

# STAT 801=830

## Distribution Theory

Richard Lockhart

Simon Fraser University

STAT 801=830 — Fall 2012



# What I assume you already know

- The change of variables for integration
- How to find densities for transformations in  $\mathbb{R}$ .
- How to compute probabilities by doing multiple integrals.
- Material on slides 4–6, 9–13.



- Basic Problem: Start with assumptions about  $f$  or CDF of random vector  $X = (X_1, \dots, X_p)$ .
- Define  $Y = g(X_1, \dots, X_p)$  to be some function of  $X$  (usually some statistic of interest).
- How can we compute the distribution or CDF or density of  $Y$ ?



# Univariate Techniques

Method 1: compute the CDF by integration and differentiate to find  $f_Y$ .

Example:  $U \sim \text{Uniform}[0, 1]$  and  $Y = -\log U$ .

$$\begin{aligned} F_Y(y) &= P(Y \leq y) = P(-\log U \leq y) \\ &= P(\log U \geq -y) = P(U \geq e^{-y}) \\ &= \begin{cases} 1 - e^{-y} & y > 0 \\ 0 & y \leq 0. \end{cases} \end{aligned}$$

so  $Y$  has standard exponential distribution.



# Chi-square

- Example:  $Z \sim N(0, 1)$ , i.e.

$$f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2}$$

and  $Y = Z^2$ .

- Then

$$\begin{aligned} F_Y(y) &= P(Z^2 \leq y) \\ &= \begin{cases} 0 & y < 0 \\ P(-\sqrt{y} \leq Z \leq \sqrt{y}) & y \geq 0. \end{cases} \end{aligned}$$

- Now differentiate

$$P(-\sqrt{y} \leq Z \leq \sqrt{y}) = F_Z(\sqrt{y}) - F_Z(-\sqrt{y})$$

to get

$$f_Y(y) = \begin{cases} 0 & y < 0 \\ \frac{d}{dy} [F_Z(\sqrt{y}) - F_Z(-\sqrt{y})] & y > 0 \\ \text{undefined} & y = 0. \end{cases}$$



## More

- Then

$$\begin{aligned}\frac{d}{dy}F_Z(\sqrt{y}) &= f_Z(\sqrt{y})\frac{d}{dy}\sqrt{y} \\ &= \frac{1}{\sqrt{2\pi}}\exp\left(-(\sqrt{y})^2/2\right)\frac{1}{2}y^{-1/2} \\ &= \frac{1}{2\sqrt{2\pi y}}e^{-y/2}.\end{aligned}$$

- (Similar formula for other derivative.) Thus

$$f_Y(y) = \begin{cases} \frac{1}{\sqrt{2\pi y}}e^{-y/2} & y > 0 \\ 0 & y < 0 \\ \text{undefined} & y = 0. \end{cases}$$



# Indicators

- We will find **indicator** notation useful:

$$1(y > 0) = \begin{cases} 1 & y > 0 \\ 0 & y \leq 0 \end{cases}$$

which we use to write

$$f_Y(y) = \frac{1}{\sqrt{2\pi y}} e^{-y/2} 1(y > 0)$$

(changing definition unimportantly at  $y = 0$ ).



# Fundamental Theorem of Calculus

- **Notice:** I never evaluated  $F_Y$  before differentiating it.
- In fact  $F_Y$  and  $F_Z$  are integrals I can't do but I can differentiate then anyway.
- Remember fundamental theorem of calculus:

$$\frac{d}{dx} \int_a^x f(y) dy = f(x)$$

at any  $x$  where  $f$  is continuous.

- Summary: for  $Y = g(X)$  with  $X$  and  $Y$  each real valued

$$\begin{aligned} P(Y \leq y) &= P(g(X) \leq y) \\ &= P(X \in g^{-1}(-\infty, y]) . \end{aligned}$$

- Take  $d/dy$  to compute the density

$$f_Y(y) = \frac{d}{dy} \int_{\{x: g(x) \leq y\}} f_X(x) dx .$$

- Often can differentiate without doing integral.





