

## Transforms of Stochastic Processes

Apply these ideas with  $f$  being stochastic process  $X$ . Several difficulties:

- $X$  is not periodic.
- $X$  often only discrete time function; data always discrete time.
- Even for continuous time  $X$ , Fourier transform integral typically doesn't converge:

$$X \not\rightarrow 0 \text{ as } t \rightarrow \pm\infty.$$

Discrete  $X$  leads to study of discrete approximation to integral:

$$\sum_{t=0}^{T-1} X_t \exp(i2\pi\omega t)$$

This object has real part

$$\sum_{t=0}^{T-1} X_t \cos(2\pi\omega t)$$

and imaginary part

$$\sum_{t=0}^{T-1} X_t \sin(2\pi\omega t).$$

So: apart from means not being 0 studying sample covariance with sines and cosines at frequency  $\omega$ .

Statistical properties and interpretation?

Suppose  $X$  mean 0 stationary time series, autocovariance function  $C$ .

**Def'n:** *discrete Fourier transform* of  $X$  is

$$\hat{X}(\omega) = \frac{1}{\sqrt{T}} \sum_{t=0}^{T-1} X_t \exp(i2\pi\omega t).$$

Division by  $\sqrt{T}$  motivated by recognition that sum of  $T$  terms typically has standard deviation on order of  $\sqrt{T}$ .

So expect SD of  $\hat{X}$  will have reasonable limit as  $T \rightarrow \infty$ .

First compute moments of  $\hat{X}$ .

Moments of complex valued  $\hat{X}$ ?

One way: view  $\hat{X}$  as vector with two components, the real and imaginary parts.

Gives  $\hat{X}$  a mean and a 2 by 2 variance covariance matrix.

Also of interest: expected modulus squared of  $\hat{X}$ ,:

$$E[|\hat{X}(\omega)|^2] = E[\hat{X}(\omega)\overline{\hat{X}(\omega)}]$$

where  $\bar{z}$  is the complex conjugate of  $z$ .

(If  $z = x + iy$  with  $x$  and  $y$  real then  $\bar{z} = x - iy$ .)

Since the  $X$ s have mean 0 we see that

$$E\hat{X}(\omega) = 0$$

Note expected value of complex valued random variable is computed by finding expected value of real and imaginary parts.

Then

$$E[|\hat{X}(\omega)|^2] = \frac{1}{T} \sum_{s=0}^{T-1} \sum_{t=0}^{T-1} \exp(i2\pi\omega(s-t))E(X_s X_t)$$

Expected values are  $C(s - t)$ .

Gather together all terms involving  $C(0)$ , all those involving  $C(1)$ , etc.:

$$\begin{aligned} \mathbb{E}[|\hat{X}(\omega)|^2] = & \frac{1}{T} \} TC(0) \\ & + (T - 1)(e^{i2\pi\omega} + e^{-i2\pi\omega})C(1) + \dots \} \end{aligned}$$

which simplifies to

$$\begin{aligned} C(0) + (1 - 1/T)C(1)(e^{i2\pi\omega} + e^{-i2\pi\omega}) \\ + (1 - 2/T)C(2)(e^{i4\pi\omega} + e^{-i4\pi\omega}) \dots \end{aligned}$$

As  $T \rightarrow \infty$  coefficient of  $C(k)$  converges to 1.

Use  $C(k) = C(-k)$  to see

$$\lim_{T \rightarrow \infty} \mathbb{E}[|\hat{X}(\omega)|^2] = \sum_{-\infty}^{\infty} C(k) \exp(i2\pi\omega k).$$

**Def'n:** *Spectral density, or power spectrum, of  $X$ :*

$$f_X(\omega) = \sum_{-\infty}^{\infty} C(k) \exp(i2\pi\omega k).$$

Interpretations of spectral density and discrete Fourier transform:

- Discrete Fourier transform is rerepresentation of the data: can recover data from transform by inverse transform:

$$\begin{aligned} & \frac{1}{\sqrt{T}} \sum_{k=0}^{T-1} \exp\left(\frac{-i2\pi kt}{T}\right) \hat{X}(k/T) \\ &= \frac{1}{T} \sum_{k=0}^{T-1} \exp\left(\frac{-i2\pi kt}{T}\right) \sum_{s=0}^{T-1} \exp\left(\frac{i2\pi ks}{T}\right) X_s \\ &= \frac{1}{T} \sum_{s=0}^{T-1} X_s \sum_{k=0}^{T-1} \exp\left(\frac{i2\pi k(s-t)}{T}\right) \end{aligned}$$

For  $s = t$  sum over  $k$  is  $T$

For  $s \neq t$  sum can be done as geometric series — get 0.

