

Supplementary File for the Manuscript Titled “Semiparametric Mixture of Binomial Regression ” by J. Cao and W. Yao

Proofs

Let $g(t)$ be the density function for t . The following technical conditions are imposed in this section. They are not the weakest possible conditions, but they are imposed to facilitate the proofs.

Technical Conditions:

- A $\pi_1(t)$ and $p(t)$ has continuous second derivative at t_0 and $0 < \pi_1(t_0) < 1$ and $0 < p(t_0) < 1$. (For the constant proportion semiparametric mixture model (3), we use the same assumption for $p(t)$ and assume $0 < \pi_1 < 1$.)
- B $g(t)$ has continuous second derivative at the point t_0 and $g(t_0) > 0$.
- C $K(\cdot)$ is a symmetric (about 0) kernel density with compact support $[-1, 1]$.
- D The bandwidth h tends to zero such that $nh \rightarrow \infty$.

Let $\alpha_n = (nh)^{-1/2} + h^2$, $\boldsymbol{\theta}_0 = \{\pi_1(t_0), p(t_0)\}$,

$$f(x, \boldsymbol{\theta}) = \pi_1 I(x = 0) + \pi_2 \binom{N}{x} p^x \{1 - p\}^{N-x},$$

$l(x, \boldsymbol{\theta}) = \log f(x, \boldsymbol{\theta})$, where $\boldsymbol{\theta} = (\pi_1, p)$. Then the objective function (4) can be written as

$$\ell(\boldsymbol{\theta}) = \frac{1}{n} \sum_{i=1}^n K_h(t_i - t_0) \log f(x_i, \boldsymbol{\theta}) = \frac{1}{n} \sum_{i=1}^n K_h(t_i - t_0) l(x, \boldsymbol{\theta}).$$

Define

$$l_1(x, \boldsymbol{\theta}) = \frac{\partial}{\partial \boldsymbol{\theta}} l(x, \boldsymbol{\theta}) \quad \text{and} \quad l_2(x, \boldsymbol{\theta}) = \frac{\partial^2}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}^T} l(x, \boldsymbol{\theta}),$$

$G(t) = E\{l_1(X, \boldsymbol{\theta}_0) \mid t\}$ and $\mathcal{I}(t) = -E\{l_2(X, \boldsymbol{\theta}_0) \mid t\}$. The moments of K and K^2 are denoted respectively by

$$\mu_j = \int t^j K(t) dt \quad \text{and} \quad \nu_j = \int t^j K^2(t) dt.$$

By some simple calculations, we can get the following results.

Lemma 1. Assume that the regularity conditions A–C hold. We have the following results

1. The $G(t)$ has continuous second derivative at t_0 and $E\{l_1(X, \boldsymbol{\theta}_0)^2 \mid t\}$ is continuous at t_0 .
2. The $\partial^3 \ell(\boldsymbol{\theta}_0) / (\partial \boldsymbol{\theta}_i \partial \boldsymbol{\theta}_j \partial \boldsymbol{\theta}_k)$ is a bounded function for all $\boldsymbol{\theta}$ in a neighborhood of $\boldsymbol{\theta}_0$ and all x .
3. $\mathcal{I}(t)$ is continuous at t_0 and positive definite at t_0 and

$$\mathcal{I}(t_0) = E\{l_1(X, \boldsymbol{\theta}_0) l_1(X, \boldsymbol{\theta}_0)^T \mid t_0\}.$$

Proof of Theorem 2.1.

