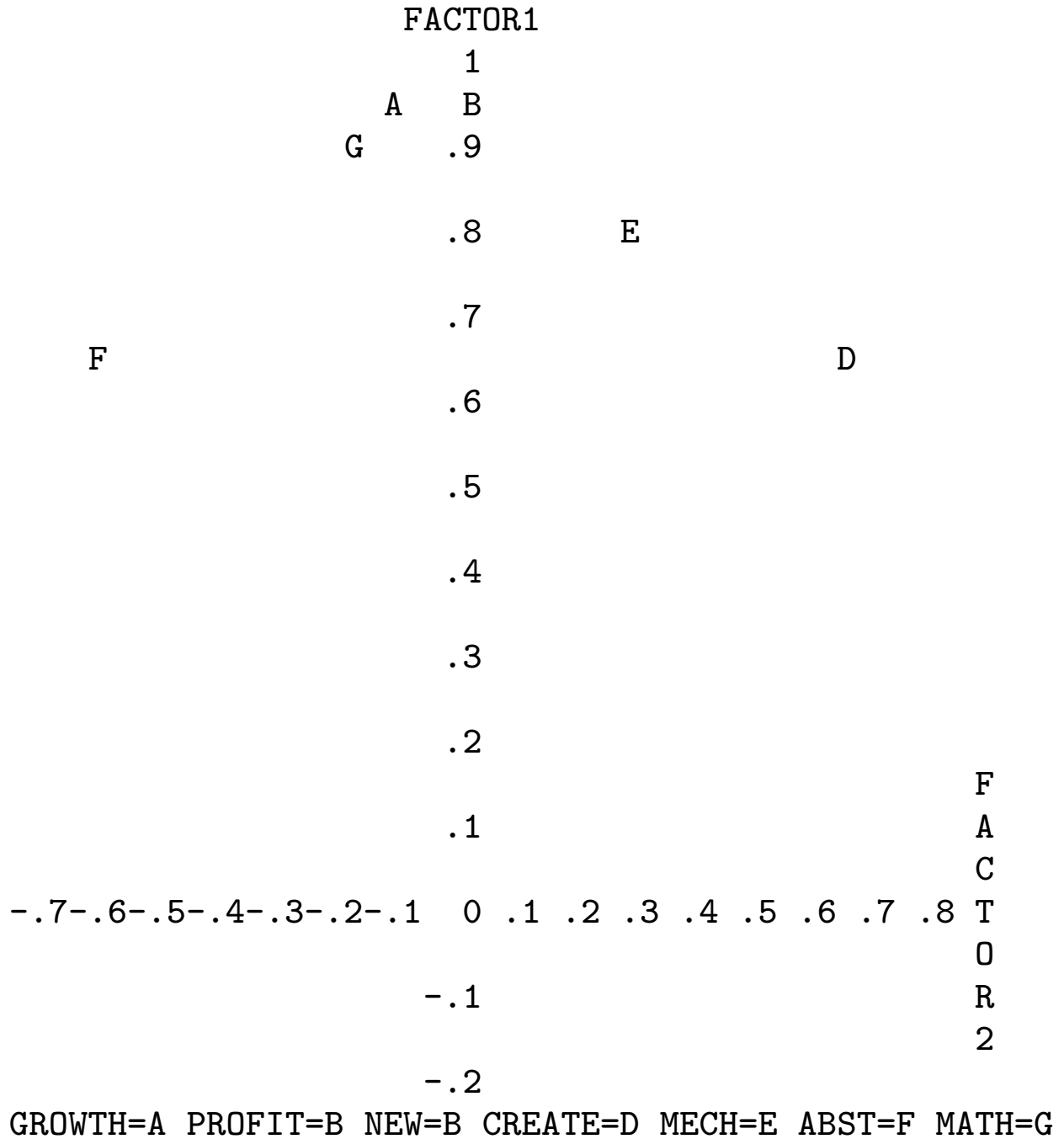
Factor Pattern

	FACTOR1	FACTOR2
GROWTH	0.97307	-0.10798
PROFIT	0.94287	0.02830
NEW	0.94475	0.00889
CREATE	0.66032	0.64581
MECH	0.78329	0.28497
ABST	0.64883	-0.62066
MATH	0.91413	-0.19359