So inside sum just picks out term  $s = t$  giving  $X_t$  as the inverse transform.

- So DFT decomposes  $X$  into trigonometric functions of various frequencies:  $\hat{X}(k/T)$  is weight on component at frequency  $k/T$ .
- Spectral density is limit of variance of that weight or an approximation to variance of component of  $X$  at frequency  $k/T$ .
- Spectral density is transform of ACF of  $X$ :

$$\int_0^1 f_X(\omega) \exp(-i2\pi\ell\omega) d\omega = C_X(\ell).$$

- Since for any integer  $t \neq 0, \pm T, \pm 2T, \dots$

$$\sum_{k=0}^{T-1} \exp(i2\pi kt/T) = 0$$

we see  $\hat{X}(k/T)$  is, apart from a factor of  $\sqrt{T}$ , a complex number:

- real part is sample covariance between  $X$  and  $\cos(2\pi kt/T)$
- imaginary part is sample covariance between  $X$  and  $\sin(2\pi kt/T)$ .

- Compute covariance between  $X$  and

$$a \cos(2\pi kt/T) + b \sin(2\pi kt/T).$$

Choose  $a, b$  to maximize covariance subject to  $a^2 + b^2 = 1$ .

Resulting coefficients found by multiple regression of  $X_t$  on the cosine and sine.



Since

$$\sum_{k=0}^{T-1} \cos(2\pi kt/T) \sin(2\pi kt/T) = 0$$

can check covariance maximized by taking  $a$  and  $b$  proportional to real and imaginary parts of  $\hat{X}(k/T)$  respectively.

Also: covariance with this linear combination is  $|\hat{X}(k/T)|^2$ .

Calculation requires  $t$  to be a non-zero integer.

In practice apply techniques to  $X - \bar{X}$ .

- Later: if series  $Y$  is a filtered version of series  $X$  then spectral densities have simple relation to one another in terms of some property of filter.

Can use this fact to estimate the filter itself when this is unknown.

## The Periodogram

Sample covariance between  $X$  and  $\sin(2\pi\omega t + \phi)$  is

$$\frac{1}{T} \sum_{t=0}^{T-1} X_t \sin(2\pi\omega t + \phi) - \bar{X} \frac{1}{T} \sum_{t=0}^{T-1} \sin(2\pi\omega t + \phi)$$

Use identity  $\sin(\theta) = (e^{i\theta} - e^{-i\theta})/(2i)$  and formulas for geometric sums to compute mean.

When  $\omega = k/T$  for an integer  $k$ , not 0, we find that  $\sum_{t=0}^{T-1} \sin(2\pi\omega t + \phi) = 0$ .

So sample covariance is simply

$$\frac{1}{T} \sum_{t=0}^{T-1} X_t \sin(2\pi\omega t + \phi).$$

For these special  $\omega$  we can also compute

$$\sum_{t=0}^{T-1} \sin^2(2\pi\omega t + \phi) = T/2.$$

So correlation between  $X$  and  $\sin(2\pi\omega t + \phi)$  is

$$\frac{\frac{1}{T} \sum_{t=0}^{T-1} X_t \sin(2\pi\omega t + \phi)}{s_x \sqrt{1/2}}$$

where  $s_x^2$  is sample variance  $\sum(X_t - \bar{X})^2/T$ .

Adjust  $\phi$  to maximize this correlation.

The sine can be rewritten as

$$\cos(\phi) \sin(2\pi\omega t) + \sin(\phi) \cos(2\pi\omega t)$$

so choose coefficients  $a$  and  $b$  to maximize correlation between  $X$  and

$$a \sin(2\pi\omega t) + b \cos(2\pi\omega t)$$

subject to the condition  $a^2 + b^2 = 1$ .

Correlations are scale invariant so drop condition on  $a$  and  $b$  and maximize the correlation between  $X$  and the linear combination of sine and cosine.

Problem solved by linear regression. Coefficients given by  $(M^T M)^{-1} M^T X$ :

$M$  is  $T$  by 2 design matrix full of sines and cosines.

