STAT 804: Lecture 5

Model identification

Model identification for time series X: select values of p,q so that the ARMA(p,q) process gives a reasonable fit to data.

Most important tool: plot of estimated autocorrelation function (ACF) of X.

Before we discuss doing this with real data we explore what plots of the ACF of various ARMA(p,q) plots should look like (in the absence of estimation error).

For an MA(p) process we found that

$$C_X(h) = \begin{cases} \sigma^2 \sum_{j=0}^{p-|h|} b_j b_{j+|h|} & |h| \le p \\ 0 & \text{otherwise} \end{cases}$$

Important *qualitative* feature: vanishes if |h| > p.

For an AR(1) process $X_t - \mu = \rho(X_{t-1} - \mu) + \epsilon_t$ the autocorrelation function is

$$\rho_X(h) = \rho^{|h|}$$

qualitative feature: decreases geometrically.

To derive ACF for general AR(p) we mimic the technique for p = 1.

If
$$X_t = \sum_{1}^{p} a_j X_{t-j} + \epsilon_t$$
 then
$$C_X(h) = \operatorname{Cov}(X_0, X_h)$$

$$= \sum_{j=1}^{p} a_j \operatorname{Cov}(X_0, X_{h-j}) + \operatorname{Cov}(X_0, \epsilon_h)$$

$$= \sum_{j=1}^{p} a_j C_X(h-j)$$

for h > 0.

Divide through by $C_X(0)$.

Remember that $\rho_X(h) = C_X(h)/C_X(0)$ and $\rho_X(-k) = \rho_X(k)$: see that the above recursions for $h = 1, \ldots, p$ are p linear equations in the p unknowns $\rho_X(1), \ldots, \rho_X(p)$.

Called the Yule Walker equations.

For instance, when p = 2 we get

$$C_X(2) = a_1 C_X(1) + a_2 C_X(0)$$

 $C_X(1) = a_1 C_X(0) + a_2 C_X(-1)$

which becomes, after division by $C_X(0)$

$$\rho_X(2) = a_1 \rho_X(1) + a_2$$
$$\rho_X(1) = a_1 + a_2 \rho_X(1)$$

Can use generating functions to get explicit formulas for the $\rho(h)$.

Here simply observe: two equations in two unknowns to solve.

The second equation shows that

$$\rho(1) = \frac{a_1}{1 - a_2}$$

Not possible if $a_2 = 1$ (unless $a_1 = 0$)

Not a correlation for some other (a_1, a_2) pairs (see homework).

The first equation then gives

$$\rho(2) = \frac{a_1^2 + a_2(1 - a_2)}{1 - a_2}$$

Note: can calculate $\rho(h)$ recursively from $\rho(1)$ and $\rho(2)$ for $h \geq 3$ via Yule Walker.

Look at characteristic polynomial $\phi(x)$:

When $a_2 = 1$ we have

$$\phi(x) = 1 - a_1 x - x^2 = (1 - \alpha_1 x)(1 - \alpha_2 x)$$

where $1/\alpha_i$, i = 1, 2 are the two roots.

Multiplying out: $\alpha_1\alpha_2 = -1$ so either:

One of two has modulus more than 1 (root $1/\alpha_i$ has modulus less than 1) or

Both have modulus 1.

Both roots real so would be ± 1 .

Since $\alpha_1 + \alpha_2 = a_1$ (again from multiplying it out and examining the coefficient of x) we would then know $a_1 = 0$. In either case there is no stationary solution.

Qualitative features: can prove solutions of Yule-Walker equations decay to 0 at a geometric rate: $|\rho_X(h)| \le a^{|h|}$ for some $a \in (0,1)$. However, for general p they are not too simple.

Periodic Processes

If Z_1, Z_2 are iid $N(0, \sigma^2)$ then we saw

$$X_t = Z_1 \cos(\omega t) + Z_2 \sin(\omega t)$$

is strictly stationary, mean 0, autocorrelation $\rho(h) = \cos(\omega h)$: perfectly periodic.