Variance explained by each factor

FACTOR1	FACTOR2
5.034598	0.933516

Plot of Factor Pattern for FACTOR1 and FACTOR2



Rotation Method: Varimax
Orthogonal Transformation Matrix

	1	2
1	0.73145	0.68189
2	-0.68189	0.73145

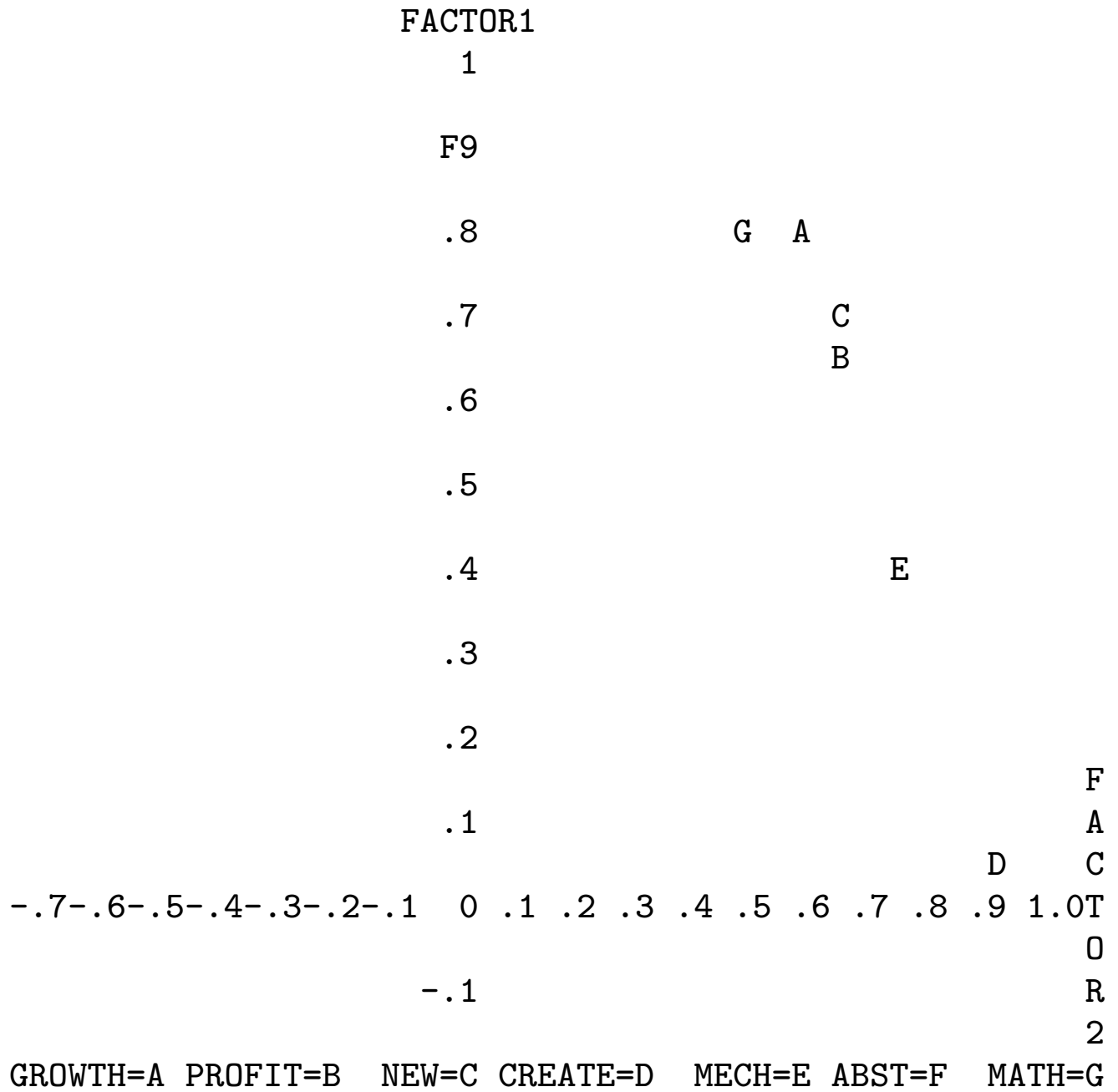
Rotated Factor Pattern

	FACTOR1	FACTOR2
GROWTH	0.78538	0.58455
PROFIT	0.67037	0.66364
NEW	0.68498	0.65072
CREATE	0.04261	0.92265
MECH	0.37862	0.74256
ABST	0.89781	-0.01155
MATH	0.80065	0.48174

Variance explained by each factor

FACTOR1	FACTOR2
3.127683	2.840431

Plot of Factor Pattern for FACTOR1 and FACTOR2



With principal components factor analysis fitting a second factor does not change the first factor (before rotation). Creativity and abstract reasoning are loaded on the second factor with opposite signs which would appear to represent a difference between people on a dimension of creativity as opposed to abstract reasoning.

After rotation everything except creativity is loaded on factor 1, though Mechanical reasoning has a rather smaller loading. Abstraction is not loaded on factor 2.

Now I tried iterated principal factor analysis with varimax rotation. The option `heywood` permits iteration to continue if the estimated uniqueness of a variable drops below 0.

```
proc factor m=prinit nfactor=2 rotate=v
            preplot plot all heywood;
run;
```