Get  $M^T M = \frac{T}{2} I_{T \times T}$ ; regression coefficients are

$$a = \frac{2}{T} \sum_{t=0}^{T-1} X_t \sin(2\pi\omega t)$$

and

$$b = \frac{2}{T} \sum_{t=0}^{T-1} X_t \cos(2\pi\omega t).$$

Covariance between  $X$  and best linear combination is

$$\begin{aligned} \frac{1}{T} \left\{ a \sum_{t=0}^{T-1} X_t \sin(2\pi\omega t) + b \sum_{t=0}^{T-1} X_t \cos(2\pi\omega t) \right\} \\ = (a^2 + b^2)/2. \end{aligned}$$

But in fact

$$a^2 + b^2 = \left| \frac{1}{T} \sum_{t=0}^{T-1} X_t \exp(2\pi\omega t i) \right|^2$$

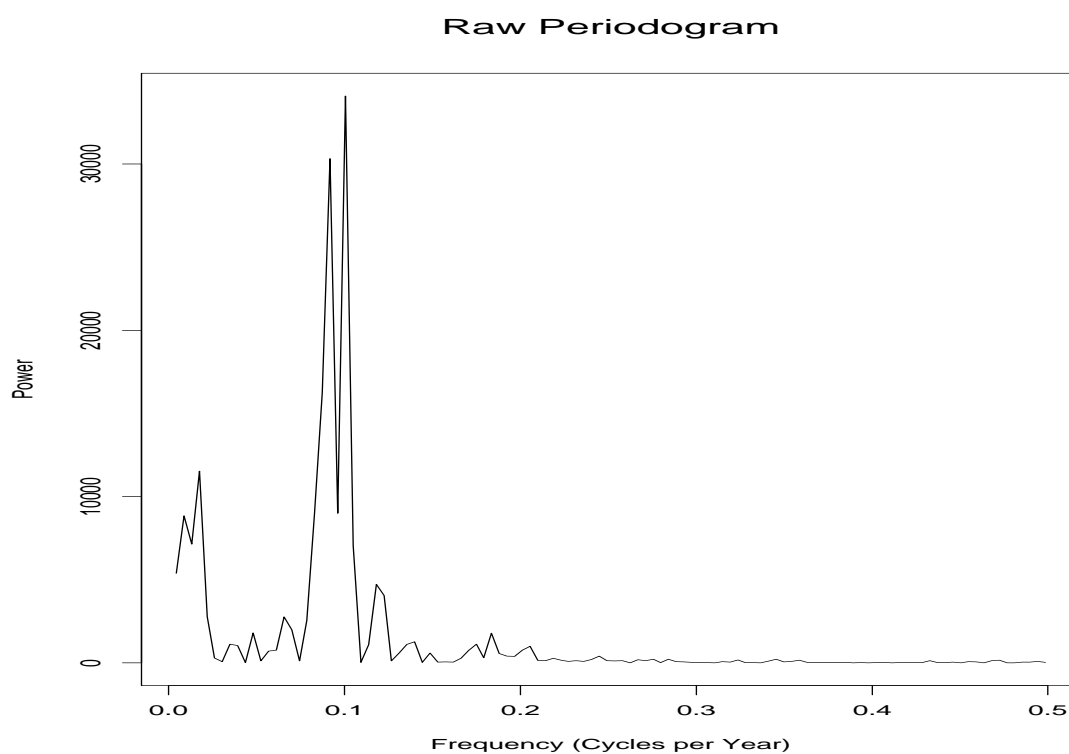
is modulus of DFT  $\hat{X}(\omega)$  divided by  $T$ .