Note that

$$\ell(\boldsymbol{\theta}) = \frac{1}{n} \sum_{i=1}^n K_h(t_i - t_0) \log f(x_i, \boldsymbol{\theta}).$$

Hence,

$$\begin{aligned} \ell(\boldsymbol{\theta}^{(k+1)}) - \ell(\boldsymbol{\theta}^{(k)}) &= \sum_{i=1}^n \log \left\{ \frac{f(x_i, \boldsymbol{\theta}^{(k+1)})}{f(x_i, \boldsymbol{\theta}^{(k)})} \right\} K_h(t_i - t_0) \\ &= \sum_{i=1}^n \log \left\{ \frac{\pi_1^{(k)} B(x_i, N, 0)}{f(x_i, \boldsymbol{\theta}^{(k)})} \frac{\pi_1^{(k+1)} B(x_i, N, 0)}{\pi_1^{(k)} B(x_i, N, 0)} \right. \\ &\quad \left. + \frac{\pi_2^{(k)} B(x_i, N, p^{(k)})}{f(x_i, \boldsymbol{\theta}^{(k)})} \frac{\pi_2^{(k+1)} B(x_i, N, p^{(k+1)})}{\pi_2^{(k)} B(x_i, N, p^{(k)})} \right\} K_h(x_i - x_0) \\ &= \sum_{i=1}^n \log \left\{ r_{i1}^{(k+1)} \frac{\pi_1^{(k+1)} B(x_i, N, 0)}{\pi_1^{(k)} B(x_i, N, 0)} + r_{i2}^{(k+1)} \frac{\pi_2^{(k+1)} B(x_i, N, p^{(k+1)})}{\pi_2^{(k)} B(x_i, N, p^{(k)})} \right\} K_h(x_i - x_0) \end{aligned}$$

Based on the Jensen's inequality, we have

$$\begin{aligned} \ell(\boldsymbol{\theta}^{(k+1)}) - \ell(\boldsymbol{\theta}^{(k)}) &\geq \sum_{i=1}^n \left[r_{i1}^{(k+1)} \log \left\{ \frac{\pi_1^{(k+1)} B(x_i, N, 0)}{\pi_1^{(k)} B(x_i, N, 0)} \right\} K_h(x_i - x_0) \right. \\ &\quad \left. + r_{i2}^{(k+1)} \log \left\{ \frac{\pi_2^{(k+1)} B(x_i, N, p^{(k+1)})}{\pi_2^{(k)} B(x_i, N, p^{(k)})} \right\} K_h(x_i - x_0) \right] \end{aligned}$$

Based on the property of M-step of (5), we have

$$\ell(\boldsymbol{\theta}^{(k+1)}) - \ell(\boldsymbol{\theta}^{(k)}) \geq 0.$$

Proof of Theorem 3.1. Denote $\alpha_n = (nh)^{-1/2} + h^2$. It is sufficient to show that for any given $\eta > 0$, there exists a large constant c such that

$$P\left\{\sup_{\|u\|=c} \ell(\boldsymbol{\theta}_0 + \alpha_n u) < \ell(\boldsymbol{\theta}_0)\right\} \geq 1 - \eta, \quad (13)$$

where $\ell(\boldsymbol{\theta})$ is defined in (4).

By using Taylor expansion, it follows that

$$\begin{aligned} \ell(\boldsymbol{\theta}_0 + \alpha_n u) - \ell(\boldsymbol{\theta}_0) &= \frac{1}{n} \sum_{i=1}^n K_h(t_i - t_0) \{l(x_i, \boldsymbol{\theta}_0 + \alpha_n u) - l(x_i, \boldsymbol{\theta}_0)\} \\ &= \frac{1}{n} \sum_{i=1}^n K_h(t_i - t_0) \left\{ l_1(x_i, \boldsymbol{\theta}_0)^T u \alpha_n + u^T l_2(x_i, \boldsymbol{\theta}_0) u \alpha_n^2 + \alpha_n^3 q(x_i, \tilde{\boldsymbol{\theta}}) \right\} \\ &= I_1 + I_2 + I_3, \end{aligned}$$

where $\|\tilde{\boldsymbol{\theta}} - \boldsymbol{\theta}_0\| \leq c\alpha_n$ and

$$q(x_i, \tilde{\boldsymbol{\theta}}) = \sum_{i=1}^2 \sum_{j=1}^2 \sum_{k=1}^2 \frac{\partial^3 l(x_i, \tilde{\boldsymbol{\theta}})}{\partial \theta_i \partial \theta_j \partial \theta_k} u_i u_j u_k,$$

where $u = (u_1, u_2)$.