Initial Factor Method:

Iterated Principal Factor Analysis

Prior Communalities Estimates: ONE

2 factors will be retained by the NFACTOR criterion.

IterChange Communalities

1	0.3052	0.958	0.889	0.892	0.853	0.694	0.806	0.873
2	0.1223	0.968	0.874	0.879	0.805	0.600	0.683	0.850
3	0.0903	0.980	0.874	0.880	0.797	0.574	0.593	0.855
4	0.0638	0.988	0.876	0.881	0.807	0.568	0.529	0.866
5	0.0401	0.993	0.877	0.882	0.824	0.564	0.489	0.878
6	0.0229	0.995	0.879	0.882	0.844	0.562	0.466	0.886
7	0.0191	0.996	0.880	0.882	0.863	0.560	0.453	0.892
8	0.0183	0.996	0.880	0.882	0.881	0.558	0.447	0.895
9	0.0174	0.996	0.880	0.881	0.899	0.556	0.443	0.897
10	0.0165	0.996	0.881	0.881	0.915	0.554	0.440	0.899
11	0.0156	0.996	0.881	0.880	0.931	0.552	0.438	0.899
12	0.0148	0.996	0.881	0.880	0.946	0.551	0.437	0.899
13	0.0141	0.996	0.881	0.879	0.960	0.549	0.436	0.899
14	0.0135	0.996	0.881	0.879	0.973	0.548	0.435	0.899
15	0.0129	0.995	0.881	0.879	0.986	0.547	0.434	0.899
16	0.0124	0.995	0.882	0.878	0.999	0.546	0.433	0.899
17	0.0010	0.995	0.882	0.878	1.000	0.545	0.432	0.899
18	0.0002	0.996	0.882	0.878	1.000	0.544	0.432	0.899

Convergence criterion satisfied.

Eigenvalues of the Reduced Correlation Matrix:

Total = 5.63350769 Average = 0.80478681

	1	2	3
Eigenvalue	4.8786	0.7663	0.2047
Difference	4.1123	0.5616	0.1209
Proportion	0.8660	0.1360	0.0363
Cumulative	0.8660	1.0020	1.0384