**Def'n: Periodogram** is function

$$|\hat{X}(\omega)|^2$$

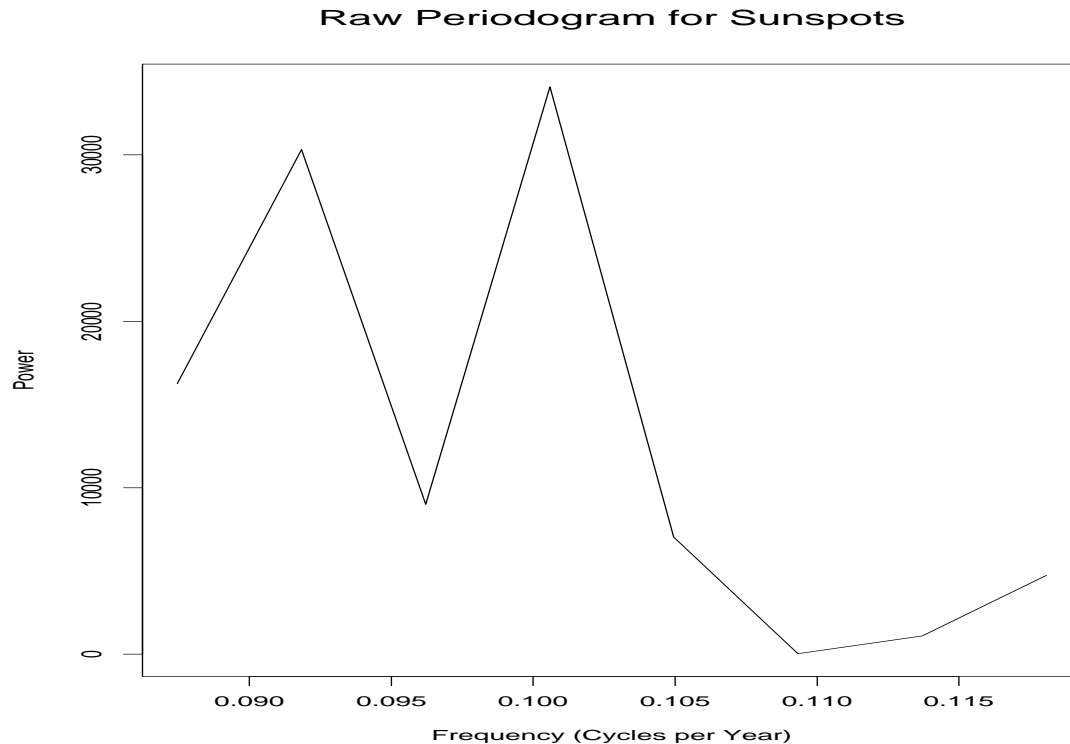
Some periodogram plots:

- $|\hat{X}|$  vs frequency for sunspots minus mean



Notice peak at frequency slightly below 0.1 cycles per year as well as peak at frequency close to 0.03.

Plot only for frequencies from  $1/12$  to  $1/8$  which should include the largest peak.



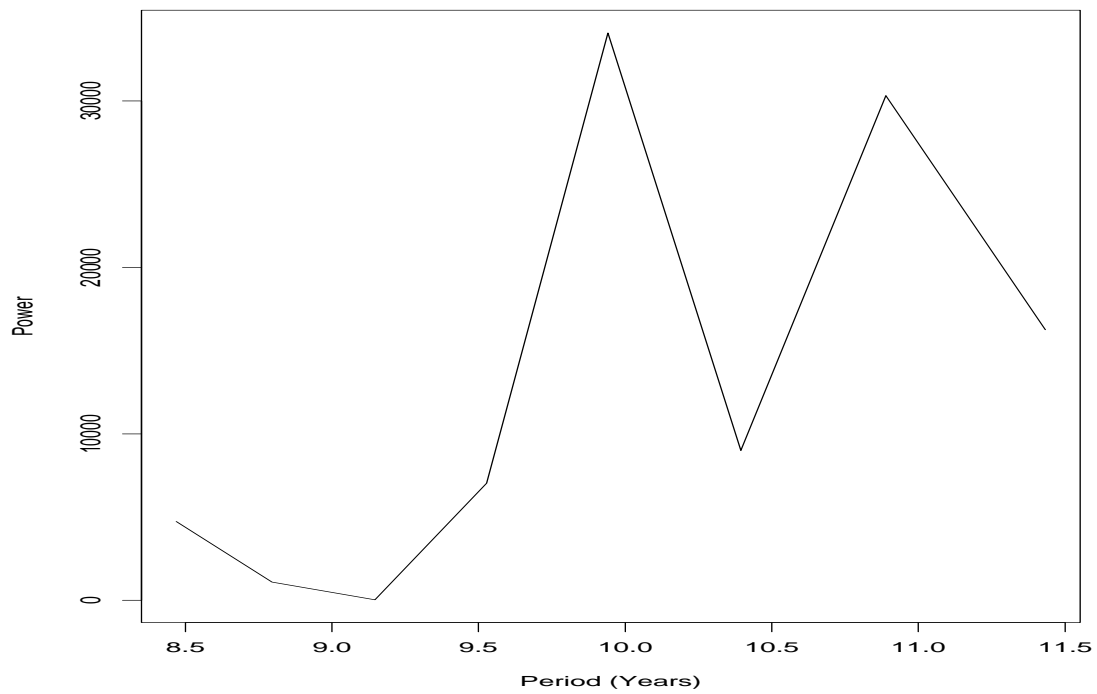
Notice: picture clearly piecewise linear.