By directly calculating the mean and variance and note that $G(t_0) = 0$, we obtain

$$\begin{aligned} E(I_1) &= \alpha_n E\{K_h(t - t_0) G(t)^T u\} = O(c\alpha_n h^2); \\ \text{var}(I_1) &= n^{-1} \alpha_n^2 \text{var}[K_h(t_i - t_0) l_1(\boldsymbol{\theta}_0, x_i)^T u] = O(c^2 \alpha_n^2 (nh)^{-1}). \end{aligned}$$

Hence

$$I_1 = O(c\alpha_n h^2) + \alpha_n c O_p((nh_1)^{-1/2}) = O_p(c\alpha_n^2).$$

Similarly,

$$I_3 = \frac{1}{n} \sum_{i=1}^n K_h(t_i - t_0) \alpha_n^3 q(x_i, \tilde{\boldsymbol{\theta}}) = O_p(\alpha_n^3).$$

and

$$I_2 = \frac{1}{n} \sum_{i=1}^n K_h(t_i - t_0) u^T l_2(x_i, \boldsymbol{\theta}_0) u \alpha_n^2 = -\alpha_n^2 g(t_0) u^T \mathcal{I}(t_0) u (1 + o_p(1)).$$

Noticing that $\mathcal{I}(t_0)$ is a positive matrix, $\|u\| = c$, we can choose c large enough such that I_2 dominates both I_1 and I_3 with probability at least $1 - \eta$. Thus (13) holds. Hence with probability approaching 1 (wpa1), there exists a local maximizer $\hat{\boldsymbol{\theta}}$ such that $\|\hat{\boldsymbol{\theta}} - \boldsymbol{\theta}_0\| \leq \alpha_n c$, where $\alpha_n = (nh)^{-1/2} + h^2$. Based on the definition of $\boldsymbol{\theta}$, we can also get, wpa1, $|\hat{\pi}(t_0) - \pi(t_0)| = O_p((nh)^{-1/2} + h^2)$ and $|\hat{p}(t_0) - p(t_0)| = O_p((nh)^{-1/2} + h^2)$.

Proof of Theorem 3.2.

Note that the estimate $\hat{\boldsymbol{\theta}}$ satisfies the equation

$$0 = \ell'(\hat{\boldsymbol{\theta}}) = \frac{1}{n} \sum_{i=1}^n K_h(t_i - t_0) \left\{ l_1(x_i, \boldsymbol{\theta}_0) + l_2(x_i, \boldsymbol{\theta}_0)(\hat{\boldsymbol{\theta}} - \boldsymbol{\theta}_0) + O_p(\|\hat{\boldsymbol{\theta}} - \boldsymbol{\theta}_0\|^2) \right\}. \quad (14)$$

The order of the third term could be derived from the (2) of Lemma 1. Let

$$\begin{aligned} W_n &= \frac{1}{n} \sum_{i=1}^n K_h(t_i - t_0) l_1(x_i, \boldsymbol{\theta}_0) \\ \Delta_n &= -\frac{1}{n} \sum_{i=1}^n K_h(t_i - t_0) l_2(x_i, \boldsymbol{\theta}_0). \end{aligned}$$

Note that

$$\begin{aligned} E(W_n) &= E\{K_h(t - t_0)G(t)\} = \frac{1}{2}(Gg)''(t_0)\mu_2 h^2(1 + o(1)), \\ \text{cov}(W_n) &= n^{-1} \text{cov}\{K_h(t_i - t_0)l_1(x_i, \boldsymbol{\theta}_0)\} \\ &= n^{-1} \{E K_h^2(t_i - t_0)l_1(x_i, \boldsymbol{\theta}_0)l_1(x_i, \boldsymbol{\theta}_0)^T - E(W_n)^2\} \\ &= (nh)^{-1}g(t_0)\mathcal{I}(t_0)\nu_0(1 + o(1)), \end{aligned} \quad (15)$$