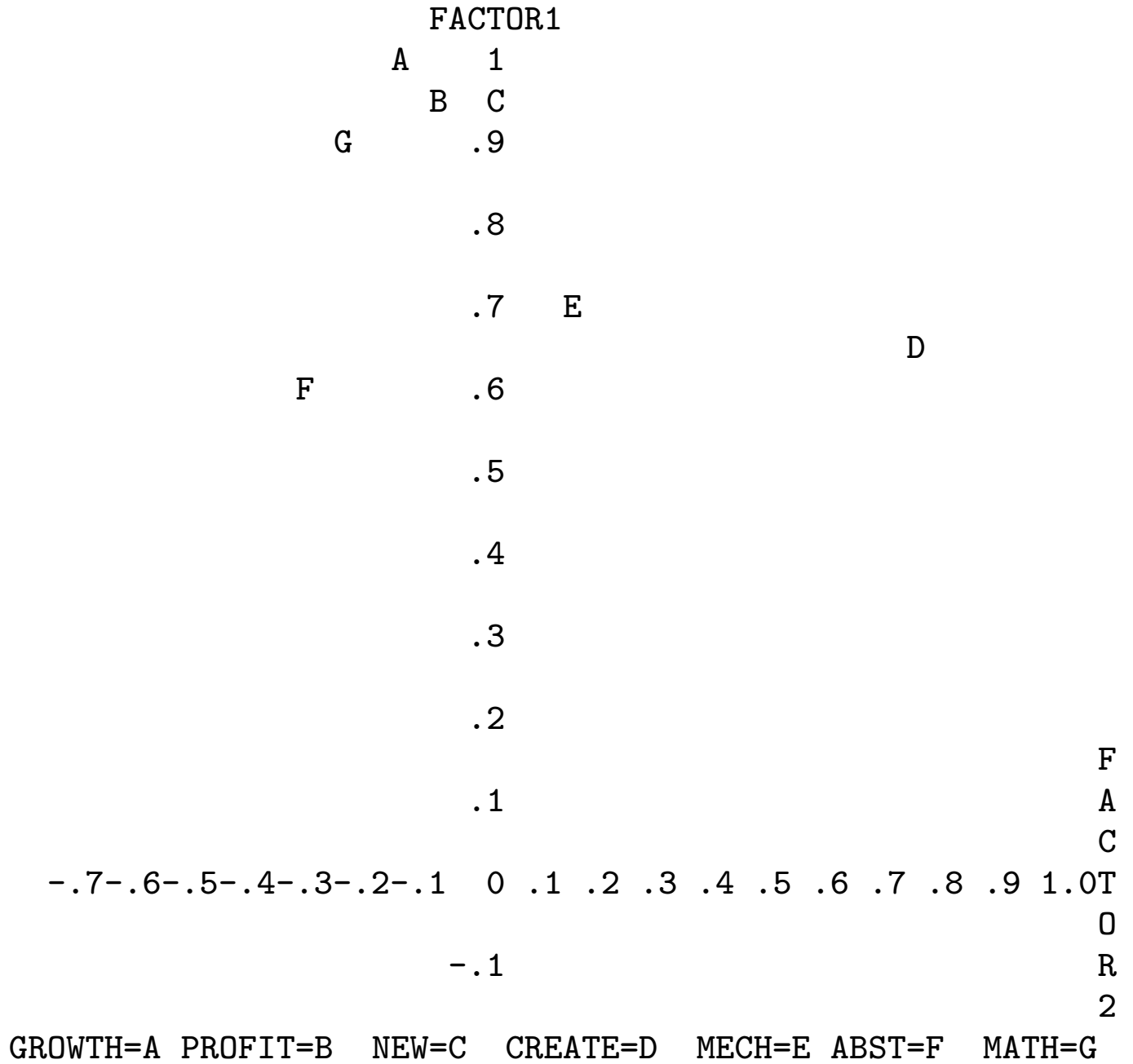
Eigenvectors

	1	2
GROWTH	0.44687	-0.16852
PROFIT	0.42379	-0.08882
NEW	0.42404	0.03803
CREATE	0.30550	0.85228
MECH	0.32820	0.15943
ABST	0.26439	-0.34496
MATH	0.41224	-0.30243

Factor Pattern

	FACTOR1	FACTOR2
GROWTH	0.98704	-0.14753
PROFIT	0.93605	-0.07775
NEW	0.93660	0.03329
CREATE	0.67478	0.74609
MECH	0.72492	0.13956
ABST	0.58396	-0.30198
MATH	0.91054	-0.26475