## Method 2: Change of variables

- Assume  $g$  is one to one.
- I do:  $g$  is increasing and differentiable.
- Interpretation of density (based on density =  $F'$ ):

$$\begin{aligned}f_Y(y) &= \lim_{\delta y \rightarrow 0} \frac{P(y \leq Y \leq y + \delta y)}{\delta y} \\&= \lim_{\delta y \rightarrow 0} \frac{F_Y(y + \delta y) - F_Y(y)}{\delta y}\end{aligned}$$

and

$$f_X(x) = \lim_{\delta x \rightarrow 0} \frac{P(x \leq X \leq x + \delta x)}{\delta x}.$$

- Now assume  $y = g(x)$ . Define  $\delta y$  by  $y + \delta y = g(x + \delta x)$ .
- Then

$$P(y \leq Y \leq g(x + \delta x)) = P(x \leq X \leq x + \delta x).$$



# Change of Variables Continued

- Get

$$\frac{P(y \leq Y \leq y + \delta y))}{\delta y} = \frac{P(x \leq X \leq x + \delta x)/\delta x}{\{g(x + \delta x) - y\}/\delta x}.$$

- Take limit to get

$$f_Y(y) = f_X(x)/g'(x) \text{ or } f_Y(g(x))g'(x) = f_X(x).$$



## Alternative view

- Each probability is integral of a density.
- First is integral of  $f_Y$  from  $y = g(x)$  to  $y = g(x + \delta x)$ .
- The interval is narrow so  $f_Y$  is nearly constant and

$$P(y \leq Y \leq g(x + \delta x)) \approx f_Y(y)(g(x + \delta x) - g(x)).$$

- Since  $g$  has a derivative  $g(x + \delta x) - g(x) \approx \delta x g'(x)$  so we get

$$P(y \leq Y \leq g(x + \delta x)) \approx f_Y(y)g'(x)\delta x.$$

- Same idea applied to  $P(x \leq X \leq x + \delta x)$  gives

$$P(x \leq X \leq x + \delta x) \approx f_X(x)\delta x$$

so that

$$f_Y(y)g'(x)\delta x \approx f_X(x)\delta x$$

or, cancelling the  $\delta x$  in the limit

$$f_Y(y)g'(x) = f_X(x).$$



## Intution continued

- If you remember  $y = g(x)$  then you get

$$f_X(x) = f_Y(g(x))g'(x).$$

- Or solve  $y = g(x)$  to get  $x$  in terms of  $y$ , that is,  $x = g^{-1}(y)$  and then

$$f_Y(y) = f_X(g^{-1}(y))/g'(g^{-1}(y)).$$

- **This is just the change of variables formula for doing integrals.**
- **Remark:** For  $g$  decreasing  $g' < 0$  but Then the interval  $(g(x), g(x + \delta x))$  is really  $(g(x + \delta x), g(x))$  so that  $g(x) - g(x + \delta x) \approx -g'(x)\delta x$ .
- In both cases this amounts to the formula

$$f_X(x) = f_Y(g(x))|g'(x)|.$$

- **Mnemonic:**

$$f_Y(y)dy = f_X(x)dx.$$



- $X \sim \text{Weibull}(\text{shape } \alpha, \text{scale } \beta)$  or

$$f_X(x) = \frac{\alpha}{\beta} \left( \frac{x}{\beta} \right)^{\alpha-1} \exp \{ -(x/\beta)^\alpha \} 1(x > 0).$$

- Let  $Y = \log X$  or  $g(x) = \log(x)$ .
- Solve  $y = \log x$ :  $x = \exp(y)$  or  $g^{-1}(y) = e^y$ .
- Then  $g'(x) = 1/x$  and  $1/g'(g^{-1}(y)) = 1/(1/e^y) = e^y$ .
- Hence

$$f_Y(y) = \frac{\alpha}{\beta} \left( \frac{e^y}{\beta} \right)^{\alpha-1} \exp \{ -(e^y/\beta)^\alpha \} 1(e^y > 0) e^y.$$

- For any  $y$ ,  $e^y > 0$  so indicator = 1. So

$$f_Y(y) = \frac{\alpha}{\beta^\alpha} \exp \{ \alpha y - e^{\alpha y} / \beta^\alpha \}.$$



## Example Continued

- Define  $\phi = \log \beta$  and  $\theta = 1/\alpha$ ; then,

$$f_Y(y) = \frac{1}{\theta} \exp \left\{ \frac{y - \phi}{\theta} - \exp \left\{ \frac{y - \phi}{\theta} \right\} \right\} .$$

- Extreme Value** density with **location** parameter  $\phi$  and **scale** parameter  $\theta$ .
- Note: several distributions are called Extreme Value.