where $(Gg)''(t)$ is the second derivative of $G(t)g(t)$, and

$$\begin{aligned} E(\Delta_n) &= E\{K_h(t - t_0)\mathcal{I}(t)\} = \mathcal{I}(t_0)g(t_0) + o(1), \\ \text{var}(\Delta_n(i, j)) &\leq n^{-1}E \left[K_h^2(t_i - t_0) \left\{ \frac{\partial^2 l(x_i, \boldsymbol{\theta}_0)}{\partial \boldsymbol{\theta}_i \partial \boldsymbol{\theta}_j} \right\}^2 \right] \\ &= O\{(nh)^{-1}\} = o(1). \end{aligned}$$

Therefore, we have

$$\Delta_n = \mathcal{I}(t_0)g(t_0) + o_p(1).$$

Note that $\|\hat{\boldsymbol{\theta}} - \boldsymbol{\theta}_0\|^2 = o_p(W_n)$. Then from (14), we have

$$\sqrt{nh}(\hat{\boldsymbol{\theta}} - \boldsymbol{\theta}_0) = g(t_0)^{-1}\mathcal{I}(t_0)^{-1}\sqrt{nh}W_n(1 + o_p(1)). \quad (16)$$

In order to prove the asymptotic normality of (16), we only need to establish the asymptotic normality of $\sqrt{nh}W_n$. Next we show, for any unit vector $d \in \mathbb{R}^2$, we prove

$$\{d^T \text{cov}(W_n^*)d\}^{-\frac{1}{2}} \{d^T W_n^* - d^T E(W_n^*)\} \xrightarrow{L} N(0, 1),$$

where $W_n^* = \sqrt{nh}W_n$. Let

$$\xi_i = \sqrt{h/n}K_h(t_i - t_0)d^T l_1(\boldsymbol{\theta}_0, x_i).$$

Then $d^T W_n^* = \sum_{i=1}^n \xi_i$. We check the Lyapunov's condition. Based on (15), we can get $\text{cov}(W_n^*) = g(t_0)\mathcal{I}(t_0)\nu_0(1+o(1))$ and $\text{var}(d^T W_n^* d) = d^T \text{cov}(W_n^*) d = g(t_0)\nu_0 d^T \mathcal{I}(t_0) d(1+o(1))$. So we only need to prove $nE|\xi_1|^3 \rightarrow 0$. Noticing that $l_1(\theta_0, x)$ is bounded for any x , and $K(\cdot)$ has compact support,

$$\begin{aligned} nE|\xi_1|^3 &\leq O(nn^{-3/2}h^{3/2})E|K_h^3(t_i - t_0)| \\ &= O(n^{-1/2}h^{3/2})O(h^{-2}) = O((nh)^{-1/2}) \rightarrow 0. \end{aligned}$$

So the asymptotic normality for W_n^* holds such that

$$\sqrt{nh} \left\{ W_n - \frac{1}{2}(Gg)''(t_0)\mu_2 h^2 + o(h^2) \right\} \xrightarrow{D} N\{0, g(t_0)\mathcal{I}(t_0)\nu_0\}.$$

Based on (16) and the Slutsky theorem, we can get the asymptotic result of $\hat{\theta}$

$$\sqrt{nh} \left\{ \hat{\theta} - \theta_0 - b(t_0)h^2 + o(h^2) \right\} \xrightarrow{D} N\{0, g^{-1}(t_0)\mathcal{I}^{-1}(t_0)\nu_0\},$$

where

$$b(t_0) = \mathcal{I}^{-1}(t_0) \left\{ \frac{G'(t_0)g'(t_0)}{g(t_0)} + \frac{1}{2}G''(t_0) \right\} \mu_2.$$

Proof of Theorem 3.3.