Plot of Factor Pattern for FACTOR1 and FACTOR2



Rotation Method: Varimax

Orthogonal Transformation Matrix

	1	2
1	0.84716	0.53133
2	-0.53133	0.84716

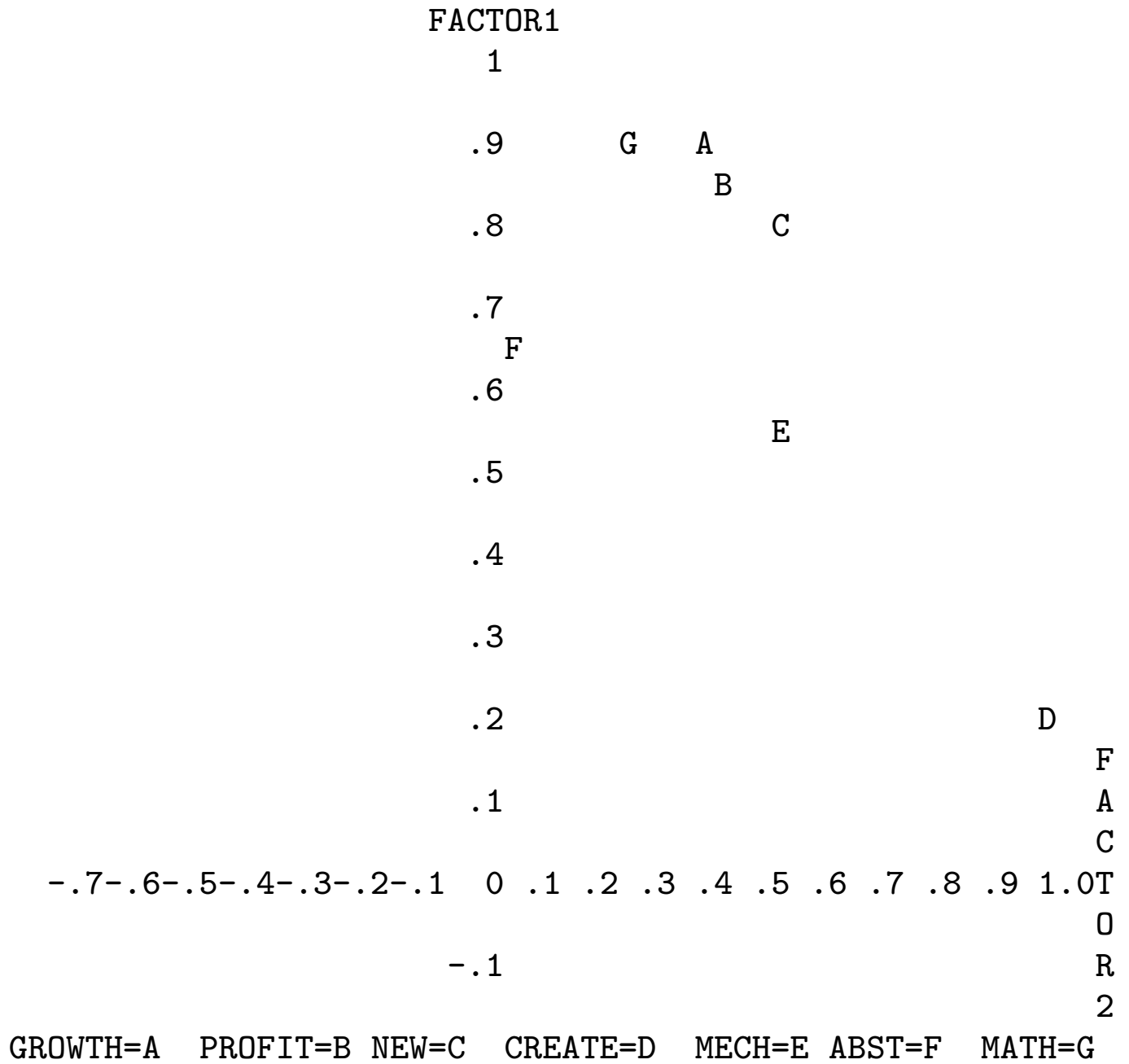
Rotated Factor Pattern

	FACTOR1	FACTOR2
GROWTH	0.91457	0.39946
PROFIT	0.83430	0.43149
NEW	0.77577	0.52585
CREATE	0.17523	0.99059
MECH	0.53997	0.50340
ABST	0.65516	0.05445
MATH	0.91205	0.25951

