## Econ 836 Midterm Exam

1. [14 points] Consider the following code and output from a log-wage regression using 2006 Census data on male residents of Toronto. The first line sets the line delimiter to ";".
. use "C:\DATA\2006 Census \pumf2006.dta", clear;
. 9 insamp=POB==1\&AGEGRP>8\&AGEGRP<18\&COW==4\&HDGREE>1\&HDGREE<88\&WAGES>100\&CFSIZE<9\&CFSIZE>0\&CMA==535\&SEX==2\&VISMIN<88;
. drop if insamp==0;
(833472 observations deleted)
. generate logwage=log(WAGES) ;
. generate not_alone=CFSIZE~=1;

- replace CFSIZE=CFSIZE-1;
(11004 real changes made)
. *NOTE for ethnic categories, everything comes from ABOID and VISMIN, and that "|" means "or";
. generate aborig=ABOID<6;
- generate white=VISMIN==13\&aborig==0;
- generate chinese=VISMIN==1\&aborig==0;
- generate southasian=VISMIN==2\&aborig==0;
- generate caribblack=VISMIN==3\&aborig==0;
- generate othvismin=VISMIN>4\&VISMIN<13\&aborig==0;
- generate vm=VISMIN<13\&aborig==0;
. generate notwhite=VISMIN<13|aborig==1;
. xi: regress logwage i.AGEGRP i.HDGREE i.MARST not_alone CFSIZE chinese southasian caribblack notwhite aborig;

| i. AGEGRP | -IAGEGRP_9-17 | (naturally coded; _IAGEGRP_9 omitted) |
| :--- | :--- | :--- |
| i.HDGREE | -IHDGREE_2-13 | (naturally coded; -IHDGREE_2 omitted) |
| i.MARST | _IMARST_1-5 | (naturally coded; _IMARST_1 omitted) |


| Source | SS | df | MS |
| :---: | :---: | :---: | :---: |
| Model | 2059.31498 | 30 | 68.6438326 |
| Residual | 7713.98876 | 10973 | . 702997245 |


| Number of obs | $=11004$ |
| ---: | :--- | ---: |
| F $(30,10973)$ | $=97.64$ |
| Prob $>$ F | $=0.0000$ |
| R-squared | $=0.2107$ |
| Adj R-squared | $=0.2086$ |
| Root MSE | $=.83845$ |


| logwage | Coef. | Std. Err. | t | $P>\|t\|$ | [95\% Conf. Interval] |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| _IAGEGRP_10 | . 3222072 | . 0283047 | 11.38 | 0.000 | . 2667248 | . 3776895 |
| _IAGEGRP_11 | . 4325841 | . 0305029 | 14.18 | 0.000 | . 3727929 | . 4923753 |
| _IAGEGRP_12 | . 518103 | . 0304299 | 17.03 | 0.000 | . 458455 | . 577751 |
| _IAGEGRP_13 | . 5740694 | . 031979 | 17.95 | 0.000 | . 5113848 | . 6367541 |
| _IAGEGRP_14 | . 5695695 | . 0349644 | 16.29 | 0.000 | . 501033 | . 638106 |
| _IAGEGRP_15 | . 5004705 | . 0408812 | 12.24 | 0.000 | . 420336 | . 5806049 |
| _IAGEGRP_16 | . 2319792 | . 050103 | 4.63 | 0.000 | . 1337683 | . 3301902 |
| _IAGEGRP_17 | -. 1181318 | . 0761101 | -1.55 | 0.121 | -. 2673213 | . 0310577 |
| _IHDGREE_3 | -. 041436 | . 0399403 | -1.04 | 0.300 | -. 1197262 | . 0368542 |
| _IHDGREE_4 | . 1697118 | . 039882 | 4.26 | 0.000 | . 0915359 | . 2478878 |
| _IHDGREE_5 | . 0435886 | . 0542511 | 0.80 | 0.422 | -. 0627535 | . 1499306 |
| _IHDGREE_6 | . 1232402 | . 0302268 | 4.08 | 0.000 | . 0639902 | . 1824901 |
| _IHDGREE_7 | . 2132195 | . 0291102 | 7.32 | 0.000 | . 1561582 | . 2702808 |
| _IHDGREE_8 | . 2578788 | . 0407753 | 6.32 | 0.000 | . 1779519 | . 3378056 |
| _IHDGREE_9 | . 4763058 | . 0228869 | 20.81 | 0.000 | . 4314433 | . 5211682 |
| _IHDGREE_10 | . 5585993 | . 0439811 | 12.70 | 0.000 | . 4723885 | . 6448101 |
| _IHDGREE_11 | . 3818988 | . 1296 | 2.95 | 0.003 | . 1278595 | . 6359382 |
| _IHDGREE_12 | . 570872 | . 0326154 | 17.50 | 0.000 | . 50694 | . 6348041 |
| _IHDGREE_13 | . 6211139 | . 0799199 | 7.77 | 0.000 | . 4644565 | . 7777714 |
| _IMARST_2 | . 2160764 | . 0361877 | 5.97 | 0.000 | . 145142 | . 2870107 |
| _IMARST_3 | -. 0304484 | . 0555723 | -0.55 | 0.584 | -. 13938 | . 0784833 |
| _IMARST_4 | -. 2207229 | . 0370627 | -5.96 | 0.000 | -. 2933724 | -. 1480733 |
| _IMARST_5 | -. 0848602 | . 1256429 | -0.68 | 0.499 | -. 3311428 | . 1614224 |
| not_alone | -. 1038007 | . 0294991 | -3.52 | 0.000 | -. 1616243 | -. 0459772 |
| CFSIZE | . 0472362 | . 0087003 | 5.43 | 0.000 | . 030182 | . 0642905 |
| chinese | . 0000648 | . 0757365 | 0.00 | 0.999 | -. 1483923 | . 1485219 |
| southasian | . 0016856 | . 0790663 | 0.02 | 0.983 | -. 1532985 | . 1566697 |
| caribblack | -. 123238 | . 073679 | -1.67 | 0.094 | -. 267662 | . 021186 |
| notwhite | -. 1239527 | . 0530695 | -2.34 | 0.020 | -. 2279785 | -. 0199269 |
| aborig | . 0332937 | . 1001404 | 0.33 | 0.740 | -. 1629995 | . 229587 |
| _cons | 10.20313 | . 0463208 | 220.27 | 0.000 | 10.11233 | 10.29392 |