Let

$$f(x_i, \pi_1, \hat{p}(t_i)) = \log \left[\pi_1 I(x_i = 0) + \pi_2 \binom{N}{x_i} \hat{p}(t_i)^{x_i} (1 - \hat{p}(t_i))^{N-x_i} \right].$$

Based on a Taylor expansion of (4), similar to the proof of Theorem 3.2, we have that

$$\sqrt{n}(\tilde{\pi}_1 - \pi_1) = B_n^{-1}A_n + o_p(1).$$

where

$$\begin{aligned} A_n &= \frac{1}{\sqrt{n}} \sum_{i=1}^n \frac{\partial f(x_i, \pi_1, \hat{p}(t_i))}{\partial \pi_1} \\ B_n &= -\frac{1}{n} \sum_{i=1}^n \frac{\partial^2 f(x_i, \pi_1, \hat{p}(t_i))}{\partial \pi_1^2} \end{aligned}$$

It can be shown that

$$\begin{aligned} B_n &= -E \left\{ \frac{\partial^2 f(x_i, \pi_1, p(t_i))}{\partial \pi_1^2} \right\} + o_p(1) \\ &= \mathcal{I}_{\pi_1} + o_p(1). \end{aligned}$$

It can be shown that

$$\begin{aligned} A_n &= \frac{1}{\sqrt{n}} \sum_{i=1}^n \frac{\partial f(x_i, \pi_1, p(t_i))}{\partial \pi_1} + \frac{1}{\sqrt{n}} \sum_{i=1}^n \frac{\partial^2 f(x_i, \pi_1, p(t_i))}{\partial \pi_1 \partial p} \{\hat{p}(t_i) - p(t_i)\} + O_p(d_{1n}) \\ &= \frac{1}{\sqrt{n}} \sum_{i=1}^n \frac{\partial f(x_i, \pi_1, p(t_i))}{\partial \pi_1} + S_{n1} + O_p(d_{1n}). \end{aligned}$$

where $d_{1n} = n^{-1/2} \|\tilde{\pi}_1 - \pi_1\|_\infty^2 = o_p(1)$. Based on the proof of Theorem 3.2, we have

$$\hat{\boldsymbol{\theta}}(t_i) - \boldsymbol{\theta}(t_i) = \frac{1}{n} g(t_i)^{-1} \mathcal{I}(t_i)^{-1} \sum_{j=1}^n K_h(t_j - t_i) l_1(x_j, \boldsymbol{\theta}(t_i)) + O_p(d_{n2}),$$

Based on Carroll et al. (1997) and Li and Liang (2008), we have that $n^{1/2} d_{n2} = o_p(1)$ uniformly in t_i , if $nh^2/\log(1/h) \rightarrow \infty$. Let $\psi(t_j, x_j)$ be the second entry of $\mathcal{I}(t_j)^{-1} l_1(x_j, \boldsymbol{\theta}(t_j))$. Since $p(t_i) - p(t_j) = O(t_i - t_j)$ and $K(\cdot)$ is symmetric about 0, we have

$$\begin{aligned} S_{n1} &= \frac{1}{n^{-3/2}} \sum_{j=1}^n \sum_{i=1}^n \frac{\partial^2 f(x_i, \pi_1, p(t_i))}{\partial \pi_1 \partial p} g(t_i)^{-1} \psi(t_j, x_j) K_h(t_j - t_i) + O_p(n^{1/2} h^2) \\ &= S_{n2} + O_p(n^{1/2} h^2). \end{aligned}$$

It can be shown, by calculating the second moment, that

$$S_{n2} - S_{n3} = o_p(1),$$

where $S_{n3} = -n^{-1/2} \sum_{j=1}^n \xi(t_j, x_j)$, with

$$\begin{aligned} \xi(t_j, x_j) &= -E \left\{ \frac{\partial^2 f(x, \pi_1, p(t_j))}{\partial \pi_1 \partial p} \mid t = t_j \right\} \psi(t_j, x_j) \\ &= \mathcal{I}_{\pi_1 p}(t_j) \psi(t_j, x_j). \end{aligned}$$