Simplest multivariate problem:

$$X = (X_1, \dots, X_p), \quad Y = X_1$$

(or in general  $Y$  is any  $X_j$ ).

If  $X$  has density  $f(x_1, \dots, x_p)$  and  $q < p$  then  $Y = (X_1, \dots, X_q)$  has density

$$f_Y(x_1, \dots, x_q) = \int_{-\infty}^{\infty} \cdots \int_{-\infty}^{\infty} f(x_1, \dots, x_p) dx_{q+1} \dots dx_p.$$

$f_{X_1, \dots, X_q}$  is the **marginal** density of  $X_1, \dots, X_q$  and  $f_X$  the **joint** density of  $X$  but they are both just densities.

“Marginal” just to distinguish from the joint density of  $X$ .



## Example

- The function  $f(x_1, x_2) = Kx_1x_21(x_1 > 0, x_2 > 0, x_1 + x_2 < 1)$  is a density provided

$$P(X \in R^2) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x_1, x_2) dx_1 dx_2 = 1.$$

- The integral is

$$\begin{aligned} K \int_0^1 \int_0^{1-x_1} x_1 x_2 dx_1 dx_2 &= K \int_0^1 x_1 (1 - x_1)^2 dx_1 / 2 \\ &= K(1/2 - 2/3 + 1/4)/2 = K/24 \end{aligned}$$

so  $K = 24$ .

- The marginal density of  $X_1$  is Beta(2, 3):

$$\begin{aligned} f_{X_1}(x_1) &= \int_{-\infty}^{\infty} 24x_1x_21(x_1 > 0, x_2 > 0, x_1 + x_2 < 1) dx_2 \\ &= 24 \int_0^{1-x_1} x_1x_21(0 < x_1 < 1) dx_2 \\ &= 12x_1(1 - x_1)^21(0 < x_1 < 1). \end{aligned}$$





# General Problem

- General problem has  $Y = (Y_1, \dots, Y_q)$  with  $Y_i = g_i(X_1, \dots, X_p)$ .
- **Case 1:**  $q > p$ .  $Y$  **won't** have density for “smooth”  $g$ .  $Y$  will have a **singular** or discrete distribution.
- Problem rarely of real interest. (But, e.g., residuals have singular distribution.)
- **Case 2:**  $q = p$ . Use multivariate change of variables formula.
- **Case 3:**  $q < p$ . Pad out  $Y$ —add on  $p - q$  more variables (carefully chosen) say  $Y_{q+1}, \dots, Y_p$ . Find functions  $g_{q+1}, \dots, g_p$ . Define for  $q < i \leq p$ ,  $Y_i = g_i(X_1, \dots, X_p)$  and  $Z = (Y_1, \dots, Y_p)$ .
- Choose  $g_i$  so that we can use change of variables on  $g = (g_1, \dots, g_p)$  to compute  $f_Z$ . Find  $f_Y$  by integration:

$$f_Y(y_1, \dots, y_q) = \int_{-\infty}^{\infty} \cdots \int_{-\infty}^{\infty} f_Z(y_1, \dots, y_q, z_{q+1}, \dots, z_p) dz_{q+1} \cdots dz_p$$



# Multivariate Change of Variables

- Suppose  $Y = g(X) \in R^p$  with  $X \in R^p$  having density  $f_X$ .
- **Assume  $g$  is a one to one (“injective”) map**, i.e.,  $g(x_1) = g(x_2)$  if and only if  $x_1 = x_2$ .
- Find  $f_Y$ :
- noindent Step 1: Solve for  $x$  in terms of  $y$ :  $x = g^{-1}(y)$ .
- Step 2: Use basic equation:

$$f_Y(y)dy = f_X(x)dx$$

and rewrite it in the form

$$f_Y(y) = f_X(g^{-1}(y)) \frac{dx}{dy}.$$

- Interpretation of derivative  $\frac{dx}{dy}$  when  $p > 1$ :

$$\frac{dx}{dy} = \left| \det \left( \frac{\partial x_i}{\partial y_j} \right) \right|$$

which is the so called **Jacobian**.