a. Do Aboriginal men have lower earnings than white men, conditional on age, education, marital status, and household size? What is the conditional expectation of the difference in log-earnings between Aboriginal and white men?
b. The constant is highly significant, with a t-value of 220 . Is this surprising? Why or why not? What is the meaning of the constant term?
c. What is the predicted difference in log-earnings for a household with 1 member versus a household with 2 members?
d. Why is R-squared (equal to $\mathrm{V}(X \widehat{\beta}) / \mathrm{V}(\mathrm{Y}))$ so low when so many coefficients have big t-values?
e. Why is _IAGEGRP_9 omitted?
f. What is the average of the residual vector $e=Y-X \widehat{\beta}$ ?
g. Is the residual $e$ correlated with household size (the variable CFSIZE)?
2. [8 points] Pendakur and Pendakur (1998) estimate models of earnings which control for education, and investigate the differences in earnings across ethnic groups.
a. If there were unobserved quality variation for people with the same reported education level, how would this affect your interpretation of the estimates?
b. Assume that 'field-of-study' is available in the data (it is). Should it be included in the regression? Does excluding it induce bias? Why?
c. Does it matter that they drop all observations for which income from wages and salaries is zero?
d. Suppose these authors wanted to investigate the conditional median of log-earnings rather than the conditional mean. Would this give them a way to deal with zeroes?
2. [8 points] Allen, Pendakur and Suen (2005) estimates a panel model with the standard deviation of the log of age at first marriage on the LHS and no-fault status and state and year dummies on the RHS.
a. They do not include any information about the population of the state in the model. Likewise, there is not information on education levels in the state. Does this matter? Under what conditions does it not matter? Are these plausible conditions?
b. Why didn't they use the random effects FGLS estimator?
c. It could be that time affects every country differently. Why didn't they interact time dummies with country dummies?
d. These authors regress median age at first marriage in a state on legal characteristics. Could they have run a quantile regression to address their question? If so, what quantile regression?
4. [4 points] Suppose that: $Y_{i}=X_{i} \beta+\varepsilon_{i}$, for $i=1, \ldots, N ; \mathrm{X}$ is a single column with X a range between 1 and 2; and $E[\varepsilon]=0_{N}, E\left[\left(\varepsilon_{i}\right)^{2}\right]=\frac{\sigma^{2}}{\left(X_{i}\right)^{2}}$ and $E\left[\varepsilon_{i} \varepsilon_{j}\right]=0$ for all $i$ not equal to $j$. Here, the variance of the disturbance decreases with $X$, and there are no correlations in disturbances across observations.
a. What is the standard error of the OLS estimate of the (scalar) parameter in this case? Is it larger or smaller than the standard error given in regression output which assumes homoskedasticity? How much bigger or smaller?
b. Derive the GLS estimator for this case, and show how you would implement it. In what way would it treat observations where $X=2$ differently from those where $X=1$ ?
5. [6 points] Jacks and Pendakur (2010) estimate the effect of freight prices on international trade volumes.
a. Suppose they regressed $Y$ on covariates $X$ but not on country dummies or decade dummies. Under what conditions is this estimator unbiased? Under what conditions is it efficient?
b. Is there a better estimator for the case when it is unbiased but not efficient? If so, what? If not, why not?
c. What about regressing $Y$ on covariates $X$ and dummies for each decade (but not each country)? What about regressing $Y$ on $X$ and dummies for each decade and each country and the interaction of these two vectors of dummy variables.

