## **Queuing Theory**

Ingredients of Queuing Problem:

1: Queue input process.

2: Number of servers

**3**: Queue discipline: first come first serve? last in first out? pre-emptive priorities?

4: Service time distribution.

**Example**: Imagine customers arriving at a facility at times of a Poisson Process N with rate  $\lambda$ . This is the input process, denoted M (for Markov) in queuing literature.

Single server case:

Service distribution: exponential service times, rate  $\mu$ .

Queue discipline: first come first serve.

X(t) = number of customers in line at time t.

X is a Markov process called M/M/1 queue:

$$v_i = \lambda + \mu \mathbf{1}(i > 0)$$
 
$$\mathbf{P}_{ij} = \begin{cases} \frac{\mu}{\mu + \lambda} & j = i - 1 \ge 0\\ \frac{\lambda}{\mu + \lambda} & j = i + 1, i > 0\\ 1 & j = 1, i = 0\\ 0 & \text{otherwise} \end{cases}$$

**Example**:  $M/M/\infty$  queue:

Customers arrive according to PP rate  $\lambda$ . Each customer begins service immediately. X(t) is number being served at time t. X is a birth and death process with

$$v_n = \lambda + n\mu$$

and

$$\mathbf{P}_{ij} = \begin{cases} \frac{i\mu}{i\mu + \lambda} & j = i - 1 \ge 0\\ \frac{\lambda}{i\mu + \lambda} & j = i + 1\\ 0 & \text{otherwise} \end{cases}$$

Stationary distributions?

For M/M/1 queue:

Solve

$$\{\lambda + \mu \mathbf{1}(n > 0)\}\pi_n = \mu \pi_{n+1} + \lambda \mathbf{1}(n > 0)\pi_{n-1}$$

Just use general birth and death process formulation:

$$\lambda_n = \lambda \quad \mu_n = \mu \mathbf{1}(n > 0)$$

SO

$$\frac{\lambda_0 \cdots \lambda_{n-1}}{\mu_1 \cdots \mu_n} = (\lambda/\mu)^n$$

and

$$\sum_{n=0}^{\infty} (\lambda/\mu)^n = 1/(1 - \lambda/\mu)$$

SO

$$\pi_n = \frac{(\lambda/\mu)^n}{1 + 1/(1 - \lambda/\mu)}$$

Exists only if  $\lambda < \mu$ .

For  $M/M/\infty$  queue:

$$\pi_n \propto \frac{\lambda^n}{\mu^n n!}$$

and

$$\sum_{n=0}^{\infty} \frac{\lambda^n}{\mu^n n!} = \exp(\lambda/\mu)$$

SO

$$\pi_n = \exp(-\lambda/\mu) \frac{\lambda^n}{\mu^n n!}$$

Notice this exists for all  $\lambda > 0$  and all  $\mu > 0$ .

Scope of Queuing Theory:

1) M/M/k queues. X(t) is number queued or in service.

Birth and Death process; death rate maxes out at  $k\mu$ .

Stationary distribution exists if  $\lambda < k\mu$ .

2) Same input / service processes as M/M/k but customers not served leave. Question of interest: customers lost per time unit?

Take X to be number in service.  $(0 \le X(t) \le k)$ .

Find stationary distribution.

Fraction of time spent in state k is  $\pi_k$ .

During time in state k lose customers at rate  $\lambda$ . So lost  $\pi_k \lambda$  customers per unit time.

- 3) G/M/1 queue. General distribution of interarrival times for input. Input is a **renewal process**. Not Markov.
- 4) M/G/1 and others.
- 5) Networks: output of 1 queue is input of next; feedback ...

Quantities of potential interest:

Average fraction of time server idle.

Average time in system for customer.

Average wait to see server.

One example calculation: in G/M/1 queue.

Compute long run fraction time system is idle.

Idea: interarrival times are iid with cdf G.

Service rate  $\mu$ .

Let  $X_n$  be number of customers in service / in line when nth customer arrives.

Claim  $X_n$  is a Markov chain.

(Example of general tactic: find simple process buried within process of interest.)

Notation:  $T_1, T_2, \cdots$  iid interarrival times.

Given  $X_n = i$  and  $T_{n+1} = t$  number served between nth arrival and n+1st arrival is

$$\min\{ \mathsf{Poisson}(\mu t), i+1 \}$$

So: if  $X_n = i$  and the Poisson variable above is j then

$$X_{n+1} = i + 1 - \min\{j, i + 1\}$$

Now to compute prob of j served must average over  $T_{n+1}$ :

$$P(j \text{ served}) = \int e^{-\mu t} \frac{(\mu t)^j}{j!} dG(t) \equiv a_j$$
 for  $j < i+1$ .

This gives:

$$P_{ik} = \begin{cases} a_{i+1-k} & 1 \le k \le i+1 \\ 1 - \sum_{0}^{i} a_{j} & k = 0 \\ 0 & \text{otherwise} \end{cases}$$