Linear Superposition

If X and Y are jointly stationary then Z=aX+bY is stationary and

$$C_Z(h) = a^2 C_X(h) + b^2 C_Y(h) + ab(C_{XY}(h) + C_{YX}(h))$$

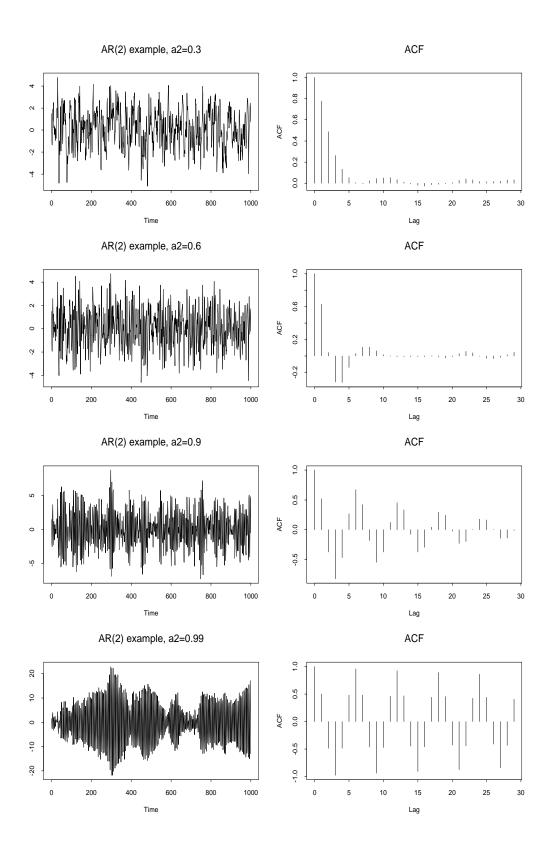
Could hope to recognize periodic component in series by looking for a periodic component in plotted ACF.

Periodic versus AR processes

In fact you can make AR processes which behave very much like periodic processes. Consider the process

$$X_t = X_{t-1} - aX_{t-2} + \epsilon_t$$

Here are graphs of trajectories and autocorrelations for a = 0.3, 0.6, 0.9 and 0.99.



Observe: slow decay of waves in autocovariances, particularly for a near 1.

When a=1 characteristic polynomial is $1-x+x^2$ which has roots

$$\frac{1\pm\sqrt{-3}}{2}$$

Both these roots have modulus 1 so there is no stationary trajectory with a=1. The point is that some AR processes have nearly periodic components.

To get more insight consider the differential equation describing a sine wave:

$$\frac{d^2}{dx^2}f(x) = -\omega^2 f(x);$$

solution is $f(x) = a \sin(\omega x + \phi)$. Replace derivative by differences: get approximation

$$\frac{d^2}{dx^2}f(x) \approx \frac{f(x+h) - 2f(x) + f(x-h)}{h^2}$$

so that

$$\frac{f(x+h) - 2f(x) + f(x-h)}{h^2} \approx -\omega^2 f(x)$$

Take h = 1 and reorganize to get

$$f(x+1) = (2 - \omega^2)f(x) - f(x-1)$$

If we add noise, change notation to t = x + 1 and replace the letter f by X we get

$$X_t = (2 - \omega^2)X_{t-1} - X_{t-2} + \epsilon_t$$

Formalism only; no stationary solution exists.

But, AR(2) processes are at least analogous to solutions of second order differential equations with added noise.

Estimates of C and ρ

In order to identify suitable ARMA models using data we need estimates of C and ρ . If $\mu=0$ is known then

$$C_X(h) = \text{Cov}(X_0, X_h) = \text{Cov}(X_1, X_{h+1}) = \cdots$$

= $\text{E}(X_0 X_h) = \text{E}(X_1 X_{h+1}) = \cdots$.

We would then be motivated to use

$$\widehat{C}(h) = \sum_{0}^{T-1-h} X_t X_{t+h} / T.$$

Average products over all pairs which are h time units apart. When μ is unknown often simply use $\hat{\mu} = \bar{X}$; take

$$\widehat{C}(h) = \sum_{0}^{T-1-h} (X_t - \widehat{\mu})(X_{t+h} - \widehat{\mu})/T$$

Alternative: only T-h terms in the sum

$$\widehat{C}(h) = \sum_{0}^{T-1-h} (X_t - \widehat{\mu})(X_{t+h} - \widehat{\mu})/(T - h).$$

so use

$$\widehat{\rho}(h) = \frac{\widehat{C}(h)}{\widehat{C}(0)}.$$

(Note, however, that when T-h is used in the divisor it is technically possible to get a $\hat{\rho}$ value which exceeds 1.)