Variance explained by each factor

FACTOR1	FACTOR2
3.717650	1.927275

Plot of Factor Pattern for FACTOR1 and FACTOR2



The rotated factor loadings seem rather similar here. There seems to be a distinct latent variable which fully explains creativity and is at play in determining other variables a bit. There also seems to be a variable on which every variable except creativity is highly loaded. I am not really sure how to interpret this variable.

Now I tried maximum likelihood.

```
proc factor m=ml nfactor=2 rotate=v
            preplot plot all heywood;
run;
```

The output is

```
Significance tests based on 50 observations:
  Test of H0: No common factors.
      vs HA: At least one common factor.
```

```
Chi-square = 499.661 df = 21 Prob>chi**2 = 0.0001
```

```
Test of H0: 2 Factors are sufficient.
```

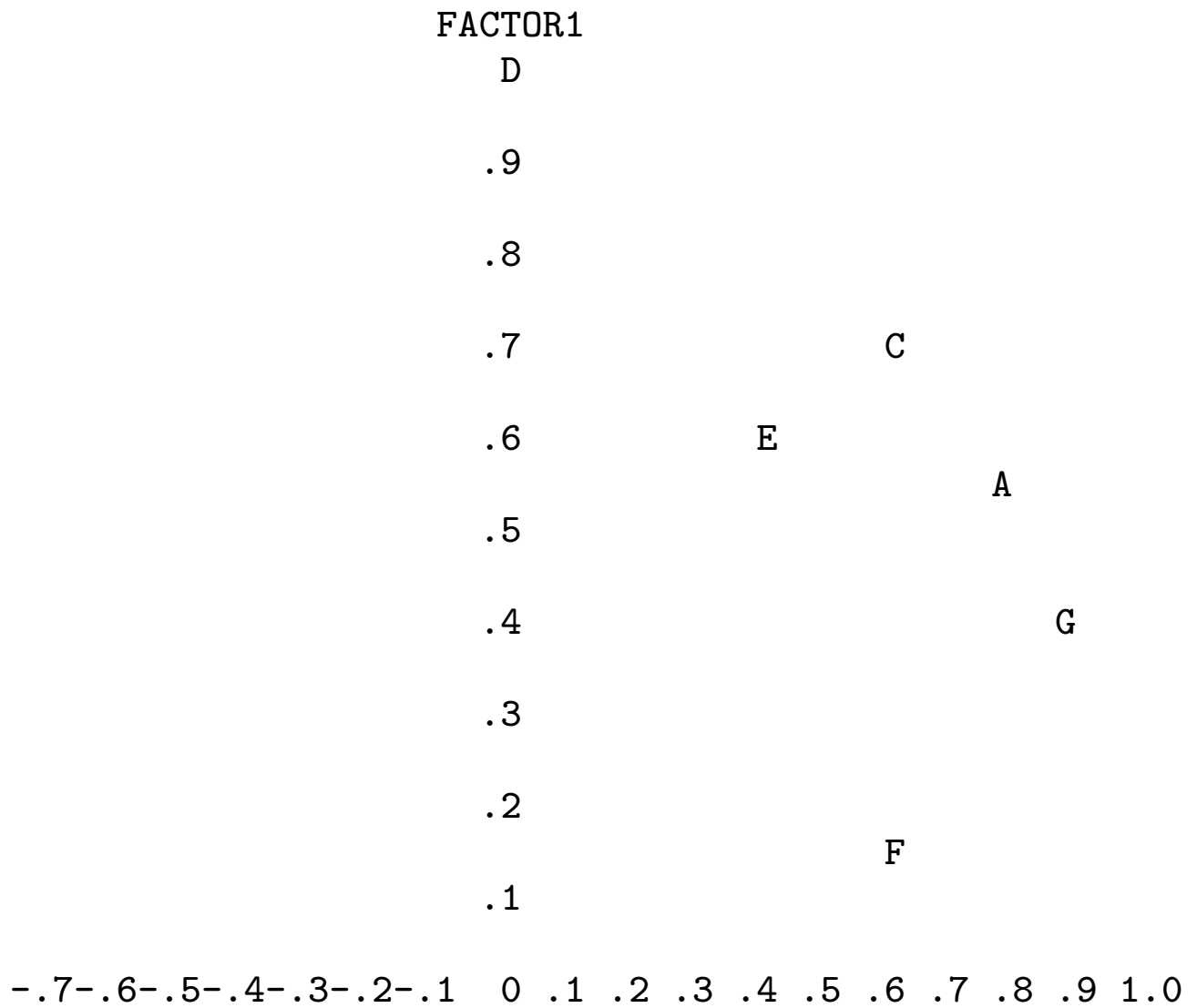
```
      vs HA: More factors are needed.
```

```
Chi-square = 117.092   df = 8   Prob>chi**2 = 0.0001
```