# Multivariate Change of Variables

- Equivalent formula inverts the matrix:

$$f_Y(y) = \frac{f_X(g^{-1}(y))}{\left| \frac{dy}{dx} \right|}$$

- This notation means

$$\left| \frac{dy}{dx} \right| = \left| \det \begin{bmatrix} \frac{\partial y_1}{\partial x_1} & \frac{\partial y_1}{\partial x_2} & \cdots & \frac{\partial y_1}{\partial x_p} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial y_p}{\partial x_1} & \frac{\partial y_p}{\partial x_2} & \cdots & \frac{\partial y_p}{\partial x_p} \end{bmatrix} \right|$$

**but** with  $x$  replaced by the corresponding value of  $y$ , that is, replace  $x$  by  $g^{-1}(y)$ .



- The **standard bivariate normal density** is

$$f_X(x_1, x_2) = \frac{1}{2\pi} \exp \left\{ -\frac{x_1^2 + x_2^2}{2} \right\}.$$

- Let  $Y = (Y_1, Y_2)$  where  $Y_1 = \sqrt{X_1^2 + X_2^2}$  and  $0 \leq Y_2 < 2\pi$  is angle from positive  $x$  axis to ray from origin to point  $(X_1, X_2)$ .
- I.e.,  $Y$  is  $X$  in polar co-ordinates.
- Solve for  $x$  in terms of  $y$ :

$$X_1 = Y_1 \cos(Y_2) \quad X_2 = Y_1 \sin(Y_2)$$



## Example

- This makes

$$\begin{aligned}g(x_1, x_2) &= (g_1(x_1, x_2), g_2(x_1, x_2)) \\&= (\sqrt{x_1^2 + x_2^2}, \text{argument}(x_1, x_2)) \\g^{-1}(y_1, y_2) &= (g_1^{-1}(y_1, y_2), g_2^{-1}(y_1, y_2)) \\&= (y_1 \cos(y_2), y_1 \sin(y_2)) \\\left| \frac{dx}{dy} \right| &= \left| \det \begin{pmatrix} \cos(y_2) & -y_1 \sin(y_2) \\ \sin(y_2) & y_1 \cos(y_2) \end{pmatrix} \right| \\&= y_1.\end{aligned}$$

- It follows that

$$f_Y(y_1, y_2) = \frac{1}{2\pi} \exp \left\{ -\frac{y_1^2}{2} \right\} y_1 1(0 \leq y_1 < \infty) 1(0 \leq y_2 < 2\pi).$$



## Marginal densities of $Y_1, Y_2$ ?

- Factor  $f_Y$  as  $f_Y(y_1, y_2) = h_1(y_1)h_2(y_2)$  where

$$h_1(y_1) = y_1 e^{-y_1^2/2} \mathbf{1}(0 \leq y_1 < \infty)$$

and

$$h_2(y_2) = \mathbf{1}(0 \leq y_2 < 2\pi)/(2\pi).$$

- Then

$$f_{Y_1}(y_1) = \int_{-\infty}^{\infty} h_1(y_1)h_2(y_2) dy_2 = h_1(y_1) \int_{-\infty}^{\infty} h_2(y_2) dy_2$$

so marginal density of  $Y_1$  is a multiple of  $h_1$ .



## Continued

- Multiplier makes  $\int f_{Y_1} = 1$  but in this case

$$\int_{-\infty}^{\infty} h_2(y_2) dy_2 = \int_0^{2\pi} (2\pi)^{-1} dy_2 = 1$$

so that  $Y_1$  has the Weibull or Rayleigh law

$$f_{Y_1}(y_1) = y_1 e^{-y_1^2/2} \mathbf{1}(0 \leq y_1 < \infty).$$

- Similarly

$$f_{Y_2}(y_2) = 1(0 \leq y_2 < 2\pi)/(2\pi)$$

which is the **Uniform**(0, 2 $\pi$ ) density.

- Exercise:  $W = Y_1^2/2$  has standard exponential distribution.
- Recall: by definition  $U = Y_1^2$  has a  $\chi^2$  dist on 2 degrees of freedom.
- Exercise: find  $\chi_2^2$  density.
- Notice that  $Y_1 \perp\!\!\!\perp Y_2$ .

