

STAT 830

Expectation

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What I assume you already know

- Continuous case: $E(X) = \int_{-\infty}^{\infty} xf(x) dx$.
- Discrete case: $E(X) = \sum_x xf(x)$
- Mean, variance, standard deviation, covariance, correlation.
- Conditional analogues of above.



What I want you to learn

- Abstract definition of $E(X)$.
- Dominated, monotone convergence theorems.
- Mean, variance, standard deviation, covariance, correlation.
- Define conditional expectation and relation to conditional density.
- Give some properties of conditional expectation.



- Two elementary definitions of expected values:
- **Def'n:** If X has density f then

$$E\{g(X)\} = \int g(x)f(x) dx.$$

- **Def'n:** If X has discrete density f then

$$E\{g(X)\} = \sum_x g(x)f(x).$$

- **FACT:** if $Y = g(X)$ for a smooth g

$$\begin{aligned} E(Y) &= \int yf_Y(y) dy = \int g(x)f_Y(g(x))g'(x) dx \\ &= E\{g(X)\} \end{aligned}$$

by change of variables formula for integration.

- Good: otherwise might have two different values for $E(e^X)$.



- There are random variables which are neither absolutely continuous nor discrete.
- **Def'n:** RV X is simple if we can write

$$X(\omega) = \sum_1^n a_i 1(\omega \in A_i)$$

for some constants a_1, \dots, a_n and events A_i .

- **Def'n:** For a simple rv X define

$$E(X) = \sum a_i P(A_i).$$

- For positive random variables which are not simple extend definition by approximation:
- **Def'n:** If $X \geq 0$ then

$$E(X) = \sup\{E(Y) : 0 \leq Y \leq X, Y \text{ simple}\}.$$



Integrable rvs

- **Def'n:** X is **integrable** if

$$E(|X|) < \infty.$$

- In this case we define

$$E(X) = E\{\max(X, 0)\} - E\{\max(-X, 0)\}.$$

- **Facts:** E is a linear, monotone, positive operator:
 - 1 **Linear:** $E(aX + bY) = aE(X) + bE(Y)$ provided X and Y are integrable.
 - 2 **Positive:** $P(X \geq 0) = 1$ implies $E(X) \geq 0$.
 - 3 **Monotone:** $P(X \geq Y) = 1$ and X, Y integrable implies $E(X) \geq E(Y)$.



Convergence Theorems

- Major technical theorems:
- **Monotone Convergence:** If $0 \leq X_1 \leq X_2 \leq \dots$ and $X = \lim X_n$ (which has to exist) then

$$E(X) = \lim_{n \rightarrow \infty} E(X_n).$$

- **Dominated Convergence:** If $|X_n| \leq Y_n$ and \exists rv X such that $X_n \rightarrow X$ (technical details of this convergence later in the course) and a random variable Y such that $Y_n \rightarrow Y$ with $E(Y_n) \rightarrow E(Y) < \infty$ then

$$E(X_n) \rightarrow E(X).$$

- Often used with all Y_n the same rv Y .
- These theorems are used in *approximation*.



Connection to integration

Theorem

With this definition of E :

- ① if X has density $f(x)$ (even in R^p say) and $Y = g(X)$ then

$$E(Y) = \int g(x)f(x)dx.$$

(Could be a multiple integral.)

- ② If X has pmf f then

$$E(Y) = \sum_x g(x)f(x).$$

- ③ First conclusion works, e.g., even if X has a density but Y doesn't.



- **Def'n:** The r^{th} moment (about the origin) of a real rv X is $\mu'_r = E(X^r)$ (provided it exists).
- We generally use μ for $E(X)$.
- **Def'n:** The r^{th} central moment is

$$\mu_r = E[(X - \mu)^r]$$

- We call $\sigma^2 = \mu_2$ the variance.
- **Def'n:** For an R^p valued random vector X

$$\mu_X = E(X)$$

is the vector whose i^{th} entry is $E(X_i)$ (provided all entries exist).