Notice ML conclusion: not enough factors.

	FACTOR1	FACTOR2
GROWTH	0.57204	0.77709
PROFIT	0.54151	0.79768
NEW	0.70036	0.62138
CREATE	1.00000	-0.00000
MECH	0.59074	0.42024
ABST	0.14691	0.60579
MATH	0.41264	0.89462

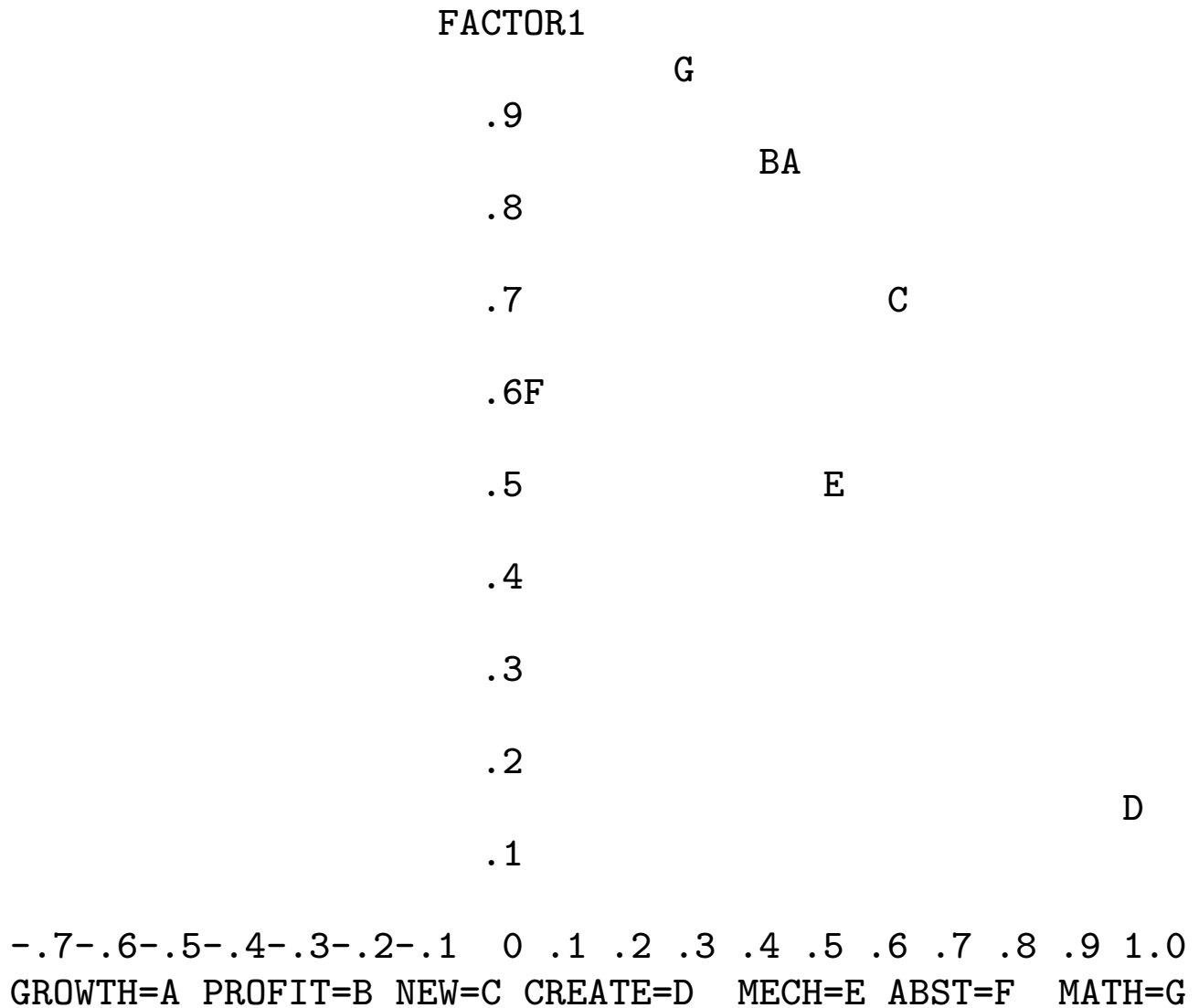
Plot of Factor Pattern for FACTOR1 and FACTOR2