Actually using DFT: computes sample spectrum only at frequencies of form  $k/T$  (in cycles per point) for integer values of  $K$ .

There are only about 10 points on this plot.

Same plot against period ( $= 1/\omega$ ) shows peaks just below 10 years and just below 11.

Raw Periodogram for Sunspots



DFT can be computed very quickly at special frequencies but to see structure clearly near a peak need to compute  $\hat{X}(\omega)$  for a denser grid of  $\omega$ .

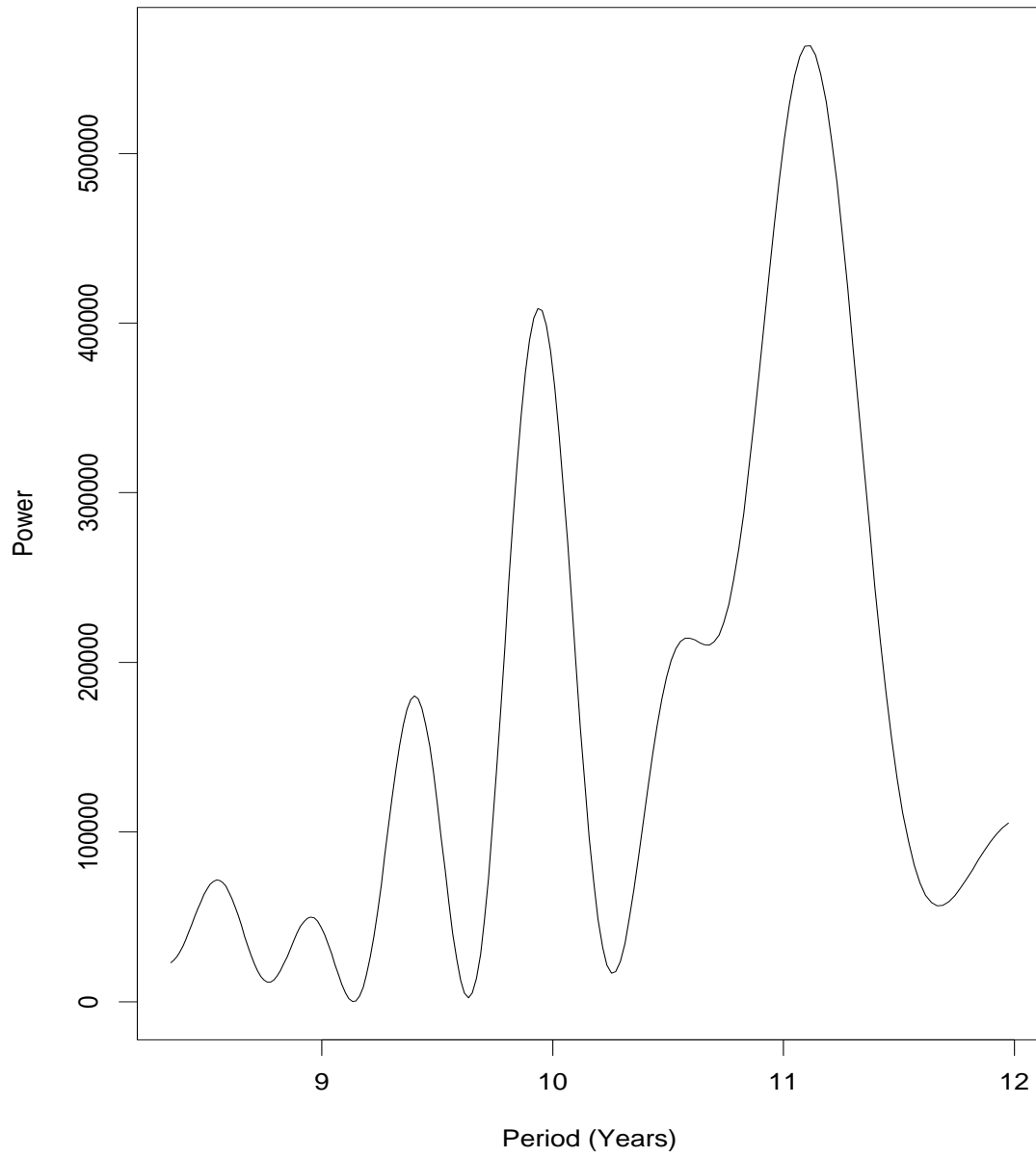
Use S-Plus function

```
transform<- function(x, a, b, n = 100)
{
  f <- seq(a, b, length = n)
  nn <- 1:length(x)
  args <- outer(f, nn, "*") * 2 * pi
  cosines <- t(cos(t(args))) * x
  sines <- t(sin(t(args))) * x
  one <- rep(1, length(x))
  ((cosines %*% one)^2
   + (sines %*% one)^2)/length(x)
}
```

to compute lots of values for periods between 8 and 12 years.