- **Def'n:** The $(p \times p)$ variance covariance matrix of X is

$$\text{Var}(X) = E [(X - \mu)(X - \mu)^t]$$

which exists provided each component X_i has a finite second moment



Moments and independence

Theorem

If X_1, \dots, X_p are independent and each X_i is integrable then $X = X_1 \cdots X_p$ is integrable and

$$E(X_1 \cdots X_p) = E(X_1) \cdots E(X_p).$$



Proof

Suppose each X_i is simple:

$$X_i = \sum_j x_{ij} 1(X_i = x_{ij})$$

where the x_{ij} are the possible values of X_i . Then

$$\begin{aligned} E(X_1 \cdots X_p) &= \sum_{j_1 \cdots j_p} x_{1j_1} \cdots x_{pj_p} E(1(X_1 = x_{1j_1}) \cdots 1(X_p = x_{pj_p})) \\ &= \sum_{j_1 \cdots j_p} x_{1j_1} \cdots x_{pj_p} P(X_1 = x_{1j_1} \cdots X_p = x_{pj_p}) \\ &= \sum_{j_1 \cdots j_p} x_{1j_1} \cdots x_{pj_p} P(X_1 = x_{1j_1}) \cdots P(X_p = x_{pj_p}) \\ &= \sum_{j_1} x_{1j_1} P(X_1 = x_{1j_1}) \cdots \sum_{j_p} x_{pj_p} P(X_p = x_{pj_p}) \\ &= \prod E(X_i). \end{aligned}$$



General Case

- General $X_i \geq 0$:
- Let X_{in} be X_i rounded down to nearest multiple of 2^{-n} (to maximum of n).
- That is: if

$$\frac{k}{2^n} \leq X_i < \frac{k+1}{2^n}$$

then $X_{in} = k/2^n$ for $k = 0, \dots, n2^n$.

- For $X_i > n$ put $X_{in} = n$.
- Apply case just done:

$$E\left(\prod X_{in}\right) = \prod E(X_{in}).$$

- Monotone convergence applies to both sides.
- General case: write each X_i as difference of positive and negative parts:

$$X_i = \max(X_i, 0) - \max(-X_i, 0).$$

- Apply positive case.



- Abstract definition of conditional expectation is:
- **Def'n:** $E(Y|X)$ is any function of X such that

$$E[R(X)E(Y|X)] = E[R(X)Y]$$

for any bounded function $R(X)$.

- **Def'n:** $E(Y|X = x)$ is a function $g(x)$ such that

$$g(X) = E(Y|X)$$

- **Fact:** If X, Y has joint density $f_{X,Y}(x, y)$ and conditional density $f(y|x)$ then

$$g(x) = \int yf(y|x)dy$$

satisfies these definitions.



Proof

$$\begin{aligned} E(R(X)g(X)) &= \int R(x)g(x)f_X(x)dx \\ &= \int R(x) \int yf(y|x)dyf_X(x)dx \\ &= \int \int R(x)yf_X(x)f(y|x)dydx \\ &= \int \int R(x)yf_{X,Y}(x,y)dydx \\ &= E(R(X)Y) \end{aligned}$$



Interpretation of conditional expectation

- **Intuition:** Think of $E(Y|X)$ as average Y holding X fixed.
- Behaves like ordinary expected value but functions of X only are like constants:

$$E\left(\sum A_i(X)Y_i|X\right) = \sum A_i(X)E(Y_i|X)$$

- Statement called Adam's law by Jerzy Neyman – he used to say it comes before all the others:

$$E[E(Y|X)] = E(Y)$$

which is just the definition of $E(Y|X)$ with $R(X) \equiv 1$.

- In regression courses we say that the total sum of squares is the sum of the regression sum of squares plus the residual sum of squares:

$$\text{Var}(Y) = \text{Var}(E(Y|X)) + E[\text{Var}(Y|X)]$$

- The conditional variance means

$$\text{Var}(Y|X) = E[(Y - E(Y|X))^2|X].$$