GROWTH=A PROFIT=A NEW=C CREATE=D MECH=E ABST=F MATH=G

Rotation Method: Varimax Rotated Factor Pattern

FACTOR1	FACTOR2	
GROWTH	0.85249	0.45205
PROFIT	0.86839	0.41884
NEW	0.71725	0.60180
CREATE	0.14646	0.98922
MECH	0.50223	0.52282
ABST	0.62078	0.05660
MATH	0.94541	0.27716



Pattern same as for other methods.

Notice however, that this procedure factors S not R .

This explains the totally different eigenvalues and so on.

Finally I ran `proc glm` regressing the sales figures on the psychological test scores.

```
proc glm ;  
  model growth profit new = create mech abst math;  
  manova h=_all_ /printh printe;  
run;
```

The output is:

Dependent Variable: GROWTH

Source	DF	Type III	SS	Mean Sq	F	Pr > F
CREATE	1	61.43176	61.4317	23.53	0.0001	
MECH	1	25.57000	25.5700	9.79	0.0031	
ABST	1	92.58581	92.5858	35.46	0.0001	
MATH	1	548.00956	548.0095	209.87	0.0001	

Dependent Variable: PROFIT

Source	DF	Type III	SS	Mean Sq	F	Pr > F
CREATE	1	6.79727	6.79727	1.80	0.1860	
MECH	1	213.27427	213.27427	56.58	0.0001	
ABST	1	48.81780	48.81780	12.95	0.0008	
MATH	1	1764.23024	1764.23024	468.05	0.0001	

Dependent Variable: NEW

Source	DF	Type III	SS	Mean Sq	F	Pr > F
CREATE	1	153.40826	153.408	92.69	0.0001	
MECH	1	1.84667	1.846	1.12	0.2965	
ABST	1	63.38081	63.380	38.30	0.0001	
MATH	1	150.95169	150.951	91.21	0.0001	

All the variables are very significant predictors of the sales indices.

Type 3 Sums of Squares, which adjust for all other variables in the model are the relevant ones.

They show that from a multivariate point of view no variables can be deleted.

But:in univariate regression for PROFIT can probability drop Creativity while for predicting NEW sales Mechanical reasoning appears unimportant.

All 4 must be retained for prediction of sales growth.

So far as I can see this data set is relatively well understood from this regression output, though the correlation structure is of some interest too.