By condition $nh^4 \rightarrow 0$, we know

$$A_n = n^{-1/2} \sum_{i=1}^n \left\{ \frac{\partial f(x_i, \pi_1, p(t_i))}{\partial \pi_1} - \xi(t_i, x_i) \right\} + o_p(1).$$

We can show that $E(A_n) = 0$. Define

$$\Sigma = \text{var}(A_n) = \text{var} \left\{ \frac{\partial f(x, \pi_1, p(t))}{\partial \pi_1} - \xi(t, x) \right\}.$$

Based on the central limit theorem, we can have

$$\sqrt{n}(\tilde{\pi}_1 - \pi_1) \rightarrow N(0, \mathcal{I}_{\pi_1}^{-2} \Sigma).$$

Proof of Theorem 3.4.

Based on a Taylor expansion of (7), similar to the proof of Theorem 3.2, we have

$$\sqrt{nh} \{\tilde{p}(t_0) - p(t_0)\} = g(t_0)^{-1} \mathcal{I}_p(t_0)^{-1} \tilde{W}_n(1 + o_p(1)),$$

where

$$\mathcal{I}_p(t) = -E \left\{ \frac{\partial^2 f(x, \pi_1, p(t_0))}{\partial p^2} \mid t \right\}$$

and

$$\tilde{W}_n = \sqrt{\frac{h}{n}} \sum_{i=1}^n \frac{\partial f(x_i, \pi_1, p(t_0))}{\partial p} K_h(t_i - t_0).$$

It can be calculated that

$$\tilde{W}_n = \sqrt{\frac{h}{n}} \sum_{i=1}^n \frac{\partial f(x_i, \tilde{\pi}_1, p(t_0))}{\partial p} K_h(t_i - t_0) + C_n + o_p(1),$$

where

$$C_n = \sqrt{\frac{h}{n}} \sum_{i=1}^n \frac{\partial^2 f(x_i, \pi_1, p(t_0))}{\partial p \partial \pi_1} (\tilde{\pi}_1 - \pi_1) K_h(t_i - t_0).$$

Since $\sqrt{n}(\tilde{\pi}_1 - \pi_1) = O_p(1)$, it can be shown that

$$C_n = o_p(1).$$

Hence

$$\sqrt{nh}\{\tilde{p}(t_0) - p(t_0)\} = g(t_0)^{-1} \mathcal{I}_p(t_0)^{-1} W_n (1 + o_p(1)),$$

where

$$W_n = \sqrt{\frac{h}{n}} \sum_{i=1}^n \frac{\partial f(x_i, \pi_1, p(t_0))}{\partial p} K_h(t_i - t_0).$$

Let

$$\Gamma(t) = \mathbb{E} \left\{ \frac{\partial f(x, \pi_1, p(t_0))}{\partial p} \middle| t \right\}.$$

Note that $\Gamma(t_0) = 0$. We can show that

$$\text{var}(W_n) = \mathcal{I}_p(t_0) g(t_0) \nu_0 (1 + o_p(1))$$

and

$$\mathbb{E}(W_n) = \frac{\sqrt{nh}}{2} \{ \Gamma''(t_0) g(t_0) + 2\Gamma'(t_0) g'(t_0) \} h^2 \mu_2 (1 + o_p(1)).$$

Similar to the proof of Theorem 3.2, we can prove the asymptotic normality of W_n . Hence, we have

$$\sqrt{nh}\{\tilde{p}(t_0) - p(t_0) - \tilde{b}(t_0)h^2\} \xrightarrow{D} N(0, g(t_0)^{-1} \mathcal{I}_p(t_0)^{-1} \nu_0),$$

where

$$\tilde{b}(t_0) = \frac{1}{2g(t_0)\mathcal{I}_p(t_0)} \{ \Gamma''(t_0) g(t_0) + 2\Gamma'(t_0) g'(t_0) \} \mu_2.$$