Plot of Spectrum vs Period for Sunspots



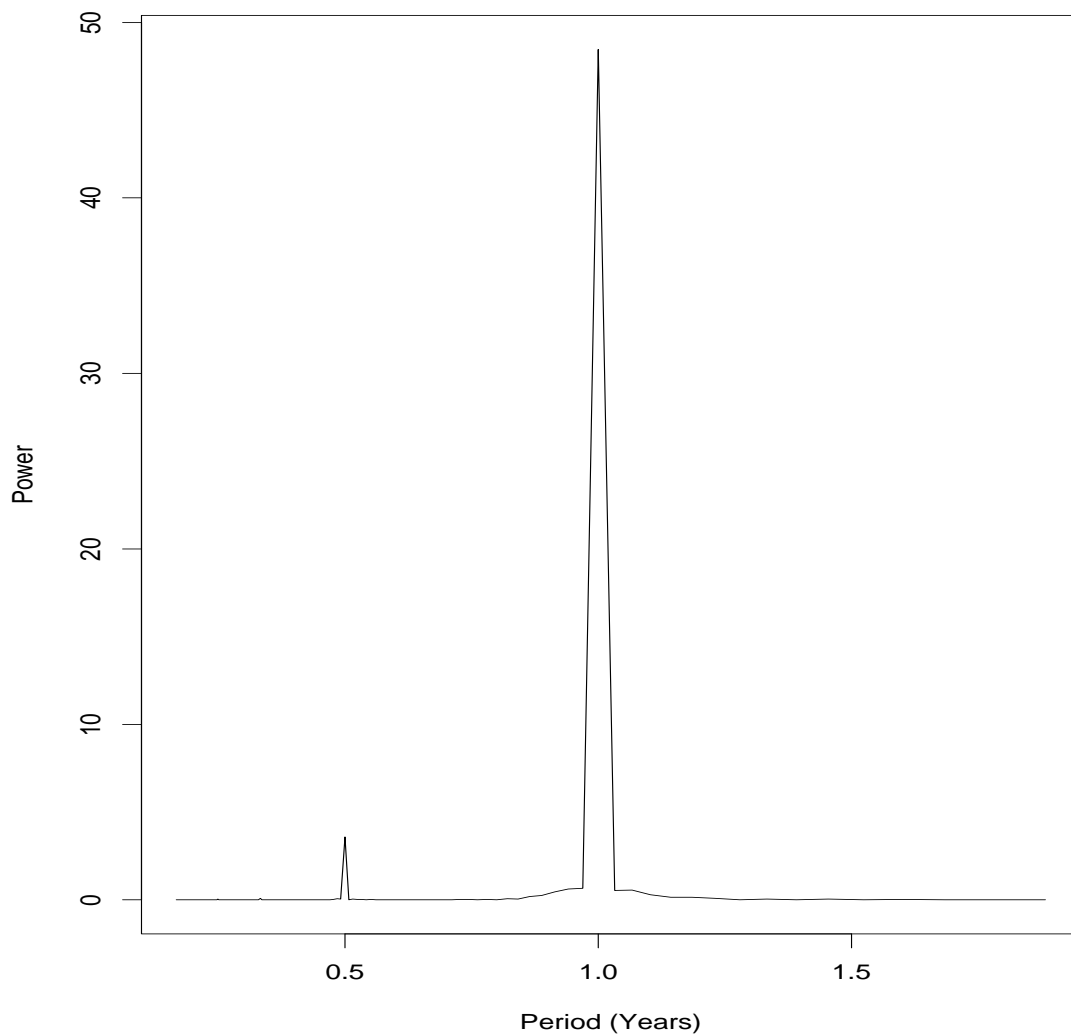
Periodogram for CO<sub>2</sub> above Mauna Loa: Linear trend removed by linear regression.

Note peaks at periods of 1 year and 6 months.

Peaks show clear annual cycle.

Annual cycle not simple sine wave – contains overtones: components whose frequency is integer multiple of basic frequency of 1 cycle per year.

Power Spectrum against Period  
CO2 Conc above Mauna Loa detrended



Now a detail of this image:

